# A MANUAL FOR THE ESTABLISHMENT AND ASSESSMENT OF HORIZONTAL SURVEY NETWORKS IN THE MARITIME PROVINCES 

D. B. THOMSON<br>E. J. KRAKIWSKY<br>B. G. NICKERSON

## PREFACE

In order to make our extensive series of technical reports more readily available, we have scanned the old master copies and produced electronic versions in Portable Document Format. The quality of the images varies depending on the quality of the originals. The images have not been converted to searchable text.

# A MANUAL FOR THE ESTABLISHMENT AND ASSESSMENT OF HORIZONTAL SURVEY NETWORKS IN THE <br> <br> MARITIME PROVINCES 

 <br> <br> MARITIME PROVINCES}
by
D. B. Thomson
E. J. Krakiwsky
B. G. Nickerson

Department of Surveying Engineering University of New Brunswick Fredericton, N.B. Canada

## PREFACE

This "manual" is the third in the series designed to assist surveyors, in the maritime provinces, on the correct and practical use of the geodetic information of the redefined Maritime Geodetic Network. It has been written as a surveyors handbook for the design, computation, and assessment of horizontal geodetic networks. In this report, a geodetic network is considered to be any geometric configuration of three or more terrestrial survey points. The points may be connected via any combination of direction, angle, azimuth, and distance observations; furthermore, there may be redundant observations leading to overdetermined cases. The networks are treated in only one environment in this manual, the conformal mapping plane. There are two sound reasons for this:
(1) this is the environment in which most practicing surveyors wish to do their network computations, (ii) derived quantities - coordinates, distances, azimuths and their associated covariance matrices - can be transformed, if required, to the 2-D ellipsoidal and 3-D environments using the methodologies outlined in "A Manual for Geodetic Coordinate Transformations in the Maritimes" [Krakiwsky et. al., 1977] and "A Manual for Geodetic Position Computations in the Maritime Provinces" [Thomson et. al., 1978] respectively. This approach (rigorous transformation of 2-D plane information to 2-D ellipsoidal or 3-D) is equivalent to carrying out the original computations in the environment itself (e.g. 2-D ellipsoidal, 3-D).

No extensive derivations or explantions of the mathematical formulae used are given. The equations required to solve certain problems are stated, the notation is explained, and numerical examples
are presented. A reader desiring extensive background information is referred to the reference material.

The authors would like to acknowledge the financial assistance given by the Surveys and Mapping Division of the Land Registration and Information Service to support the preparation of this report. We would like to thank R. Steeves, a graduate student in Surveying Engineering, for his excellent work in the preparation of the numerical examples and associated computer programs. S. Biggar is acknowledged for her patience and dedication in the typing and final preparation of this report.

TABLE OF CONTENTS
Preface ..... i
Acknowledgements ..... ii
Table of Contents ..... iv
List of Figures ..... vi
List of Tables ..... viii

1. Introduction ..... 1
2. Observations and their Reductions ..... 6
2.1 Accuracies of Observed Azimuths, Directions, Angles and Distances ..... 6
2.2 Reduction of Observations to a Conformal Mapping Plane ..... 20
2.3 Data Screening ..... 32
3. Mathematical Models for Azimuth, Direction, and Angle Observations ..... 37
3.1 Azimuth Mathematical Model ..... 37
3.2 Direction Mathematical Model ..... 40
3.3 Angle Mathematical Model ..... 43
4. Mathematical Models for Distance Observations ..... 47
5. Solution of Unique Cases ..... 50
5.1 Direct Problem ..... 51
5.2 Azimuth Intersection ..... 57
5.3 Distance Intersection ..... 61
5.4 Angle Resection ..... 65
5.5 Open Traverse ..... 70
6. Solution of Overdetermined Cases ..... 77
6.1 Closed Traverse ..... 78
6.2 Network ..... 86
7. A Priori Knowledge of Parameters ..... 97
8. Preanalysis ..... 109
8.1 Traverse Design ..... 110
8.2 Property Survey Design ..... 118

## Table of Contents (continued)

9. Postanalysis ..... 125
References ..... 129
Appendix I: Taylor's Series ..... 131
Appendix II: Least Squares Method ..... 134
Appendix III: Error Ellipses ..... 139
Figure 2.1 Angles and Directions ..... 15
Figure 2.2 Reduction of Observed Astronomic Azimuth to a Conformal Mapping Plane ..... 24
Figure 2.3 Reduction of an Observed Direction to a Conformal Mapping Plane ..... 26
Figure 2.4 Reduction of an Observed Angle to a Conformal Mapping Plane ..... 28
Figure 2.5 Reduction of as Terrain Spatial Distance to a Conformal Mapping Plane ..... 30
Figure 3.1 Mapping Plane Azimuth ..... 38
Figure 3.2 Direction on the Mapping Plane ..... 41
Figure 3.3 Angle on the Mapping Plane ..... 44
Figure 4.1 Distance on the Mapping Plane ..... 48
Figure 5.1 Direct Problem ..... 51
Figure 5.2 Confidence Ellipse for Direct Problem ..... 56
Figure 5.3 Azimuth Intersection on the Plane ..... 57
Figure 5.4 95\% Confidence Ellipse for Azimuth Intersection ..... 60
Figure 5.5 Distance Intersection ..... 61
Figure 5.6 95\% Confidence Ellipse for Distance Intersection ..... 65
Figure 5.7 Angle Resection ..... 66
Figure 5.8 95\% Confidence Ellipse from Angle Resection ..... 69
Figure 5.9 Open Traverse ..... 71
Figure 5.10 95\% Confidence Ellipses for Open Traverse ..... 76
Figure 6.1 Least Squares Line Fitting ..... 77
Figure 6.2 Closed Traverse ..... 79
Figure 6.3 Plotof 95\% Error Ellipses for the Closed Traverse ..... 85

## List of Figures (Continued)

Figure 6.4 Network ..... 87
Figure 6.5 Plot of 95\% Error Ellipses for the Network ..... 96
Figure 7.1 Direct Problem with Weighted Point ..... 98
Figure 7.2 Plot of 95\% Error Ellipses for Direct Case with Weighted Parameters ..... 101
Figure 7.3 Initial Configuration for Closed Traverse with Weighted Points ..... 103
Figure 7.4 Plot of 95\% Error Ellipses for Closed Traverse with Weighted points ..... 107
Figure 8.1 Initial Data Plot ..... 111
Figure 8.2 Initial Observables Plot ..... 112
Figure 8.3 Plot of $99 \%$ Confidence Ellipses from Initial Observables ..... 115
Figure 8.4 Plot of Update Observables ..... 116
Figure 8.5 Plot of Updated 99\% Error Ellipses ..... 117
Figure 8.6 Initial Data Plot for Property Survey Design ..... 119
Figure 8.7 Plot of Initial Observables for Property Survey Design ..... 120
Figure 8.8 Initial Design Results ..... 121
Figure 8.9 Final Plot of Property Survey Design ..... 123
Figure AI. 1 Geometric Interpretation of Linear Taylor's Series ..... 131
Figure AIII: 1 Error Ellipse ..... 140
Figure AIII. 2 Error Circle and Confidence Ellipse ..... 147
Table 2.1 Expected Values of $\sigma_{A}$ Using Hour Angle Method for $\phi=A=Z=45^{\circ}$ ..... 8
Table 2.2 Expected Values of $\sigma_{A}$ Using Star Altitude Method for $\phi=A=45^{\circ}$ ..... 10
Table 2.3 Expected Centering Error ..... 12
Table 2.4 Expected Values of $\sigma_{d}$ for $h=5^{\circ}$ and $\sigma_{c}=1 \mathrm{~mm}$ ..... 13
Table 2.5 Major Features of Some Modern Theodolites ..... 14
Table 2.6 Characteristics of Modern EDM ..... 17
Table 2.7 Effect of Meteorological Errors on Measured Distances ..... 20
Table 2.8 Expected Values for $\sigma_{r}$ ..... 21
Table 2.9 Reduction of Observed Astronomic Azimuth to a Conformal Mapping Plane ..... 25
Table 2.10 Reduction of an Observed Direction to a Conformal Mapping Plane ..... 27
Table 2.11 Reduction of an Observed Angle to a Conformal Mapping Plane ..... 29
Table 2.12 Reduction of a Terrain Spatial Distance to a Confromal Mapping Plane ..... 31
Table 2.13 Testing for Outliers ..... 33
Table 2.14 Cumulative Normal Distributions - Values of Pr ..... 35
Table 2.15 Astronomic Azimuth Data Series ..... 36
Table 5.1 Initial Data for Direct Problem ..... 51
Table 5.2 Initial Data for Azimuth Intersection ..... 58
Table 5.3 Initial Data for Distance Intersection ..... 61
Table 5.4 Initial Data for Angle Resection ..... 65
Table 5.5 Initial Data for Open Traverse ..... 71
Table 5.6 Station and Relative Ellipses for Open Traverse ..... 75
Table 6.1 Initial Data for Closed Traverse ..... 80
Table 6.2 95\% Error Ellipses for Closed Traverse ..... 86
Table 6.3 Initial Coordinates for Network Stations ..... 88
Table 6.4 Observations on the Mapping Plane for the Network ..... 89
Table 6.5 95\% Error Ellipses for Network ..... 95
Table 7.1 95\% Error Ellipses for Closed Traverse with Weighted Parameters ..... 106
Table 8.1 Initial Data for Traverse Design ..... 110
Table 8.2 Initial Observables ..... 111
Table 8.3 Confidence Ellipses from Initial Observables ..... 115
Table 8.4 Update observables ..... 116
Table 8.5 99\% Error Ellipses After Update ..... 117
Table 8.6 Observable Summary ..... 118
Table 8.7 Initial Data for Property Survey Designs ..... 118
Table 8.8 Initial Observables for Property Survey Design ..... 119
Table 8.9 Final 95\% Error Ellipses for Property Survey Design ..... 123
Table AIII. $1 \quad \chi^{2}$ Distribution ..... 141
Table AIII. 2 F Distribution for $(1-\alpha)=0.95$ ..... 144
Table AIII. 3 Horizontal Survey Classification ..... 148

## 1. INTRODUCTION

It has been shown that the coordinate definition and associated accuracy estimates for any terrain point can be expressed equivalently in three dimensional cartesian coordinates ( $X, Y, Z ; C_{x, y, z}$ ), in ellipsoidal coordinates ( $\sigma, \lambda ; C_{\phi_{,} \lambda}$ ), or in conformal mapping plane coordinates ( $x, y ; C_{x, y}$ ), since the rigorous transformations between these quantities are well known [e.g. Krakiwsky et. al., 1977]. When an unknown terrain point is observed (e.g. an azimuth and a distance) from a known terrain point, the determination of the unknown coordinates and associated covariance matrix can be done in three dimensional space, on the surface of a reference ellipsoid, or on a conformal mapping plane [e.g. Thomson et. al, 1978]. The equivalence of results (coordinates, covariance matrix) in the threeenvironments is attained through the rigorous reduction (recall that no reductions are required for three dimensional computations) of the spatial measurements to the chosen computation surface. We can conclude from this that the choice of an environment in which to carry out position computations is, from a mathematical point of view, arbitrary. This fact is very important in the present context as it permits us to study the establishment and assessment of horizontal geodetic networks in the conformal mapping plane environment with the assurance that the procedures used yield results equivalent to those used in the three dimensional and ellipsoidal surface environments. Since the conformal mapping plane mathematical models involved with the establishment of horizontal geodetic networks are easy to understand, and since in practice a majority of surveyors prefer to use plane coordinates, the entire subject matter of this manual is treated
in only one envixonment - the conformal mapping plane. For a treatment of this subject matter in the three-dimensional and ellipsoidal surface environments, the interested reader is referred to, for example, Vincenty [1973] and Krakiwsky and Thomson [1978] respectively. The establishment and assessment of one-dimensional vertical networks is beyond the scope of this present work. A knowledge of the treatment of vertical networks is vital for surveyors, and the reader is referred, for example, to Vanicek and Krakiwsky [in prep.].

In this manual, a horizontal geodetic network is considered to be any geometric configuration of three or more terrain points connected via any combination of azimuth, direction, angle, and distance observations. The horizontal network may be such that only a unique solution for the coordinates of unknown points is possible (no. observations $n=$ no. unknowns u), or there may be redundant observations in which case we say the network is overdetermined (no. observations $n>$ no. unknowns $u$ ). For both cases, contemporary mathematical and statistical concepts and methodology are used.

The fundamental concept utilized is that of a mathematical model. A mathematical model is defined as a functional relationship between some unknown parameters $x$ (coordinates of unknown points) and some observables \& (azimuths, directions, angles, distances). There are two mathematical models that are of interest to us: the direct (explicit) form

$$
\begin{equation*}
x=g(\ell) \tag{1-1}
\end{equation*}
$$

in which $g$ is an explicit, functional relationship, and the inverse (explicit) form

$$
\begin{equation*}
\ell=h(x) \tag{1-2}
\end{equation*}
$$

in which $h$ is another explicit, functional relationship. Both forms of these models are well-known to surveyors. For example, the direct (explicit) form is used in position computations on a conformal mapping plane [e.g. Thomson et. al., 1978; p. 112, eqs. (4-39) and (4-40)]

$$
\begin{aligned}
& x_{j}=x_{i}+\ell_{i j} \sin t_{i j}, \\
& y_{j}=y_{i}+\ell_{i j} \cos t_{i j}
\end{aligned}
$$

The solution for either of the coordinates of the new point ( $x_{j}, y_{j}$ ) simply involves the evaluation of either of the two equations $\left(x_{i}, Y_{i}\right.$ are known, $l_{i j}\left(c h o r d\right.$ length) and $t_{i j}$ (grid azimuth of chord) are reduced measured quantities). This type of mathematical model (direct explicit) lends itself to geodetic position computations (only one unknown point to be considered). The inverse (explicit) form is also well known, for example, the expressions for a distance and azimuth respectively [e.g. Thomson et. al., 1978; p. 112, eqs. (4-41) and (4-42)]

$$
\begin{aligned}
\ell_{i j}= & {\left[\left(x_{j}-x_{i}\right)^{2}+\left(y_{j}-y_{i}\right)^{2}\right]^{1 / 2} } \\
& t_{i j}=\tan ^{-1} \frac{\left(x_{j}-x_{i}\right)}{\left(y_{j}-y_{i}\right)}
\end{aligned}
$$

Recall, however, that the objective is to solve for the coordinates of an unknown point (say $x_{j}, y_{j}$ ) using observed $\ell_{i j}$ and $t_{i j}$. In this instance, neither equation can be solved directly. They must be used together to get a solution for $\left(X_{j}, y_{j}\right)$. This requires special techniques which become extremely important when there are redundant observations. The inverse
(explicit) form of the mathematical model leads itself particularly to the computation of geodetic networks, and can also be easily used for geodetic position computations. In this manual, the inverse (explicit) model is used exclusively.

There are, of course, other forms of mathematical models. These are considered to be outside the scope of this work, but for the solution of certain special problems are important. For a complete coverage of the topic of mathematical models, the reader is referred, for example, to Vanicek and Krakiwsky [in prep].

This manual presents the step-by-step mathematical and computational procedures required for the establishment and assessment of a horizontal geodetic network on a conformal mapping plane. The procedures with respect to different conformal mapping planes only vary in the reduction of measured quantities; therefore, this distinction is only made in Chapter 2 which covers Observations and Their Reductions. In addition to a review of the reductions of measurements to a conformal mapping plane, Chapter 2 also includes (i) a brief review of the instrumentation used to observe azimuths, directions, angles, and distances and the accuracy estimates (variances) one should expect to be associated with the measured quantities, and (ii) an introduction to the concept of screening (assessing) observations prior to their reduction and use in network computations. Chapters 3 and 4 respectively cover the topics of Mathematical Models for Azimuth, Direction and Angle Observations and Mathematical Models for Distance Observations. For each observable, the inverse explicit model is presented, first in it's original non-linear form, then in it's linearized form. The linearized equation is often
referred to as the observation equation. The relationships of the elements of the linearized model with the matrix expressions of the method of leastsquares used for solving a set (2 or more) of these equations is given for each case. In Chapter 5, entitled the Solution of Unique Cases, the linearized mathematical models of Chapters 3 and 4 are taken in several practical combinations to yield unique (no. unknowns $u=n o$. observations $n$ ) solutions of well-known surveying problems (e.g. direct problem, azimuth intersection, distance intersection, resection, special traverses). Numerical examples for each problem are presented. The Solution of Overdetermined Cases, which constitute the main body of work in the establishment of a surveying network, is the subject of Chapters 6 and 7. The advantages of using the method of least squares is given; the combination of the observation equations (Chapters 3 and 4) in several practical situations are shown. In addition, the implications of differenct conditions imposed on a horizontal network (e.g. fixed or weighted coordinates, orientation, scale) are discussed. Numerical examples for several types of survey networks are given (e.g. traverse, triangulation, trilateration). Chapters 8 and 9 deal with the analyses of networks. The Preanalysis of a network, which is basically an optimization process, is important for surveyors when considering geometric design, economics, tolerances, etc. The Postanalysis, treated in the final chapter (9) is most important to a surveyor. It is here that a certain "confidence" in the work done can be ascertained. The manual is concluded by Three Appendices deemed to be necessary for a complete understanding of this work, namely Taylor Series (I), Least Squares Method (Parametric) (II), and Error Ellipses (III).

## 2. OBSERVATIONS AND THEIR REDUCTIONS

The planning, execution, and treatment of observed azimuths, directions, angles, and distances are important aspects of the establishment and analysis of a horizontal geodetic network. The execution (actual field measurement techniques) are not covered in this manual; the interested reader is referred to, for example, Faig [1972], Thomson [1978], Cooper [1971], Burnside [1971], and Saastamoinen [1967]. The planning of observations is treated in Chapter 8 (Preanalysis). The mathematical treatment of the observed quantities is given here in three sections, namely (i) the accuracies of observed azimuths, directions, angles, and distances, (ii) a review of the reductions of observations to a conformal mapping plane, and (iii) data screening.

### 2.1 Accuracies of Observed Azimuths, Directions, Angles and Distances

A knowledge of the accuracy of an observable (a proposed measurement) is an important aspect of the preanalysis of survey networks (Chapter 8), and a knowledge of the accuracy of an observation (a completed measurement) is important for network computations. The determination of these accuracies, expressed as variances ( $\sigma_{\ell}{ }^{2}$ ), is the subject of this section. Note that the effects of systematic errors are assumed to have been removed by either observing or mathematical procedures or a combination of the two.

The variance $\sigma_{A}^{2}$ of an astronomic azimuth determination by observation of celestial bodies (e.g. stars or sun) is dependent on the method used. For astronomic azimuths determined by the hour angle method $\sigma_{A}{ }^{2}$ is given as [e.g. Nickerson, 1978; Mueller, 1969]

$$
\begin{equation*}
\sigma_{A}^{2}=\frac{1}{n} F \sigma_{t}^{2}+\frac{1}{n}\left(\sigma_{p}^{2}+\sigma_{c}^{2}\right) \tag{2-1}
\end{equation*}
$$

where $\quad n=$ number of pointings on the star,

$$
\begin{aligned}
\sigma_{t}^{2}= & \text { variance of the observation of time in arcseconds }\left(l^{\prime \prime}=0.067 s\right) \\
\sigma_{p}^{2}= & \text { variance of a single pointing on a star (cf. eq. }(2-5)), \\
\sigma_{c}^{2}= & \text { combined variance of two readings of the horizontal circle and } \\
& \text { pointing on the reference mark (cf. eqs. (2-6) and (2-4)), }
\end{aligned}
$$

$$
\begin{aligned}
& F=\cos ^{2} \Phi(\tan \Phi-\cos A \cot Z)^{2}+m\left(2 \tan ^{2} \Phi+\cot ^{2} Z-2 \tan \Phi \cos A \cot Z\right) \\
& Z=\text { zenith angle of star, } \\
& \Phi=\text { astronomic latitude of station, } \\
& A=\text { measured astronomic azimuth, } \\
& m=\left(\sigma_{p}^{2}+\sigma_{v}^{2}\right) / \sigma_{t}^{2} . \\
& \sigma_{V}^{2}=\text { variance of levelling the theodolite (cf. eq. (2-9)). }
\end{aligned}
$$

Some typical default values for $\sigma_{A}$ assuming $\Phi=A=Z=45^{\circ}$ and different typical theodolites are shown in Table 2.1.

Astronomic azimuths determined by observing star altitude have
the following expected variance [e.g. Nickerson, 1978; Mueller, 1969]:

$$
\begin{align*}
\sigma_{A}^{2}= & \frac{1}{n}\left\{\left(\sigma_{p}^{2}+\sigma_{v}^{2}\right) \tan ^{2} a+(\tan \Phi-\cos A \tan a)^{2}\left[\left(\sigma_{v c}^{2}+\sigma_{p}^{2}\right) \operatorname{cosec}^{2} A+\sigma_{\operatorname{tr}}^{2} \cos ^{2} \Phi\right]\right. \\
& \left.+\left(\sigma_{p}^{2}+\sigma_{c}^{2}\right)\right\} \tag{2-2}
\end{align*}
$$

where $\sigma_{v c}^{2}=$ combined variance of levelling the vertical circle bubble and reading the vertical circle,
$\sigma_{t r}^{2}=$ variance of tracking for simultaneous horizontal and vertical pointing on a star $\simeq 1^{\prime \prime}$,
$a=$ altitude of the observed star corrected for refraction.

|  | Typical 20" Inst. |  |  |  | Typical 1 " Inst. |  |  |  | Typical 0:5 inst. |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\mathrm{M}=20, \mathrm{~d}=20^{\prime \prime}, \mathrm{v}=30^{\prime \prime}$ |  |  |  | $\mathrm{M}=30, \mathrm{~d}=1, \mathrm{v}=20^{\prime \prime}$ |  |  |  | $M=40, d=0.5, V=10^{\prime \prime}$ |  |  |  |
|  | yields |  |  |  | yields |  |  |  | yields |  |  |  |
|  | $\sigma_{p}=3.5, \sigma_{c}=8 " .78, \sigma_{v}=6 "$ |  |  |  | $\sigma_{p}=2: 33, \sigma_{c}=3.24, \sigma_{v}=4 "$ |  |  |  | $\sigma_{p}=1: 88, \sigma_{c}=2.095, \sigma_{v}=2 "$ |  |  |  |
|  | $\mathrm{n}=2$ | $\mathrm{n}=4$ | $\mathrm{n}=8$ | $\mathrm{n}=16$ | $\mathrm{n}=2$ | $\mathrm{n}=1$ | $\mathrm{n}=8$ | $\mathrm{n}=16$ | $\mathrm{n}=2$ | $\mathrm{n}=4$ | $\mathrm{n}=8$ | $\mathrm{n}=16$ |
| $\sigma_{t}=0.5(0.03 s)$ | 9.11 | 6. 44 | 4.55 | 3. 22 | 5.20 | 3.68 | 2.60 | 1:84 | 3.15 | 2. 23 | 1:58 | 1:11 |
| $\sigma_{t}=1.5(0.10 s)$ | 9:11 | 6.44 | 4:55 | 3. 22 | 5.21 | 3.68 | 2:60 | 1:84 | 3.16 | 2"23 | 1:58 | 1.12 |
| $\sigma_{t}=15 \prime$ (1.0s) | 9:37 | 6:'62 | 4:38 | 3.31 | 5.65 | 3:99 | 2:82 | 2.00 | 3.84 | 2:72 | 1:92 | 1:36 |
| $\sigma_{t}=2^{\prime} 30^{\prime \prime}(10.0 \mathrm{~s})$ | 23:78 | 16:81 | 11:89 | 8:41 | 22:57 | 15:96 | 11:29 | 7:98 | 22.19 | 15:69 | 11:10 | 7:85 |

Table 2.1. Expected Values of $\sigma_{A}$ Using Hour Angle Method for $\Phi=A=Z=45^{\circ}$

Table 2.2 shows some typical values of $\sigma_{A}$ (for the same types of theodolites considered in Table 2.1) assuming $A=\Phi=45^{\circ}$.

For azimuths observed with gyro-theodolites, the most reliable method for obtaining a variance is to compute the sample variance of the mean of a set of many observation of the azimuth. This yields

$$
\begin{equation*}
\sigma_{A}^{2}=\frac{1}{n-1} \sum_{i=1}\left(A_{i}-\bar{A}\right) \tag{2-3}
\end{equation*}
$$

where $n=$ number of observed gyro-azimuths,
$A_{i}=$ individual gyro-azimuths,
$\bar{A}=$ mean of the set of observed gyro-azimuths.
Some default values to be expected are $\sigma_{A}=20^{\prime \prime}$ to $30^{\prime \prime}$ for a single observed azimuth determined by a gyro attachment similar to the Wild GAKl
[Bomford, 1975], and $\sigma_{A}=3^{\prime \prime}$ for a single gyro-azimuth observed with gyrotheodolites such as the MOM Gi-B2 or GYMO-GI-Bl/A which have electronic time registration [Halmos, 1977].

The expected variance of direction observations is [e.g Nickerson, 1978]

$$
\begin{equation*}
\sigma_{d}^{2}=\frac{\sigma_{p}^{2}+\sigma_{r}^{2}}{n}+\sigma_{L}^{2}+\rho^{2} \frac{2 \sigma_{c}^{2}}{D^{2}} \tag{2-4}
\end{equation*}
$$

where $\sigma_{p}^{2}=$ variance of pointing the telescope on the target (cf. eqs. (2-6) and (2-5)),
$\sigma_{r}^{2}=$ variance of reading the horizontal circle of the theodolite (cfe eqs. (2-7) and (2-8)),
$\sigma_{L}^{2}=$ effect of variance of levelling the theodolite (cf. eq. (2-9)),
$\sigma_{c}^{2}=$ variance of centering the instrument and target (see Table 2.3),
$n$ number of pointings and readings for the direction,
$\rho=206264.8=$ number of arc seconds in one radian,
D = distance between instrument and target.

Typical 20" Inst.
Typical 1" Inst.
Typical 0."5 Inst.


Table 2.2. Expected Values of $\sigma_{A}$ Using Star Altitude Method for $\Phi=A=45^{\circ}$

The pointing error is dependent on the magnification $M$ of the particular theodolite being used (see Table 2.5) and is given as

$$
\begin{equation*}
\sigma_{p}^{\prime \prime} \simeq \frac{45^{\prime \prime}}{M} \tag{2-5}
\end{equation*}
$$

for stationary targets and good observing conditions. For moving targets (e.g. star), the pointing error is

$$
\begin{equation*}
\sigma_{p}^{\prime \prime}=\frac{70^{\prime \prime}}{M} \tag{2-6}
\end{equation*}
$$

The reading error is a function of the least count of the theodolite and the readout system. For theodolites with a least count of $d$ ", and using coincidence micrometers (usually the case for 1 " and 0.5 instruments (see Table 2.5)),

$$
\begin{equation*}
\sigma_{r}=2.5 \mathrm{~d}^{\prime \prime} \tag{2-7}
\end{equation*}
$$

and for theodolites with a microscope or direct reading system (typically for $d=10^{\prime \prime}$ to $l^{\prime \prime}$ ),

$$
\begin{equation*}
\sigma_{r}=0.3 \mathrm{~d}^{\prime \prime} \tag{2-8}
\end{equation*}
$$

The effect $\sigma_{L}^{2}$ of the variance of levelling the instrument $\sigma_{v}^{2}$ is dependent on the vertical angle $h$ to the target, and is given as

$$
\begin{equation*}
\sigma_{L}=\sigma_{v}^{\prime \prime} \tan h \tag{2-9}
\end{equation*}
$$

where

$$
\begin{equation*}
\sigma_{v}=0.2 \mathrm{v}^{\prime \prime} \tag{2-10}
\end{equation*}
$$

and $v^{\prime \prime}$ is the value of one division ( 2 mm ) of the: plate level (see Table 2.5). The centering error $\sigma_{c}$ summarised in

| Method of Centering | Expected Error $\sigma_{C}$ |
| :--- | :--- |
| spring plumb-bob | $1 \mathrm{~mm} / \mathrm{m}$ |
| optical plummet | $0.5 \mathrm{~mm} / \mathrm{m}$ |
| plumbing rods | $0.5 \mathrm{~mm} / \mathrm{m}$ |
| forced or self-centering | 0.1 mm |

Table 2.3 Expected Centering Error

Table 2.3 is for normal conditions (i.e no wind, equipment in good adjustment). Table 2.4 lists typical default values for $\sigma_{d}$ assuming $\sigma_{c}=0.5 \mathrm{~mm} / \mathrm{m}$ (i.e. $\sigma_{c}=1 \mathrm{~mm}$ for instrument height $=2 \mathrm{~m}$ ) and a vertical angle $h=5^{\circ}$. It should be noted that this table is representative of good observing conditions. If observing conditions were poor, then the pointing error $\sigma_{p}$ would increase accordingly. From equation (2-4), it is obvious that centering error $\sigma_{c}$ is more critical for short lines of sight. As well, $\sigma_{L}$ (eq. 2-9) will contribute increasingly for steeper lines of sight.

Horizontal angles $B$ can be considered as the difference of two direction observations. Propagation of errors through the formula

$$
\begin{equation*}
\mathrm{B}=\mathrm{d}_{2}-\mathrm{d}_{1} \tag{2-11}
\end{equation*}
$$

yields the expected variance for an observed angle as

$$
\begin{equation*}
\sigma_{B}^{2}=2\left\{\frac{\sigma_{p}^{2}+\sigma_{r}^{2}}{n}+\sigma_{L}^{2}+\rho^{2} \frac{2 \sigma_{c}^{2}}{D^{2}}\right\} \tag{2-12}
\end{equation*}
$$

or twice the variance of a single direction. Thus, typical values for observed angles can be obtained from Table 2.4 by multiplying the values by $\sqrt{2}$.

|  | Typical 20" Instrument$\begin{gathered} M=20, d=20 ", v=30 " \\ \therefore \\ \sigma_{p}=2.25, \sigma_{r}=6!0, \sigma_{L}=0.52 \end{gathered}$ |  |  |  | Typical 1" Instrument$\begin{gathered} M=30, d=1 ", v=20 " \\ \therefore \\ \sigma_{p}=1: 5, \sigma_{r}=2: 5, \sigma_{L}=0: 35 \end{gathered}$ |  |  |  | Typical 0:5 Instrument$\begin{gathered} M=40, d=0: 5, \quad v=10 " \\ \therefore \\ \sigma_{p}=1: 13, \sigma_{r}=1: 25, \sigma_{L}=0.18 \end{gathered}$ |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $D(m)$ | $\mathrm{n}=2$ | $\mathrm{n}=4$ | $\mathrm{n}=8$ | $\mathrm{n}=16$ | $\mathrm{n}=2$ | $\mathrm{n}=4$ | $\mathrm{n}=8$ | $\mathrm{n}=16$ | $\mathrm{n}=2$ | $\mathrm{n}=4$ | $\mathrm{n}=8$ | $\mathrm{n}=16$ |
| 100 | 5.41 | 4.36 | 3.73 | 3. 37 | 3.59 | 3"28 | 3:11 | 3.03 | 3:16 | 3:04 | 2!98 | 2:95 |
| 200 | 4.79 | 3.56 | 2.74 | 2". 23 | 2:55 | 2.09 | 1.88 | 1.67 | 1:89 | 1:69 | 1:59 | 1.53 |
| 400 | 4:62 | 3. 33 | 2. 44 | 1:84 | 2.21 | $1: 67$ | 1:31 | 1.09 | 1:41 | 1:13 | 0.96 | 0.86 |
| 800 | 4.58 | 3'27 | 2. 35 | 1.72 | 2.12 | 1:54 | 1:15 | 0.89 | 1:26 | 0.94 | 0. 72 | 0. 59 |
| 1600 | 4.56 | 3. 25 | 2. 33 | 1.69 | 2.10 | 1.51 | 1.10 | 0.83 | 1:22 | 0.88 | 0.65 | 0.49 |
| 3200 | 4:56 | 3.25 | 2.33 | 1:69 | 2.09 | 1:50 | 1.09 | 0:81 | 1:21 | 0:87 | 0.63 | 0.47 |

Table 2.4. Expected Values of $\sigma_{d}$ for $h=5^{\circ}$ and $\sigma_{c}=1 \mathrm{~mm}$


Table 2.5 Major Features of Some Modern Theodolites

The covariance between angles derived from a set of three or more directions cannot be overlooked. Considering the situation illustrated in Figure 2.1, the angles are usually derived from the directions as

$$
\begin{align*}
& B_{i j k}=d_{i k}-d_{i j}  \tag{2-13}\\
& B_{i k \ell}=d_{i \ell}-d_{i k}
\end{align*}
$$



Figure 2.1 Angles and Directions

Use of the covariance law to propagate errors from equations (2-13) into the angles $B_{i j k}$ and $B_{i k \ell}$ gives the variance covariance matrix $C_{B}$ of the angles as

$$
C_{B}=\left[\begin{array}{ccc}
\sigma_{d_{i j}}^{2} & +\sigma_{d i k}^{2} & -\sigma_{d_{i k}}^{2}  \tag{2-14}\\
-\sigma_{d_{i k}}^{2} & & \sigma_{d_{i k}}^{2}+\sigma^{2} d_{i l}
\end{array}\right]
$$

i.e. covariances equal to minus the variance of the common direction between the angles will exist. For angles not derived from directions, but measured independently, $C_{B}$ will be a diagonal matrix.

The variance $\sigma_{r}^{2}$ of spatial distances observed with EDM is characterized by [e.g. Nickerson, 1978]

$$
\begin{equation*}
\sigma_{r}^{2}=\frac{\sigma_{p h}^{2}}{m_{l}}+\sigma_{z}^{2}+\frac{r_{n}^{2} \sigma_{n}^{2}}{m_{2} n^{2}} \tag{2-15}
\end{equation*}
$$

where $\sigma_{\mathrm{ph}}^{2}=$ variance of phase difference determination (cf. eq. (2-16)), $\sigma_{z}^{2}=$ variance of the so-called zero error, $\sigma_{n}^{2}=$ variance of determination of the index of refraction $n$ (cf. eqs. (2-17) and (2-18)),
$m_{1}$ and $m_{2}=$ number of determinations of phase difference and meteorological readings, respectively. The variance of phase determination is computed as

$$
\begin{equation*}
\left.\sigma_{\mathrm{ph}}^{2}=\left\{\frac{1}{2} \lambda\right\}\right\}^{2} \sigma_{\theta}^{2} \tag{2-16}
\end{equation*}
$$

where $\lambda=$ modulation wavelength used by the specific
instrument (see Table 2.6),
$\sigma_{\theta}^{2}=$ variance of determination of the phase difference for one distance measurement in fractions of a wavelength.

Most modern EDM equipment can easily achieve an accuracy of phase difference determination $\sigma_{\theta}=0.001$ [Burnside, 1971], but more accurate values for an individual instrument should be available from the manufacturer's specifications. The zero error $\sigma_{z}$ results from inaccurate knowledge of the electrical center of the instrument with respect to the geometric center which is aligned over the point. This value is usually small (e.g. 5 mm ) for instruments using light waves as the carrier frequency, but for

| Model | Manufacturer | Radiation Source | Modulation Frequency |  | Modulator | Power <br> Consumed (W) | Method of Phase Measurement | Range ( Km ) | Standard Deviation |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Base ( $\mathrm{MH}_{\mathrm{Z}}$ ) | Total \# |  |  |  | Day Night |  |
| Geodimeter | AGA | $5 \mathrm{ml} \mathrm{He}-\mathrm{Ne}$ | 30 | 4 | KDP Crystal | 75 | null meter | $30 \quad 60$ | $\pm\left(5 \mathrm{~mm}+1 \cdot 10^{-6} \mathrm{~s}\right)$ |
| Model 8 Geodolite 3 G | Sweden Spectra-Physics |  |  |  |  |  | digital | 6080 | $\pm 1.10^{-6} \mathrm{~s}$ or 1 mm |
| Geodolite 3 G | Spectra-Physics U.S.A. | $5 \mathrm{~mW} \mathrm{He}-\mathrm{Ne}$ Laser | 49 | 5 |  | . 400 | digital |  | whichever greater |
| Geodimeter | AGA | 30 W | 30 | 3 | Kerr Cell | 70 | resolver | $3 \quad 15$ | $\pm\left(1 \mathrm{~cm}+2.10^{-6 \mathrm{~s}}\right)$ |
| Model 6 | Siveden | Nercury Lamp |  |  |  | 300 | null meter | ${ }_{3} \quad^{25}$ | $+\left(1 \mathrm{~cm}+1.10^{-6} \mathrm{~s}\right)$ |
| Geodimeter 76 | AGA Sweden | $\begin{aligned} & 2 \mathrm{~mW} \\ & \text { Laser } \end{aligned}$ |  | 2 | Kerr Cell |  |  | (3 prisms) | $\pm\left(1 \mathrm{~cm}+1.10^{-6} \mathrm{~s}\right)$ |
| DM 1000 | Kern | GaAs-Diode | 15 | 2 | - . | 11 | digital | 2.5 | $\pm 1 \mathrm{~cm}$ |
| Mekometer | Kern | 900 nm Xenon-flash | 500 | 5 | ADP Crystal | 18 | optomechanical | 3 Refl.) 3 | $\pm\left(0.2 \mathrm{~mm}+1.10^{-6} \mathrm{~s}\right)$ |
| ME 3000 | Kern | ( 100 Hz ) |  |  | ADP Crystar |  | null meter | (3 prisms) |  |
| Di 500 | Kern | GaAs Diode | 15 | 2 | - | 11 | digital | 0.5 | $\pm 1 \mathrm{~cm}$ |
| S4 11 | Zeiss | 875 nm GaAs Diode | 15 | 2 | - | 12 | automatic | (3 prisms) | $\pm 5$ to 10 mm |
|  | Oberkochen | 910 nm |  |  |  |  | digital | (19 prisms) |  |
| ELDI 2 | Zeiss |  |  | . |  | 4 |  |  | $\pm 5 \mathrm{~mm}$ |
| A 100 | Oberkochen Tellurometer | GaAs Diode | 75 | 4 | - | 14 | digital | 2 | $\pm\left(1.5 \mathrm{~mm}+2.10^{-6} \mathrm{~s}\right)$ |
|  |  | 930 nm |  |  |  |  |  |  | $+\left(5 \mathrm{~mm}+5.10^{-6} \mathrm{~s}\right)$ |
| CD 6 | Tellurometer | GaAs Diode |  |  |  |  | digital | 2 | $\stackrel{+}{+}\left(5 \mathrm{~mm}+5.10{ }^{\text {cm }}\right.$ |
| 50:4-3 | Sokkisha Ltd, Tokyo | GaAs Diode 900 nm | 15 | 2 | - | 10 | digital | (3 ${ }_{\text {prisms }}$ | $\pm 1 \mathrm{~cm}$ |
| 013 | Wild Hearbrugg | GaAs Diode | 7.5 | 2 | - | 14 | digital | 0.6 | $\pm\left(5 \mathrm{~mm}+5.10^{-6} \mathrm{~s}\right)$ |
| 2il-60 | Cuhic Ind. | 375 nm GaAs Diode | 75 | 3 | - | 15 | automatic | (3 prisms) | $\pm\left(5 \mathrm{~mm}+1.10^{-5} \mathrm{~s}\right)$ |
| Cubitape | Co., LISA | 9 O nm |  |  |  |  | digital |  | $\pm\left(5 \mathrm{~mm}+1.10^{-5} \mathrm{~s}\right)$ |
| 3300 B | Hewlett- Packard, USA | GaAs Diode | 15 | 4 | - | 12 | digital null meter | (3 prisms) | $\pm\left(5 \mathrm{~mm}+1.10^{-3} \mathrm{~s}\right)$ |
| Ranger II | Laser Syst. \& E?ectronics USA | $3 \mathrm{~m}!: \mathrm{He}-\mathrm{Ne}$ Laser | 15 | 4 | KDP Crystal |  | automatic digital | ${ }_{6}$ | $\pm\left(5 \mathrm{~mm}+2.10^{-5} \mathrm{~s}\right)$ |

Table 2.6 Characteristics of Modern EDM

| Model | Manufacturer | Carrier Frequency $\left(\mathrm{GH}_{z}\right)$ | Measuring Frequency $\left(\mathrm{MH}_{\mathrm{z}}\right)$ | Antenna |  | Power Consumed (w) | Readout | Measuring <br> Range (Km) | Standard Deviation |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | Diameter (cm) | Divergence ( ${ }^{\circ}$ ) |  |  |  |  |
| MRA 101 | Tellurometer Ltd. | 10.05 to 10.45 | 7.5 | 33 | 6 | 38 | digital | 0.1 to 50 | $\pm\left(1.5 \mathrm{~cm}+3.10^{-6} \mathrm{~s}\right)$ |
| MRA 3 | Tellurometer Ltd. | 10.025 to 10.45 | 7.5 | 33 | 9 |  | digital | 0.1 to 50 | $\pm\left(1.5 \mathrm{~cm}+3.10^{-6} \mathrm{~s}\right)$ |
| MRA 4 | Tellurometer Ltd. | 34.5 to 35.1 | 75 | 33 | 2 |  | digital | 0.05 to 60 | $\pm\left(3 \mathrm{~mm}+3.10^{-6} \mathrm{~s}\right)$ |
| CA 1000 | Tellurometer Ltd. | 10.1 to 10.45 | 19 to 25 |  |  |  | digital | 0.05 to 30 | $\pm\left(1.5 \mathrm{~cm}+5.10^{-6} \mathrm{~s}\right)$ |
| Electrotape DM2O | Cubic Corp. U.S.A. | 10.5 to 10.5 | 7.5 | 33 | 6 |  | digital | 0.05 to 50 | $\pm\left(1 \mathrm{~cm}+3.10^{-6} \mathrm{~s}\right)$ |
| Distomat DI50 | Wild Heerbrugg | 10.2 to 10.5 | 15 | 36 | 6 | 50 | digital | 0.1 to 50 | $\pm\left(2 \mathrm{~cm}+5.10^{-6} \mathrm{~s}\right)$ |
| Distomat DI60 | Siemens-Albiswerk | 10.3 | 150 | 35 | 6 | 38 | digital | 0.02 to 150 | $\pm\left(1 \mathrm{~cm}+3.10^{-6} \mathrm{~s}\right)$ |

Table 2.6 Continued.
microwave instruments the value may be up to 20 mm . This value is normally supplied with the instrument. The variance of the refractive index is different for lightwaves and microwaves. For lightwaves, the variance is [e.g. Nickerson, 1978; Laurila, 1976]

$$
\sigma_{n}^{2}=\left[\left\{\frac{1}{T^{2}}\left(\frac{-N_{G} P}{3.709}+11.27 e\right)\right\}^{2} \sigma_{T}^{2}+\left\{\frac{N_{G}}{3.709 T}\right\}^{2} \sigma_{p}^{2}+\left\{\frac{11.27}{T}\right\}^{2} \sigma_{e}^{2}\right] \cdot 10^{-12}
$$

where $T=$ temperature in degrees Kelvin ( $t^{\circ} \mathrm{C}+273.15$ ),

$$
\begin{aligned}
& \mathrm{N}_{\mathrm{G}}=\left(287.604+\frac{4.8864}{\lambda_{\mathrm{C}}^{2}}+\frac{0.068}{\lambda_{\mathrm{C}}^{4}}\right) \text { for } \lambda_{\mathrm{C}}=\underset{\text { carrier wavelength }}{\text { (see Table } 2.6),} \\
& \mathrm{p}=\text { air pressure in millibars ( } 1 \text { mbar } \simeq 0.75 \mathrm{~mm} \mathrm{Hg} @ 0^{\circ} \mathrm{C} \text { ), } \\
& \mathrm{e}=\text { water vapour pressure in mbar (for detailed computation see Bomford }
\end{aligned}
$$ [1975], p. 54),

$\sigma_{T} \stackrel{2}{2}=$ variance of temperature measurement in ${ }^{\circ} C^{2}$,
$\sigma_{p}^{2}=$ variance of pressure measurements in mbar ${ }^{2}$,
$\sigma_{e}^{2}=$ variance of water vapour determination in mbar ${ }^{2}$.
For microwaves the variance of the refractive index is [e.g. Nickerson, 1978; Laurila, 1976]

$$
\begin{array}{r}
\sigma_{n}^{2}=\left[\left\{\frac{-77.62 P}{T^{2}}+\left(\frac{12.92}{T^{2}}-\frac{74.38 \cdot 10^{4}}{T^{3}}\right) e\right\}^{2} \sigma_{T}^{2}+\left\{\frac{77.62}{T}\right\}^{2} \sigma_{p}^{2}+\right. \\
\left.\quad+\left\{\frac{-12.92}{T}+\frac{37.19 \cdot 10^{4}}{T^{2}}\right\} \quad \sigma_{e}^{2}\right] \cdot 10^{-12} \tag{2-18}
\end{array}
$$

where the elements in this equation are defined the same as those in equation (2-17). Table 2.7 summarizes the effect of errors in meteorological measurements on observed distances.

| METEOROLOGICAL ERROR | EFFECT ON DISTANCE |  |
| :--- | :---: | :---: |
|  | Light waves | Microwaves |
| $\pm 1$ mbar in air pressure | 0.22 ppm | 0.22 ppm |
| $\pm 1^{\circ} \mathrm{C}$ in temperature | 1.0 ppm | 1.6 ppm |
| $\pm 1^{\circ} \mathrm{C}$ in the difference |  |  |
| between dry and wet bulbs |  |  |

Table 2.7 Effect of Meteorological Errors on Measured Distances
. Table 2.8 below lists some expected values of $\sigma_{r}$ for both lightwave and microwave instruments.

For treatment of distancesobserved by mechanical or optical means, one is referred to e.g. Nickerson [1978] or Smith [1970].

The above discussion has treated only the accuracy of observed azimuths, directions, angles and distances. However, the observations used in horizontal network computations are considered to be reduced to the plane. Any inaccuracies resulting from these reductions must also be accounted for. This propagation of errors through the reduction formulae (see section 2.2) has already been covered in section 3.2.8 of Thomson et al [1978].

### 2.2 Reduction of Observations to a Conformal Mapping Plane

The reduction of observed azimuths, directions, angles, and distances to a conformal mapping plane is essentially a two-phased process: terrain to reference ellipsoid, and reference ellipsoid to conformal mapping plane. Each phase, depending on the observed quantity, may contain one or more reduction steps. These procedures, for the. $3^{\circ}$ Transverse Mercator and

|  | Lightwaves; $\lambda_{c}=900 \mathrm{~nm}$$\lambda=20 \mathrm{~m}, \sigma_{z}=0.005 \mathrm{~m}, \sigma_{\theta}=0.001$ |  |  |  | Microwaves; $\lambda_{c}=3 \mathrm{~cm}$ $\lambda=40 \mathrm{~m}, \sigma_{z}=0.015 \mathrm{~m}, \sigma_{\theta}=0.001$ |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\begin{gathered} \sigma_{p}=1 \text { mbar } \\ \sigma_{T}=\sigma_{\Delta T}=0.2{ }^{\circ} \mathrm{C} \end{gathered}$ |  | $\begin{aligned} & \sigma_{p}=5 \mathrm{mbar} \\ & \sigma_{T}=\sigma_{\Delta T}=1^{\circ} \mathrm{C} \end{aligned}$ |  | $\begin{gathered} \sigma_{p}=1 \text { mbar } \\ \sigma_{T}^{2}=\sigma_{\Delta T}=0: 2 C \end{gathered}$ |  | $\begin{gathered} \sigma_{p}=5 \mathrm{mbar} \\ \sigma_{T}=\sigma_{\Delta T}=1^{\circ} \mathrm{C} \end{gathered}$ |  |
| S (m) | $m_{1}=2, m_{2}=1$ | $\mathrm{m}_{1}=4, \mathrm{~m}_{2}=2$ | $m_{1}=2, m_{2}=1$ | $\mathrm{m}_{1}=4, \mathrm{~m}_{2}=2$ | $\mathrm{m}_{1}=2, \mathrm{~m}_{2}=1$ | $\mathrm{m}_{1}=4, \mathrm{~m}_{2}=2$ | $m_{1}=2, m_{2}=1$ | $\mathrm{m}_{1}=4, \mathrm{~m}_{2}=2$ |
| 100 | 0.009 | 0.007 | 0.009 | 0.007 | 0.021 | 0.018 | 0.021 | 0.018 |
| 200 | 0.009 | 0.007 | 0.009 | 0.007 | 0.021 | 0.018 | 0.021 | 0.018 |
| 400 | 0.009 | 0.007 | 0.009 | 0.007 | 0.021 | 0.018 | 0.021 | 0.018 |
| 800 | 0.009 | 0.007 | 0.009 | 0.007 | 0.021 | 0.018 | 0.022 | 0.019 |
| 1600 | 0.009 | 0.007 | 0.009 | 0.007 | 0.021 | 0.018 | 0.024 | 0.020 |
| 3200 | 0.009 | 0.007 | 0.010 | 0.008 | 0.021 | 0.018 | 0.033 | 0.026 |
| 6400 | 0.009 | 0.007 | 0.013 | 0.010 | 0.023 | 0.020 | 0.057 | 0.041 |
| 12800 | 0.010 | 0.008 | 0.021 | 0.015 | 0.029 | 0.023 | 0.107 | 0.077 |
| 25600 | 0.012 | 0.009 | 0.040 | 0.028 | 0.045 | 0.035 | 0.212 | 0.150 |

Table 2.8 Expected Values for $\sigma_{r}$

Double Stereographic conformal map projections, are given, for example, in Thomson et. al. [1978]. The entire process is reviewed here in the context of horizontal geodetic networks. The primary reason for this is that for network computations the sequence of events is different than that used for position computations; in addition, the software used to generate the numerical examples given in this report follows the sequence given here.

The first problem to be solved is the determination of the approximate coordinates, $\left(X^{a}, Y^{a}\right)$ and $\left(\phi^{a}, \lambda^{a}\right)$, for each unknown point in the network. This can be done in several ways, but the most often used are (i) to determine them graphically using a large scale map or a plan, or (ii) to compute them using observed quantities, well known geometric/ trigonometric solutions and coordinate transformation procedures. The coordinate transformations are given in, for example, Krakiwsky, et. al [1977]. The main point to bear in mind when determining approximate coordinates is that they must be sufficiently close to the final values so that the effects on the reduction of observations will be negligible. A conservative estimate of "sufficiently close" is 20 m . This can be achieved easily, in most instances, using observed quantities and unique geometric/trigonometric solutions. Well determined approximate coordinates are also important in the solution for final coordinates as this will minimize the number of required iterations [e.g. Steeves, 1978].

The reduction of an astronomic azimuth ( $A_{i j}$ ), obtained from astronomic observations or a gyrotheodolite, to a conformal mapping plane (grid) azimuth ( $t_{i j}$ ) is outlined in Figure 2.2. The observed, known, and
computed quantities required are given in Table 2.9. The equation numbers listed in Figure 2.2 and Table 2.9 refer to those found in Thomson et. al [1978].

The reduction of a measured terrain normal section direction ( $d_{i j}^{t}$ ) to a conformal mapping plane (grid) direction $\left(d_{i j}\right)$ of the corresponding chord is outlined in Figure 2.3. The observed, known, and computed quantities required are given in Table 2.10. The equation numbers listed in Figure 2.3 and Table 2.10 refer to those found in Thomson et. al. [1978].

A measured angle ( $\mathrm{B}_{\mathrm{jik}}$ ), since it is simply the difference of two terrain normal section directions ( $d_{i k}^{t}-d_{i j}^{t}$ ) follows the same reduction procedure as the directions themselves. The procedure is outlined in Figure 2.4. The observed, known, and computed quantities involved are given in Table 2.11. The equation numbers listed in Figure 2.4 and Table 2.11 refer to those found in Thomson et. al. [1978].

The reduction of a terrain spatial distance ( $r_{i j}$ ) (measured distance corrected for atmospheric and instrumental effects) to a conformal mapping plane (grid) distance ( $\ell_{i j}$ ) of the corresponding chord is outlined in Figure 2.5. The observed, known, and computed quantities involved are given in Table 2.12. The equation numbers given in Figure 2.5 and Table 2.12 refer to those found in Thomson et. al [1978].

An examination of Figures 2.2 to 2.5 and Tables 2.9 to 2.12 inclusive shows a significant overlap in observed (e.g. $Z_{i j}$ ), computed (e.g. approximate coordinates), and known (e.g. ellipsoidal and conformal mapping system constants) quantities involved in the reduction of measured azimuths, directions, angles, and distances. In practice, these quantities need only be specified once. For example, in the program GEOPAN [Steeves, 1978],


Figure 2.2
Reduction of Observed Astronomic Azimuth to a Conformal Mapping Plane

| Status | Quantity | Remarks |
| :---: | :---: | :---: |
| ® 0 d 0 0 0 0 | ```Aij: Astronomic Azimuth Z ij (or hij or Hij) zenith distance(or elliposidal or orthometric height difference)``` | ōbtained via stellar, solar, or gyrotheodolite observations. <br> needed for reduction purposes |
| $\begin{aligned} & 0 \\ & 0 \\ & 0 \\ & 0 \\ & 0 \\ & 0 \\ & 0 \\ & 0 \end{aligned}$ | $\begin{aligned} & \left(x_{i}^{a}, y_{i}^{a}\right),\left(x_{j}^{a}, y_{j}^{a}\right) \\ & \left(\phi_{i}^{a}, \lambda_{i}^{a}\right),\left(\phi_{j}^{a}, \lambda_{j}^{a}\right) \\ & \left(x_{i}^{a}, \Lambda_{i}^{a}\right) \\ & z_{i, j}: \text { reduced zenith distance } \\ & S_{i j}^{a}: \text { approximate ellipsoidal }_{\text {distance }} \\ & (T-t)_{i j} \\ & \gamma_{i}: \text { Meridian convergence } \\ & h_{j}: \text { ellipsoidal height of } \\ & \text { target } \end{aligned}$ | see Krakiwsky et al. [1977] re coordinate transformations. <br> (3-14); use partially reduced azimuth, from (3-15), for this computation (3-39) or (3-62); use approximate coordinates ( $\sigma^{a}, \lambda^{a}$ ) to compute all quantities, including auxilaries. (4-11) or (4-31) for Double Stereographic, (4-53) for $3^{\circ}$ Transverse Mercator; use approximate coordinates for all computations. (4-7) or (4-27) for Double Stereographic, (4-47) for $3^{\circ}$ Tranverse Mercator; use approximate coordinates for all computations. |
|  | ```\(a, b\) (or \(a, f\) ) \(\psi_{0}, \Lambda_{o}, R ;\) \(\phi_{0}, \lambda_{0} ; x_{o}, y_{0} ;\) \({ }^{k}\) 。 \(H_{i}\left(\right.\) or \(\left.h_{i}\right)\) : orthonetric \({ }^{i}\) (or elfipsoid) height \(N_{i}^{*}, N_{j}^{*}\) \(\xi_{i}, \eta_{i}\) \{deflection of vertical components at observed station``` | parameters of reference ellipsoid. <br> all defining parameters of the conformal mapping system. <br> geoidal heights required to determine $\Delta h_{i j}$. |

Table 2.9

Reduction of Observed Astronomic Azimuth to a Conformal Mapping Plane


Figure 2.3

| Status | Quantity | Remarks |
| :---: | :---: | :---: |
| 0 0 0 0 0 0 0 | $\begin{aligned} & \mathrm{d}_{\mathrm{ij}}^{\mathrm{t}}: \begin{array}{l} \text { terrain normal section } \\ \text { direction } \end{array} \\ & \mathrm{z}_{\mathrm{ij}} \quad\left(\text { or } \Delta \mathrm{H}_{i j} \text { or } \Delta \mathrm{h}_{\mathrm{ij}}\right) \end{aligned}$ | needed for reduction purposes |
| $\begin{aligned} & \text { O} \\ & \text { H } \\ & 0 \\ & 0 \\ & \text { O } \\ & 0 \end{aligned}$ | $\begin{aligned} & \left(x_{i}^{a}, y_{i}^{a}\right),\left(x_{j}^{a}, y_{j}^{a}\right) \\ & \left(\phi_{i}^{a}, \lambda_{i}^{a}\right),\left(\phi_{j}^{a}, \lambda_{j}^{a}\right) \\ & S_{i j}^{a}: \text { approximate ellipsoidal } \\ & \alpha_{i j}^{a}: \text { appronce } \\ & z_{i j}: \text { reduchath } \\ & (T-t)_{i j} \\ & \begin{array}{l} \text { rede geodetic } \\ h_{j}: \text { ellipsoidal height of } \\ \text { target } \end{array} \\ & \end{aligned}$ | see Krakiwsky et al. [1977] re coordinate transformations. <br> $\left.\begin{array}{l}\text { (3-39) or (3-62) } \\ \text { (3-37) or (3-59) }\end{array}\right\} \begin{aligned} & \text { use approximate } \\ & \text { coordinates } \\ & \text { for all computations }\end{aligned}$ (3-14) : use $\alpha_{i j}^{a}$ for this computation <br> (4-11) or (4-31) for Double Stereographic, (4-53) for $3^{\circ}$ Transverse Mercator; use approximate coordinates for all computations. |
| K K 気 | $\begin{aligned} & a, b \text { (or } a, f) \\ & \psi_{0}, \Lambda_{o}, R_{i} \\ & \phi_{o}, \lambda_{o} ; x_{o}, y_{o} ; \\ & k_{o} \\ & H_{i}\left(o r h_{i}\right): \text { orthometric (or } \\ & \text { ellipsoidal) height of } \\ & \text { instrument } \\ & N_{i}^{*}, N_{j}^{*}: \text { geoidal heights } \\ & \xi_{i}, \eta_{i}: \text { deflection of vertical } \\ & \text { components } \end{aligned}$ | parameters of reference ellipsoid <br> all defining parameters of the particular conformal mapping system. <br> required to determine $\Delta h_{i j}$ |

Table 2.10
Reduction of an Observed Direction to a Conformal Mapping Plane


Figure 2.4

| Status | Quantity | Remarks |
| :---: | :---: | :---: |
| $\begin{aligned} & \text { D} \\ & 0 \\ & 0 \\ & 0 \\ & 0 \\ & 0 \\ & 0 \end{aligned}$ | $\begin{aligned} & \beta_{j i k}: \text { angle } \\ & z_{i j}, z_{i k} \text { (or } \Delta H_{i j}, \Delta H_{i k}, \\ & \quad \text { or } \Delta h_{i j}, \Delta h_{i k} \end{aligned}$ | $\beta_{j i k}=d_{i k}^{t}-d_{i j}^{t}$ <br> needed for reduction purposes |
| $\begin{aligned} & \text { © } \\ & \text { H } \\ & \text { H } \\ & \text { Z } \\ & 0 \end{aligned}$ | $\begin{aligned} & \left(x_{i}^{a}, y_{i}^{a}\right),\left(x_{j}^{a}, y_{j}^{a}\right),\left(x_{k}^{a}, y_{k}^{a}\right) \\ & \left(\phi_{i}^{a}, \lambda_{i}^{a}\right),\left(\phi_{j}^{a}, \lambda_{j}^{a}\right),\left(\phi_{k}^{a}, \lambda_{k}^{a}\right) \\ & S_{i j}^{a}, S_{i k}^{a}: \text { approx. ellipsoidal } \\ & \alpha_{i j}^{a}, \alpha_{i k}^{a}: \begin{array}{l} \text { distances } \\ \text { azimuths } \end{array} \\ & Z_{i j}, Z_{i k}: \begin{array}{l} \text { reduced zenith } \\ \text { distances } \end{array} \\ & (T-t)_{i j}, \quad(T-t)_{i k} \\ & h_{j}, h_{k}: \text { ellipsoidal heights } \\ & \text { of targets } \end{aligned}$ | see Krakiwsky et al. [1977] re coordinate transformations. <br> (3-39) or (3-62) $\}$ use approximate coordinates (3-37) or (3-59) $\}$ for all computations (3-14): use $\alpha_{i j}^{a}, \alpha_{i k}^{a}$ for these computations (4-11) or (4-31) for Double Stereographic (4-53) for $3^{\circ}$ Transverse Mercator, use approximate coordinates for all computations |
|  | $\begin{aligned} & a, b,(o r a, f) \\ & \psi_{o}, \Lambda_{o}, R_{i} \\ & \phi_{o}, \lambda_{o} ; x_{o}, y_{o} ; k_{o} \\ & H_{i} \text { (or } h_{i} \text { ) : orthometric } \\ & \text { (or ellipsoidal) height } \\ & o f \text { instrument } \\ & N_{i}^{*}, N_{j}^{*}, N_{k}^{*}: \text { geoidal heights } \\ & \xi_{i}, \eta_{i}: \text { deflection of } \\ & \text { vertical components } \end{aligned}$ | parameter of reference ellipsoid all defining parameters of the particular conformal mapping system. $\text { required to determine } \Delta h_{i j}, \Delta h_{i k}$ |

Table 2.11
Reduction of an Observed Angle to a Conformal Mapping Plane


Figure 2.5

| Status | Quantity | Remarks |
| :---: | :---: | :---: |
| © 0 0 0 0 0 | $\begin{aligned} & r_{i j}: \text { terrain spatial distance } \\ & Z_{i j}\left(\text { or } \Delta H_{i j} \text { or } \Delta h_{i j}\right) \end{aligned}$ | Instrumental and atmospheric effects have been removed needed for reduction purposes. |
| $\begin{aligned} & \text { O} \\ & \text { H } \\ & 3 \\ & 0 \\ & 0 \\ & 0 \\ & 0 \end{aligned}$ | $\begin{aligned} & \left(x_{i}^{a}, y_{i}^{a}\right),\left(x_{j}^{a}, y_{j}^{a}\right) \\ & \left(\phi_{i}^{a}, \lambda_{i}^{a}\right),\left(\phi_{j}^{a}, \lambda_{j}^{a}\right) \\ & \alpha_{i j}^{a}, \alpha_{j i}^{a}: \text { approx. geodetic } \\ & h_{j}: \text { ellipsoidal height of } \\ & \bar{k}_{i j}: \text { target } \end{aligned}$ | see Krakiwsky et al. [1977] re coordinate transformations. <br> (3-37), (3-35) and (3-36), or (3-59), (3-60) and (3-61); use approximate coordinates for all computations. <br> (4-13) or (4-33) for Double Stereographic, (4-57) for $3^{\circ}$ Transverse Mercator; use approximate coordinates for all computations |
| K 年 K | $\begin{aligned} & a, b \text { (or } a, f) \\ & x_{o}, Y_{o} ; k_{o} \\ & H_{i} \text { (or } h_{i} \text { ) : orthometric (or } \\ & \begin{array}{l} \text { ellipsoidal) height } \\ \\ \text { of instrument } \end{array} \\ & N_{i}^{*}, N_{j}^{*}: \text { geoidal heights } \end{aligned}$ | parameters of reference ellipsoid three parameters pertaining to the particular conformal mapping system <br> required to determine $\Delta h_{i j}$ |

Table 2.12

Reduction of a Terrain Spatial Distance to a Conformal Mapping Plane
used for the numerical computations in this report, one set of approximate coordinates are used for any one project.

### 2.3 Data Screening

Prior to being used in network computations, each piece of data (azimuth, direction, angle, distance) should be tested individually to ensure that it is self-consistent. These tests are accomplished by methods of univariate analysis, which means the examination of the repeated measurement of the same observable (e.g. a distance). These repeated measurements are represented by a data series $\ell_{i}, i=1, N$ where $N$ is the sample size. The problem here is to discover which individual observations $\ell_{i}$ are statistically incompatible with the rest of the series. This subject is commonly known as the detection of outliers [e.g. Krakiwsky, 1978; Pope, 1976].

The specific test which is used to detect outliers depends on the underlying assumptions about the population mean $\mu$ and population variance $\sigma^{2}$. If $\mu$ and $\sigma^{2}$ are assumed unknown, they are estimated by the sample mean $\bar{\ell}$ and sample variance $S^{2}$. The following interpretations can be made: (a) ' $\mu$ known' corresponds to measuring a line of known length (e.g. a a calibration baseline) ; (b) ' $\sigma^{2}$ known' corresponds to measuring with an instrument of known accuracy; (c) ' $\mu$ unknown' corresponds to measuring a line of unknown length; (d) ' $\sigma$ unknown' corresponds to measuring with an instrument of unknown accuracy. The four possible combinations of the above cases are shown in Table 2.13.

The so-called null hypothesis $H_{o}$ being tested is

```
Ho: l i is a member of a sample with normal distribution.
```

| Name | Situation |  | $\mathrm{H}_{\mathrm{O}}$ (null hypothesis) | $\begin{gathered} \text { Statistic } \\ y \\ \hline \end{gathered}$ | $\begin{aligned} & \text { pdf* } \\ & \phi(y) \end{aligned}$ | 1-a** Confidence Interval for the Quantity Tested | Remarks |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | ${ }_{1}$ | ${ }^{8}$ |  |  |  |  |  |
| Normal <br> Test of <br> a Single <br> observation | $\stackrel{\mu}{\text { known }}$ | $\begin{gathered} \sigma^{2} \\ \text { known } \end{gathered}$ | $\ell_{i}$ belongs to a sample having the paf $\eta\left(l, \mu, \sigma^{2}\right)$ | $\frac{\ell-\mu}{\sigma}$ | standard <br> normal <br> $n(0,1)$ | $\mu-\sigma n_{\frac{a}{2}}<\ell_{i}<\mu+\sigma n_{1-} \frac{a}{2}$ | $\sigma$ known thus the normal distribution. |
| Student's t <br> Test of <br> a Single <br> Observation | $\stackrel{\mu}{\text { known }}$ | $\mathrm{s}^{2}$ | $\ell_{i}$ belongs to a sample having the pdf $2\left(l, \mu, s^{2}\right)$ | $\frac{\ell-\mu}{s}$ | $\begin{aligned} & \text { Student's } t \\ & t_{\mathrm{N}-1} \end{aligned}$ | $\mu-s t_{N-1, \frac{a}{2}}<\ell_{i}<\mu+s t_{N-1,1-\frac{a}{2}}$ | s is computed using $\bar{\ell}$ estimated from sample of size N thus t distribution. |
| Normal <br> Test of <br> a Single <br> Observation | $\bar{\ell}$ | $\begin{gathered} \sigma^{2} \\ \text { known } \end{gathered}$ | $\ell_{i}$ belongs to a sample having the paf $\bar{z}\left(\ell ; \bar{l}, \sigma^{2}\right)$ | $\frac{l-\bar{l}}{\left(\frac{N-1}{N}\right)^{1 / 2} \sigma}$ | standard <br> normal <br> $n(0,1)$ | $\bar{\ell}-\left(\frac{N-1}{N}\right)^{\frac{1}{2}} \sigma n_{\frac{a}{2}}<\ell_{1}<\bar{\ell}+\left(\frac{N-1}{N}\right)^{\frac{1}{2}} \sigma n_{1}-\frac{\ddot{a}}{2}$ | $\sigma$ known thus the normal distribution. |
| $\tau$ Test <br> of a <br> Single <br> Observation | $\bar{\ell}$ | $s^{2}$ | $\ell_{i}$ belongs to a sample having the paf $\eta\left(l ; \bar{l}, s^{2}\right)$ | $\frac{\ell-\bar{l}}{\left(\frac{N-1}{N}\right)^{1 / 2} s}$ | $\begin{aligned} & \text { Tau } \\ & \tau_{\mathrm{N}-1} \end{aligned}$ | $\overline{l-\left(\frac{N-1}{N}\right)^{\frac{1}{2}}{ }^{s \tau_{N-1}}, \frac{a}{2}<l_{i}<\bar{l}+\left(\frac{N-1}{N}\right)^{\frac{1}{2}}{ }_{s \tau_{N-1,1}-\frac{a}{2}} . \frac{a}{} .}$ | $\bar{l}$ and $s$ computed from the same sample thus the $\tau$ distribution. |
| ```* Pope [1976]. n(0,1) - standard normal distribution of 0 mean and variance 1. t TN - tau distribution with N-1 degrees of freedom.``` |  |  |  |  |  |  |  |

The test which is applicable most often is for $\sigma^{2}$ known and $\mu$ unknown. This corresponds to the third test in Table 2.13, i.e.

$$
\begin{equation*}
\bar{\ell}-\left(\frac{N-1}{N}\right)^{1 / 2} \sigma_{n_{1-a / 2}}<\ell_{i}<\bar{\ell}+\left(\frac{N-1}{N}\right)^{1 / 2} \sigma_{n_{1-a / 2}} \tag{2-19}
\end{equation*}
$$

where $\bar{l}=$ sample mean,
$\mathrm{N}=$ sample size,
$\sigma=$ known standard deviation,
$\mathrm{n}=$ standard normal distribution (see Table 2.14),
$a=\alpha / N$ for $\alpha=$ significance level (e.g. $\alpha=0.5$ ).

If the observation $l_{i}$ being tested does not lie within the limits given by equation (2-19), then the null hypothesis $H_{0}$ is rejected at the $(1-\alpha) \%$ confidence level.

For example, Table 2.15 lists 11 observed values of astronomic azimuth for the same line. The mean value of the series is $\bar{\ell}=85^{\circ} 36^{\prime} 18.71$, and the assumed known standard deviation $\sigma=5: 77$. For a significance level $\alpha=0.05, a / 2=\alpha / 2 N=0.05 / 22=2.273 \cdot 10^{-3}$, and $1-a / 2=0.99773$. From Table 2.14, the value for $n_{1-a / 2}$ is 2.83. Thus, the rejection limits for an individual observation $\ell_{i}$ are

$$
\begin{equation*}
85^{\circ} 36^{\prime} 03^{\prime \prime} 14<\ell_{i}<85^{\circ} 36^{\prime} 34: 28 \text {. } \tag{2-20}
\end{equation*}
$$

Performing the test for each astronomic azimuth in Table 2-15, it is seen that azimuth numbers 1, 2, 3 and 10 are rejected at the $95 \%$ confidence level, i.e. the hypothesis that they are members of a sample with normal distribution is rejected. Thus, only the seven remaining azimuths are taken as representative of the sample, and the mean value $\bar{\ell}=85^{\circ} 36^{\prime} 25^{\prime \prime} 14$ computed from these seven remaining azimuths is used for further computations.


Values of $\operatorname{Pr}$ corresponding to $c$ for the normal curve.
The value of $\operatorname{Pr}$ for ( $-c$ ) equals one minus the value of $\operatorname{Pr}$ for ( $+c$ ).

| C | . 00 | . 01 | . 02 | . 03 | . 04 | . 05 | . 06 | . 07 | . 08 | . 09 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| . 0 | . 5000 | . 5040 | . 5080 | . 5120 | . 5160 | . 5199 | . 5239 | . 5279 | . 5319 | . 5359 |
| . 1 | . 5398 | .it3s | . 5478 | . 5517 | . 5557 | . 55996 | . 5636 | . 5675 | . 5714 | . 5753 |
| . 2 | . 5793 | . 5832 | . 5871 | . 5910 | . 5948 | . 5987 | . 6026 | . 6064 | . 6103 | . 6141 |
| . 3 | . 6179 | . 6217 | . 6255 | . 6293 | . 6331 | . 6368 | . 6406 | . 6443 | . 6480 | . 6517 |
| . 4 | . 6554 | . 6591 | . 6628 | . 6664 | . 6700 | . 6736 | . 6772 | . 6808 | . 6844 | . 6879 |
| .5 | . 6915 | . 6950 | . 6985 | . 7019 | . 7054 | . 7088 | . 7123 | . 7157 | . 7190 | . 7224 |
| . 6 | . 7257 | . 7291 | . 7324 | . 7357 | . 7389 | . 7422 | . 7454 | . 7486 | . 7517 | . 7549 |
| . 7 | . 7580 | . 7611 | . 7642 | . 7673 | . 7704 | . 7734 | . 7764 | . 7794 | . 7823 | . 7852 |
| . 8 | . 7881 | . 7910 | . 7939 | . 7967 | . 7995 | . 8023 | . 8051 | . 8078 | . 8106 | . 8133 |
| . 9 | . 8159 | .81s6 | . 8212 | . 8238 | . 8264 | . 8289 | . 8315 | . 8340 | . 8365 | . 8389 |
| 1.0 | . 8413 | . 8438 | . 8461 | . 8485 | . 8508 | . 8531 | . 8554 | . 8577 | . 8599 | . 8621 |
| 1.1 | .8643 | . 8663 | . 8686 | . 8708 | . 8729 | . 8749 | . 8770 | . 8790 | . 8810 | . 8830 |
| 1.2 | . 8849 | . 8869 | . 8888 | . 8907 | . 89825 | . 8944 | . 8962 | . 8980 | . 8997 | . 9015 |
| 1.3 | . 9032 | . 9049 | . 9066 | . 9082 | . 9099 | . 9115 | . 9131 | . 9147 | . 9162 | . 9177 |
| 1.4 | . 9192 | . 9207 | . 9222 | . 9236 | . 9251 | . 9265 | . 9279 | . 9292 | . 9306 | . 9319 |
| 1.5 | . 9332 | . 9345 | . 9357 | . 9370 | . 9382 | . 9394 | . 9406 | . 9418 | . 9429 | . 9441 |
| 1.6 | . 9452 | . 9463 | . 9474 | . 9484 | . 9495 | . 9505 | . 9515 | . 9525 | . 9535 | . 9545 |
| 1.7 | . 9554 | .9564 | . 9573 | . 9582 | . 9591 | . 9599 | . 9608 | . 9616 | . 9625 | . 9633 |
| 1.8 | . 9641 | . 9649 | . 9656 | . 9664 | . 9671 | . 9678 | . 9686 | . 9693 | . 9699 | . 9706 |
| 1.9 | . 9713 | . 9719 | . 9726 | . 9732 | . 9738 | . 9744 | . 9750 | . 9756 | . 9761 | . 9767 |
| 2.0 | . 9772 | . 9738 | . 9783 | . 9788 | . 9793 | . 9798 | . 9803 | . 9808 | . 9812 | . 9817 |
| 2.1 | . 9821 | .98:6 | . 9830 | . 98834 | . 9838 | . 9842 | . 9846 | . 9850 | . 9885 | . 9857 |
| 2.2 | . 9861 | . 0864 | . 9868 | . 8871 | . 9875 | . 9878 | . 9881 | . 9884 | . 9887 | . 9890 |
| 2.3 | . 0893 | .asal | . 9898 | . 9901 | . 99004 | . 9906 | . 9909 | . 99911 | . 9931 | . 9916 |
| 2.4 | . 9918 | .99\% | . 9322 | . 9325 | . 9927 | . 9922 | . 9931 | . 3932 | . 9934 | . 9936 |
| 2.5 | . 9938 | . 98.40 | . 9941 | . 9943 | . 9945 | . 9946 | . 9948 | . 9949 | . 9951 | . 9952 |
| 2.6 | . 9953 | . 9953 | . 9956 | . 9955 | . 9959 | . 9960 | . 99961 | . 9962 | . 9963 | . 9964 |
| 2.7 | . 9965 | . 9366 | . 9967 | . 9968 | . 9969 | . 9970 | . 9971 | . 9972 | . 9973 | . 9974 |
| 2.8 | . 9874 | . 9975 | . 9976 | . 9977 | . 9977 | . 9978 | . 9979 | . 9979 | . 9980 | . 9981 |
| 2.9 | . 9981 | . 0982 | . 9982 | . 9983 | . 9984 | . 9984 | . 9985 | . 9985 | . 9986 | . 9986 |
| 3.0 | . 9987 | . 9097 | . 9987 | . 9388 | . 3988 | . 9389 | . 9989 | . 9989 | . 9990 | . 9990 |
| 3.1 | . 9990 | . 0991 | . 9991 | . 9991 | . 9392 | . 9992 | . 99392 | . 9932 | . 9993 | . 3993 |
| 3.2 | . 9993 | . 01913 | . 9894 | . 9993 | . 9331 | . 9993 | . 9994 | . 9995 | . 9795 | . 9995 |
| 3.3 | . 9995 | . $939 \%$ | . 9995 | . 99996 | . 9999 | . 99996 | . 9998 | . 9996 | . 9396 | . 9997 |
| 3.4 | . 0997 | . | . 9997 | . 9997 | . 9997 | . 9997 | . 9997 | . 9397 | . 9997 | . 9998 |

Table 2.14

Cumulative Normal Distribution - Values of $\operatorname{Pr}$

| SET NO. | OBS ' D | AZIMUTH |  |
| :---: | :---: | :---: | :---: |
| 1 | 85 | 35 | 27.84 |
| 2 | 85 | 35 | 34.36 |
| 3 | 85 | 36 | 51.15 |
| 4 | 85 | 36 | 27.95 |
| 5 | 85 | 36 | 28.89 |
| 6 | 85 | 36 | 28.43 |
| 7 | 85 | 36 | 20.46 |
| 9 | 85 | 36 | 22.29 |
| 10 | 85 | 36 | 24.02 |
| 11 | 85 | 36 | 36.46 |

Table 2.15 Astronomic Azimuth Data Series

If the underlying assumptions for the univariate test are different (e.g. $\sigma$ unknown, $\mu$ known), then one of the three other tests of Table 2.13 should be used. Essentially, this changes only the rejection limits (cf. eq. (2-20)). The necessary tables (i.e. Student's $t$ and tau distributions) can be found in e.g. Rainsford [1957] and Pope [1976].

## 3. MATHEMATICAL MODELS FOR AZIMUTH, DIRECTION AND ANGLE OBSERVATIONS

In this chapter, the mathematical models relating angular observations and coordinates are given. Coordinates are the $x(e a s t i n g)$ and $y$ (northing) coordinates referred to a conformal mapping plane (e.g. Krakiwsky et al., 1977; Thomson et al., 1978]. Both the linear and nonlinear forms of the mathematical model for azimuths, directions, and angles are given.

### 3.1 Azimuth Mathematical Model

The nonlinear form of the azimuth mathematical model is

$$
\begin{equation*}
F_{i j}=\arctan \left(\frac{x_{j}-x_{i}}{y_{j}-y_{i}}\right)-t_{i j}=0 \tag{3-1}
\end{equation*}
$$

where the first term is a nonlinear function of the coordinates of two points $i$ and $j$ (see Figure 3.1), and $t_{i j}$ is the observed azimuth from point $i$ to point $j$ reduced to the mapping plane [e.g. Thomson et al., 1978, section 4.2.3]. A linear Taylor series (see Appendix I) is used to approximate this nonlinear model. The resulting equation is

$$
F_{i j}=F_{i j}^{0}+d F_{i j}=\arctan \left(\frac{x_{j}^{0}-x_{i}^{o}}{y_{j}^{0}-y_{i}^{0}}\right)-t_{i j}+d t_{i j}-v_{t_{i j}}+\ldots=0
$$

where $\arctan \left(\frac{x_{j}^{0}-x_{i}^{0}}{y_{j}^{0}-y_{i}^{0}}\right)=$ computed value of the azimuth based on approximate values of the coordinates ( $\mathrm{x}^{\mathrm{O}}, \mathrm{y}^{\mathrm{o}}$ ),
$d t_{i j}=$ differential change in the computed azimuth resulting from differential changes in the approximate coordinates (see eq. (3-3)),


Figure 3.1 Mapping Plane Azimuth

$$
\mathrm{v}_{\mathrm{t}_{\mathrm{ij}}}=\text { correction to the observed mapping plane (grid) azimuth. }
$$

The differential change in azimuth $d t_{i j}$ is given as

$$
\begin{equation*}
d t_{i j}=\frac{\partial t_{i j}}{\partial x_{i}} d x_{i}+\frac{\partial t_{i j}}{\partial y_{i}} d y_{i}+\frac{\partial t_{i j}}{\partial x_{j}} d x_{j}+\frac{\partial t_{i j}}{\partial y_{j}} d y_{j} \tag{3-3}
\end{equation*}
$$

Evaluation of the partial derivatives in equation (3-3) yields

$$
\begin{align*}
& \frac{\partial t_{i j}}{\partial x_{i}}=\frac{-\left(y_{j}^{0}-y_{i}^{0}\right)}{\left(l_{i j}^{0}\right)^{2}}=a_{i j},  \tag{3-4}\\
& \frac{\partial t_{i j}}{\partial y_{i}}=\frac{\left(x_{j}^{0}-x_{i}^{O}\right)}{\left(l_{i j}^{0}\right)^{2}}=b_{i j},  \tag{3-5}\\
& \frac{\partial t_{i j}}{\partial x_{j}}=\frac{\left(y_{j}^{0}-y_{i}^{0}\right)}{\left(l_{i j}\right)^{2}}=-a_{i j},  \tag{3-6}\\
& \frac{\partial t_{i j}}{\partial y_{j}}=\frac{-\left(x_{j}^{0}-x_{i}^{0}\right)}{\left(l_{i j}^{0}\right)^{2}}=-b_{i j}, \tag{3-7}
\end{align*}
$$

where $\ell_{i j}^{0}$ is the mapping plane distance between points $i$ and $j$ computed using the approximate coordinates. Substituting equations (3-3) to (3-7) into equation (3-2) yields the so-called observation equation as

$$
v_{t_{i j}}^{\prime \prime}=\left[\arctan \left(\frac{x_{j}^{o}-x_{i}^{o}}{y_{j}^{\circ}-y_{i}^{o}}\right)-t_{i j}\right]^{\prime \prime}+\rho " a_{i j} \delta x_{i}+\rho " b_{i j} \delta y_{i}-\rho " a_{i j} \delta x_{j}-\rho " b_{i j} \delta y_{j},
$$

where $\rho "=206264$ " 8062 is used to proportion the elements, and
$\delta x_{i} \rightarrow \delta y_{j}=$ differential changes in the coordinates.
Converting equation (3-8) to matrix notation yields the matrix form of the observation equation as

$$
v_{t_{i j}}^{\prime \prime}=\left[\arctan \left(\frac{x_{j}^{o}-x_{i}^{o}}{y_{j}^{0}-y_{i}^{o}}\right)-t_{i j}\right] "+\rho "\left[a_{i j} b_{i j}-a_{i j}-b_{i j}\right]\left[\begin{array}{l}
\delta x_{i}  \tag{3-9}\\
\delta y_{i} \\
\delta x_{j} \\
\delta y_{j}
\end{array}\right] .
$$

The matrix form of the evaluated partial derivatives (i.e. $\left.a_{i j}, b_{i j}, e t c.\right)$ is called the design matrix $A$, the difference between computed and observed azimuths is called the misclosure vector $W$, the vector of differential changes in coordinates $(\delta x, \delta y)$ is called the solution vector $\hat{X}$, and the correction to the observed azimuth is called the residual vector $V$. Rewriting equation (3-9) in this notation yields

$$
\begin{equation*}
\underset{(1,1)}{v_{t_{i j}}}=\mathrm{w}_{t_{i j}}+{ }_{(1,1)}^{A_{t_{i j}}} \quad \hat{(1,4)} \quad \hat{x} . \tag{3-10}
\end{equation*}
$$

### 3.2 Direction Mathematical Model

Direction observations are relative to the 'zero' direction of the horizontal circle of a theodolite. The azimuth of this zero direction is called the orientation unknown $Z$ (see Figure 3.2) and it must be solved for along with the unknown coordinates. Orientation unknowns are not desired quantities and thus are called nuisance parameters. The nonlinear form of the direction mathematical model is

$$
\begin{equation*}
F_{i j}=\arctan \left(\frac{x_{j}^{-x_{i}}}{y_{j}-y_{i}}\right)-\left(d_{i j}+z_{i}\right)=0 \tag{3-11}
\end{equation*}
$$

where $d_{i j}=$ observed direction from point $i$ to point $j$ reduced to the mapping plane [e.g. Thomson et al., 1978, section 4.2.1], $z_{i}=$ orientation unknown at point i.


Figure 3.2 Direction on the Mapping Plane

The linear Taylor series expansion of equation (3-1l) gives the linearized direction mathematical model as

$$
\begin{equation*}
F_{i j}=\arctan \left(\frac{x_{j}^{0}-x_{i}^{0}}{y_{j}^{0}-y_{i}^{0}}\right)-z_{i}^{0}-d_{i j}+d t_{i j}-d z_{i}-v_{d_{i j}}+\ldots=0 \tag{3-12}
\end{equation*}
$$

where $z_{i}^{0}$ = approximate value of the orientation unknown,

$$
\begin{aligned}
\mathrm{dz}_{\mathrm{i}}= & \text { differential change in the orientation unknown } \mathrm{z}_{\mathrm{i}} \text { caused by an } \\
& \text { inaccurate approximate value } \mathrm{z}_{\mathrm{i}}^{0} \\
\mathrm{~V}_{\mathrm{d}_{\mathrm{ij}}}= & \text { correction to the observed direction. }
\end{aligned}
$$

The approximate value of the orientation unknown $z_{i}^{0}$ is obtained by subtracting the observed direction to a station $j$ from the azimuth to the same station computed from the approximate coordinates, i.e.

$$
\begin{equation*}
z_{i}^{o}=\arctan \left(\frac{x_{j}^{0}-x_{i}^{0}}{y_{j}^{0}-y_{i}^{o}}\right)-d_{i j} \tag{3-13}
\end{equation*}
$$

Realizing that $d t_{i j}$ has already been evaluated for the linearized azimuth mathematical model (see eqs. (3-3) to (3-7)), the observation equation for a direction is

$$
\begin{align*}
v_{d_{i j}}^{\prime \prime}=\left[\arctan \left(\frac{x_{j}^{o}-x_{i}^{o}}{y_{j}^{o}-y_{i}^{o}}\right)-z_{i}^{o}-d_{i j}\right] " & +\rho " a_{i j} \delta x_{i}+\rho " b_{i j} \delta y_{i}-\rho " a_{i j} \delta x_{j} \\
& -\rho " b_{i j} \delta y_{j}-\delta z_{i}, \tag{3-14}
\end{align*}
$$

or, in matrix form
$v_{d_{i j}}^{\prime \prime}=\left[\arctan \left(\frac{x_{j}^{o-x_{i}^{o}}}{y_{j}^{o}-y_{i}^{o}}\right)-z_{i}^{o}-d_{i j}\right]^{\prime \prime}+\rho "\left[a_{i j} b_{i j}-a_{i j}-b_{i j}-1\right]\left[\begin{array}{c}\delta x_{i} \\ \delta y_{i} \\ \delta x_{j} \\ \delta y_{j} \\ \delta z_{i}\end{array}\right]$.

Using the symbolic matrix notation of section 3.1 , equation (3-15) becomes

$$
\begin{aligned}
& \mathrm{v}_{\mathrm{d}_{i j}}=\mathrm{w}_{\mathrm{d}_{i j}}+\mathrm{A}_{\mathrm{d}_{i j}} \hat{\mathrm{X}} . \\
& (1,1) \quad(1,1) \quad(1,5)(5,1)
\end{aligned}
$$

### 3.3 Angle Mathematical Model

The nonlinear mathematical model for an angle is (see Figure 3.3)

$$
\begin{equation*}
F_{i j k}=\arctan \left(\frac{x_{k}^{-x_{i}}}{y_{k}^{-y_{i}}}\right)-\arctan \left(\frac{x_{j}^{-x_{i}}}{y_{j}^{-y_{i}}}\right)-B_{i j k}=0 \tag{3-17}
\end{equation*}
$$

where $B_{i j k}=$ angle observed at point $i$ from point $j$ to point $k$ reduced to the mapping plane [e.g. Thomson et al., 1978, section 4.2.2]. The difference between two direction mathematical models $F_{i k}$ and $F_{i j}$ has the same form as the angle mathematical model $F_{i j k}$. The linearized form of the angle mathematical model is

$$
F_{i j k}=\arctan \left(\frac{x_{k}^{0}-x_{i}^{o}}{y_{k}^{0}-y_{i}^{o}}\right)-\arctan \left(\frac{x_{j}^{0}-x_{i}^{o}}{y_{j}^{0}-y_{i}^{o}}\right)-B_{i j k}+d B_{i j k}-v_{B_{i j k}}+\ldots=0,(3-18)
$$

where $d B_{i j k}=$ differential change in the computed angle resulting from differential changes in the approximate coordinates,


The differential change in the angle $d B_{i j k}$ is given as

$$
d B_{i j k}=\frac{\partial B_{i j k}}{\partial x_{i}} d x_{i}+\frac{\partial B_{i j k}}{\partial y_{i}} d y_{i}+\frac{\partial B_{i j k}}{\partial x_{j}} d x_{j}+\frac{\partial B_{i j k}}{\partial y_{j}} d y_{j}+\frac{\partial B_{i j k}}{\partial x_{k}} d x_{k}+\frac{\partial B_{i j k}}{\partial y_{k}} d y_{k}
$$

Evaluation of the partial derivatives in equation (3-19) yields

$$
\begin{equation*}
\frac{\partial B_{i j k}}{\partial x_{i}}=\frac{-\left(y_{k}^{0}-y_{i}^{\circ}\right)}{\left(\ell_{i k}^{0}\right)^{2}}+\frac{\left(y_{j}^{0}-y_{i}^{\circ}\right)}{\left(\ell_{i j}^{0}\right)^{2}}=c_{i j k} \tag{3-20}
\end{equation*}
$$



Figure 3.3 Angle on the Mapping Plane

$$
\begin{align*}
& \frac{\partial B_{i j k}}{\partial y_{i}}=\frac{\left(x_{k}^{0}-x_{i}^{0}\right)}{\left(l_{i k}^{0}\right)^{2}}-\frac{\left(x_{j}^{0}-x_{i}^{0}\right)}{\left(l_{i j}^{0}\right)^{2}}=d_{i j k},  \tag{3-21}\\
& \frac{\partial B_{i j k}}{\partial x_{j}}=\frac{-\left(y_{j}^{\circ}-y_{i}^{\circ}\right)}{\left(l_{i j}^{O}\right)^{2}}=a_{i j},  \tag{3-22}\\
& \frac{\partial B_{i j k}}{\partial y_{j}}=\frac{\left(x_{j}^{O}-x_{i}^{O}\right)}{\left(l_{i j}^{O}\right)^{2}}=b_{i j},  \tag{3-23}\\
& \frac{\partial B_{i j k}}{\partial x_{k}}=\frac{\left(y_{k}^{0}-y_{i}^{O}\right)}{\left(l_{i k}^{O}\right)}=-a_{i k},  \tag{3-24}\\
& \frac{\partial B_{i j k}}{\partial y_{k}}=\frac{-\left(x_{k}^{0}-x_{i}^{0}\right)}{\left(l_{i k}^{O}\right)^{2}}=-b_{i k}, \tag{3-25}
\end{align*}
$$

where $l_{i k}^{0}$ is the mapping plane distance between points $i$ and $k$ computed using the approximate coordinates. Substituting the above values into equations (3-19) and subsequently evaluating the linearized mathematical model (eq. (3-18)) gives the observation equation as

$$
\begin{align*}
v_{B_{i j k}}^{\prime \prime}= & {\left[\arctan \left(\frac{x_{k}^{\circ}-x_{i}^{\circ}}{y_{k}^{\circ}-y_{i}^{\circ}}\right)-\arctan \left(\frac{x_{j}^{0}-x_{i}^{o}}{y_{j}^{\circ}-y_{i}^{\circ}}\right)-B_{i j k}\right] "+\rho " c_{i j k} \delta x_{i}+\rho " d_{i j k} \delta y_{i} } \\
& +\rho " a_{i j} \delta x_{j}+\rho " b_{i j} \delta y_{j}-\rho " a_{i k} \delta x_{k}-\rho " b_{i k} \delta \dot{y}_{k} \tag{3-26}
\end{align*}
$$

or, in matrix form
$Y_{B_{i j k}}^{\prime \prime}=\left[\arctan \left(\frac{x_{k}^{0}-x_{i}^{o}}{y_{k}^{0}-y_{i}^{0}}\right)-\arctan \left(\frac{x_{j}^{-}-x_{i}}{y_{j}^{0}-y_{i}^{0}}\right)-B_{i j k}\right]+\rho^{\prime \prime}\left[c_{i j k}{ }^{d_{i j k}}{ }^{a}{ }_{i j j} b_{i j}-a_{i k}-b_{i k}\right]\left[\begin{array}{l}\delta x_{i} \\ \delta y_{i} \\ \delta x_{j} \\ \delta y_{j} \\ \delta x_{k} \\ \delta y_{k}\end{array}\right] \cdot$
Again, using the symbolic matrix notation, equation (3-27) becomes

$$
\begin{aligned}
& \mathrm{V}_{B_{i j k}}=W_{B_{i j k}}+A_{B_{i j k}} \quad \hat{X} . \\
& (1,1) \quad(1,1) \quad(1,6) \quad(6,1)
\end{aligned}
$$

## 4. MATHEMATICAL MODELS FOR DISTANCE OBSERVATIONS

This chapter describes the nonlinear and linearized mathematical models relating plane coordinates $(x, y)$ to observed distances $\ell_{i j}$ from point $i$ to point $j$ reduced to a conformal mapping plane [e.g. Thomson et al., 1978, section 4.2.4]. The nonlinear form of the distance mathematical model is (see Figure 4.1)

$$
\begin{equation*}
F_{i j}=\left(\left(x_{j}-x_{i}\right)^{2}+\left(y_{j}-y_{i}\right)^{2}\right)^{1 / 2}-\ell_{i j}=0 \tag{4-1}
\end{equation*}
$$

Linearization of equation (4-1) by a linear Taylor series
(see Appendix I) expansion gives the linearized form of the distance mathematical model as

$$
F_{i j}=F_{i j}^{o}+d F_{i j}=\left(\left(x_{j}^{0}-x_{i}^{o}\right)^{2}+\left(y_{j}-y_{i}\right)^{2}\right)^{1 / 2}-\ell_{i j}+d \ell_{i j}-v_{\ell}+\ldots=0
$$

where $\left(\left(x_{j}^{0}-x_{i}^{0}\right)^{2}+\left(y_{j}^{0}-y_{i}^{o}\right)^{2}\right)^{l / 2}=$ computed value of the distance based on approximate values of the coordinates $\left(x^{\circ}, y^{\circ}\right)$,
$d \ell_{i j}=$ differential change in the computed distance resulting from differential changes in the approximate coordinates (see eq. (4-3)),
$\dot{\mathrm{V}}_{\ell}{ }_{i j}=$ correction to the observed mapping plane distance.

The differential change in distance $\mathrm{d}_{\mathrm{ij}}$ is given as

$$
\begin{equation*}
d \ell_{i j}=\frac{\partial \ell_{i j}}{\partial x_{i}} d x_{i}+\frac{\partial \ell_{i j}}{\partial y_{i}} d y_{i}+\frac{\partial \ell_{i j}}{\partial x_{j}} d x_{j}+\frac{\partial \ell_{i j}}{\partial y_{j}} d y_{j} \tag{4-3}
\end{equation*}
$$

where the partial derivatives are


Figure 4.1 Distance on the Mapping Plane

$$
\begin{align*}
& \frac{\partial \ell_{i j}}{\partial x_{i}}=\frac{-\left(x_{j}^{0}-x_{i}^{0}\right)}{\ell_{i j}^{0}}=e_{i j},  \tag{4-4}\\
& \frac{\partial \ell_{i j}}{\partial y_{i}}=\frac{-\left(y_{i}^{\left.-y_{i}\right)}\right.}{\ell_{i j}^{0}}=f_{i j},  \tag{4-5}\\
& \frac{\partial \ell_{i j}}{\partial x_{j}}=\frac{\left(x_{j}^{0}-x_{i}^{o}\right)}{\ell_{i j}^{0}}=-e_{i j},  \tag{4-6}\\
& \frac{\partial \ell_{i j}}{\partial y_{j}}=\frac{\left(y_{j}^{0}-y_{i}^{o}\right)}{\ell_{i j}^{0}}=-f_{i j}, \tag{4-7}
\end{align*}
$$

where $\ell_{i j}^{0}$ is the mapping plane distance between points $i$ and $j$ computed using the approximate coordinates. Substituting these partial derivatives into equation (4-3) to obtain $d \ell_{i j}$, and, in turn, substituting $d \ell_{i j}$ back into equation (4-2) yields the distance observation equation as

$$
\begin{equation*}
v_{\ell}=\left[\left(\left(x_{j}^{o}-x_{i}^{o}\right)^{2}+\left(y_{j}^{o}-y_{i}^{o}\right)^{2}\right)^{1 / 2}-\ell_{i j}\right]+e_{i j} \delta x_{i}+f_{i j} \delta y_{i}-e_{i j} \delta x_{j}-f_{i j} \delta y_{j} \tag{4-8}
\end{equation*}
$$

or in matrix form

$$
v_{\ell j}=\left[\left(\left(x_{j}^{\circ}-x_{i}^{\circ}\right)^{2}+\left(y_{j}^{0}-y_{i}^{\circ}\right)^{2}\right)^{1 / 2}-\ell_{i j}\right]+\left[e_{i j} f_{i j}-e_{i j}-f_{i j}\right]\left[\begin{array}{l}
\delta x_{i}  \tag{4-9}\\
\delta y_{i} \\
\delta x_{j} \\
\delta y_{j}
\end{array}\right]
$$

The units of both equations (4-8) and (4-9) are metres. Converting the observation equation into symbolic matrix notation as in section (3-1) (cf. eq. (3-10)) yields

$$
\begin{align*}
& v_{\ell j}=W_{\ell i j}+A_{\ell j} \quad \hat{X} .  \tag{4-10}\\
& (1,1) \quad(1,1) \quad(1,4) \quad(4,1)
\end{align*}
$$

## 5. SOLUTION OF UNIQUE CASES

This chapter covers the unique cases (i.e number of observations $n=$ number of parameters $u$ ) of coordinate determination encountered in practice (e.g. direct problem, intersection, resection, traverse). The inverse (explicit) mathematical models developed in chapters 3 and 4 are combined using the method of least squares (see Appendix II) to solve these unique cases. This leads to a unified approach when the overdetermined ( $n>u$ ) cases are considered in chapter 6. For treatment of these unique cases by the direct method (cf. eq. (1-1)), the reader is referred to e.g. Faig [1972], Thomson et al. [1978], Richardus [1974].

For all of the examples considered in this chapter, it is assumed that the observations have been reduced to a conformal mapping plane as explained in section 2.2. Thus, although the examples are not explicitly spelled out for each of the three existing Maritime conformal mapping planes [Krakiwsky et al., 1977], the methods used are equally applicable to all three provinces. The only differences are in the actual values of the initial approximate coordinates and the final adjusted coordinates.

All of the examples in this and following chapters have been performed by program GEOPAN [Steeves, 1978].

### 5.1 Direct Problem

The direct problem considered here (see Figure 5.1) is essentially the same as that used in section 4.8.1 of Thomson et al. [1978]. The only difference is that here point 1 is considered fixed (i.e. its covariance matrix is zero) whereas in Thomson et al. [1978], point 1 had a covariance matrix associated with it.


Figure 5.1 Direct Problem
The approximate coordinates of point 2, the fixed coordinates of point 1 , the reduced observations and their standard deviations are given in Table 5.1. The standard deviations are derived through the formulae developed in section 2.1. For instance, the standard


Table 5.1. Initial Data for Direct Problem
deviation of $5: .0$ for the observed azimuth could result from the azimuth observed three times by the hour angle method with a $l^{\prime \prime}$ theodolite and $\sigma_{t}=1.0 s(c f$. Table 2.1). Similarly, the distance standard deviation of 0.03 m could result (assuming $\sigma_{r}=\sigma_{\ell}$ ) from the distance observed by a CA 1000 with $\sigma_{p}=5$ mbar and $\sigma_{T}=\sigma_{\Delta T}=1^{\circ} \mathrm{C}$ (cf. Table 2.8). The approximate coordinates are determined graphically or analytically as suggested in section 2.2.

The mathematical model used here is a combination of equations (3-10) and (4-10). The residual vector $V$ is defined as

$$
\begin{gathered}
v \\
v \\
(2,1)
\end{gathered} \underset{(2,1)}{W}+\underset{(2,2)}{ } \quad \hat{X},
$$

or explicitly

$$
V(2,1)=\left[\begin{array}{c}
\arctan \left(\frac{x_{2}^{\circ}-x_{1}}{y_{2}^{\circ}-y_{1}}\right)-t_{1,2}  \tag{5-2}\\
\left(\left(x_{2}^{\circ}-x_{1}\right)^{2}+\left(y_{2}^{\circ}-y_{1}\right)^{2}\right)^{1 / 2}-\ell_{1,2}
\end{array}\right]+\left[\begin{array}{cc}
\frac{\left(y_{2}^{\circ}-y_{1}\right)}{\left(\ell_{1,2}^{0}\right)^{2}} & 0 " \frac{-\left(x_{2}^{\circ}-x_{1}\right)}{\left(\ell_{1,2}^{0}\right)^{2}} \\
\frac{\left(x_{2}^{\circ}-x_{1}\right)}{\ell_{1,2}^{\circ}} & \frac{\left(y_{2}^{\circ}-y_{1}\right)}{\ell_{1,2}^{\circ}}
\end{array}\right]\left[\begin{array}{l}
\delta x_{2} \\
\delta y_{2}
\end{array}\right]
$$

where the units are

$$
\underset{(2,1)}{V}=\left[\begin{array}{l}
"  \tag{5-3}\\
m
\end{array}\right]+\left[\begin{array}{cc}
" m^{-1} & " m^{-1} \\
- & -
\end{array}\right] \quad\left[\begin{array}{l}
m \\
m
\end{array}\right]
$$

The solution $\hat{X}$ is given from Appendix II, equation (AIII-11)
as

$$
\begin{equation*}
\hat{X}=-\left[A^{T} P A\right]^{-1} \quad A^{T} P W \tag{5-4}
\end{equation*}
$$

Thus, using the above coordinates and the observations with their standard deviations, the $A, P$ and $W$ matrices are numerically evaluated yielding

$$
A=\left[\begin{array}{cc}
59.17941 & -57.65335 \\
0.6978111 & 0.7162819
\end{array}\right], W=\left[\begin{array}{c}
-23.0324 \\
0.11655
\end{array}\right],
$$

and, assuming $\sigma_{0}^{2}=1$ (cf. eq. (AII-3)),

$$
P=\left[\begin{array}{cc}
0.04 & 0 \\
0 & 1111.11
\end{array}\right] \text { in units of }\left[\begin{array}{ll}
(n)^{-2} & \\
& m^{-2}
\end{array}\right]
$$

Using these matrices, evaluation of equation (5-4) yields

$$
\hat{x}=\left[\begin{array}{c}
0.11835 \mathrm{~m} \\
-0.27802 \mathrm{~m}
\end{array}\right]
$$

for the first iteration, and the corresponding values of the parameters are (eq. (AII-12))

$$
x=x^{0}+\hat{x}=\left[\begin{array}{c}
378907.0  \tag{5-5}\\
864184.0
\end{array}\right]+\left[\begin{array}{c}
0.11835 \\
-0.27802
\end{array}\right]=\left[\begin{array}{c}
378907.118 \mathrm{~m} \\
864183.722 \mathrm{~m}
\end{array}\right]
$$

These values for the parameters are now taken as new approximate coordinates, and the $A$ and $W$ matrices are reevaluated. They are

$$
A=\left[\begin{array}{cc}
59.17573 & -57.66265 \\
0.6978911 & 0.7162039
\end{array}\right], W=\left[\begin{array}{c}
0.001075 \\
0.000016
\end{array}\right]
$$

Evaluating equation (5-4) for the second time (iteration) yields

$$
\hat{x}=\left[\begin{array}{c}
-0.00002 \\
0.00000
\end{array}\right]
$$

which is insignificant (i.e. less than 0.001 m ), and thus the solution has converged. The final least squares estimate of the coordinates of point 2 are

$$
x=\left[\begin{array}{l}
x_{2} \\
y_{2}
\end{array}\right]=\left[\begin{array}{cc}
378907.118 & m \\
864183.722 & m
\end{array}\right]
$$

These are identical to the results obtained in Thomson et al. [1978], section 4.8.1.

The variance covariance matrix $C_{x}$ of the parameters is given by equation (AII-16) as

$$
\begin{equation*}
C_{X}=\left[A^{T} C_{L}^{-1} A\right]^{-1} \tag{5-6}
\end{equation*}
$$

If $\hat{X}$ has been computed according to formula (5-4), then this is a byproduct of computing the solution vector $\hat{X}$. In this case it is

$$
C_{x}=\left[\begin{array}{cc}
0.2305 \cdot 10^{-2} & -0.13925 \cdot 10^{-2} \\
-0.13925 \cdot 10^{-2} & 0.2233 \cdot 10^{-2}
\end{array}\right]
$$

in units of $\mathrm{m}^{2}$. Computing the standard error ellipse according to formulae (AIII-8), (AIII-5), (AIII-6) and (AIII-14) gives

$$
\begin{aligned}
& \mathrm{a}_{\mathrm{s}}=0.061 \mathrm{~m} \\
& \mathrm{~b}_{\mathbf{s}}=0.030 \mathrm{~m} \\
& \theta=-45^{\circ} 44^{\prime} 32^{\prime \prime} .
\end{aligned}
$$

Assuming the a priori variance factor known, the $c$ factor to increase the confidence level to $95 \%$ is (see Table AIII.l)

$$
c=\left(x_{2,0.95}^{2}\right)^{1 / 2}=(5.99)^{1 / 2}=2.45 .
$$

Thus, the $95 \%$ confidence ellipse has a semi-major axis

$$
\mathrm{a}=0.149 \mathrm{~m}
$$

and a semi-minor axis

$$
\mathrm{b}=0.074 \mathrm{~m}
$$

The orientation remains the same as the standard ellipse. The 95\% confidence ellipse is depicted in Figure 5.2.

In this as in all of the unique cases, the residual vector V computed (eq. (AII-17)) after the final iteration is equal to zero. The observations can give only one value for the parameters, and thus there are no residual corrections for the observations. The a posteriori variance factor $\hat{\sigma}_{0}^{2}$ is zero in this case as well.


Figure 5.2 Confidence Ellipse for Direct Problem
5.2 Azimuth Intersection

Figure 5.3 shows the azimuth intersection example considered in this section. The mathematical model used.here is that of

## O $=$ Unknown Point

 $\Delta=$ Fixed Point

Figure 5.3 Azimuth Intersection on the Plane
section 3.1 , specifically equations $(3-8),(3-9)$ and (3-10). In this case, there are two azimuth observations (namely $t_{3,1003}$ and $t_{4,1003}$ ) and two unknowns (coordinates of station 1003). Thus, the matrix form of the observation equation is

$$
\begin{equation*}
\underset{(2,1)}{v}=\underset{(2,1)}{W}+\underset{(2,2)}{A} \underset{(2,1)}{\hat{x}} \tag{5-8}
\end{equation*}
$$

or, explicitly

The approximate coordinates of points 1003 and the known coordinates of points 3 and 4 as well as the values and standard deviations of the observed azimuths reduced to the mapping plane are listed in Table 5.2. The standard deviation of 4.0 could result from 4 determinations using the hour angle method with a $l^{\prime \prime}$ theodolite and $\sigma_{t}=1.0$ s (see Table 2.1).

## Coordinates of Points

| Station | $\underline{X(m)}$ | $\underline{Y(m)}$ | From | To | Value | $\underline{\sigma}$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1003 | 3265.0 | 645.0 | 3 | 1003 | $272^{\circ} 10^{\prime} 29.71$ | $4: 0$ |  |
| 3 | 3660.0 | 630.0 | 4 | 1003 | $308^{\circ} 15^{\prime}$ | 15.94 | 4.0 |
| 4 | 3635.0 | 355.0 |  |  |  |  |  |

Table 5.2 Initial Data for Azimuth Intersection

The solution vector $\hat{\mathrm{x}}$ is given as (cf. eq. (AII-1l))

$$
\hat{\dot{x}}=-\left[A^{T} P A\right]^{-1} A^{T} P W
$$

Thus, using the above approximate coordinates for point 1003 and the observed azimuths and their standard deviations, the $A, P$ and $W$ matrices and vector are evaluated. They are

$$
A=\left[\begin{array}{cc}
19.80142 & 521.4374 \\
270.6642 & 345.3302
\end{array}\right], \quad W=\left[\begin{array}{c}
-0.63035 \\
0.36394
\end{array}\right]
$$

and, assuming $\sigma_{0}^{2}=1$,

$$
P=\left[\begin{array}{cc}
0.0625 & 0 \\
0 & 0.0625
\end{array}\right]
$$

The units for $P$ are (") ${ }^{-2}$ and for $A$ are (" /m).. Evaluation of the solution vector $\hat{\mathrm{X}}$ gives

$$
\hat{x}=\left[\begin{array}{cc}
-0.01575 & \mathrm{~m} \\
0.00181 & \mathrm{~m}
\end{array}\right]
$$

Thus, the least squares estimate of the parameters after this first iteration are

$$
\begin{gather*}
x=x^{0}+\hat{x} \\
x=\left[\begin{array}{c}
3265.0 \\
645.0
\end{array}\right]+\left[\begin{array}{c}
-0.01575 \\
0.00181
\end{array}\right]=\left[\begin{array}{c}
3264.984 \mathrm{~m} \\
645.002 \mathrm{~m}
\end{array}\right] \tag{5-10}
\end{gather*}
$$

These values for the parameters are now taken as new approximate values $x^{\circ}$, and the $A$ and $W$ matrices are reevaluated. They are

$$
A=\left[\begin{array}{ll}
19.8022 & 521.4165 \\
270.6504 & 345.3251
\end{array}\right], \quad W=\left[\begin{array}{c}
0.0001046 \\
-0.00002522
\end{array}\right]
$$

This results in a solution vector $\hat{X}$ of

$$
\hat{x}=\left[\begin{array}{l}
0.00000 \\
0.00000
\end{array}\right]
$$

Thus, the parameters $X$ are unchanged by the results of this second solution or iteration, and it has converged. The variance covariance matrix $C_{x}$ of the parameters is

$$
c_{x}=\left[\begin{array}{cc}
0.347049 .10^{-3} & -0.920918 .10^{-4} \\
-0.920918 .10^{-4} & 0.6534486 .10^{-4}
\end{array}\right]
$$

The standard ellipse computed according to Appendix III is

$$
\begin{aligned}
& \mathrm{a}_{\mathrm{S}}=0.019 \mathrm{~m} \\
& \mathrm{~b}_{\mathrm{S}}=0.006 \mathrm{~m}, \\
& \theta=-72^{\circ} 24^{\prime} 41^{\prime \prime} .
\end{aligned}
$$

Assuming $\sigma_{0}^{2}$ known, the $95 \%$ confidence ellipse is
$a=0.047 \mathrm{~m}$,
$\mathrm{b}=0.015 \mathrm{~m}$.
Figure 5.4 shows the $95 \%$ confidence ellipse.


Figure 5.4 95\% Confidence Ellipse for Azimuth Intersection

### 5.3 Distance Intersection

Figure 5.5 shows the distance intersection example used
for


Figure 5.5 Distance Intersection
this section. Points 3 and 4 are fixed, and point 1003 is unknown. Thus, there are two unknown parameters and two observations giving the unique case again.

The approximate coordinates for point 1003, the fixed coordinates of

| Coordinates of Points |  |  | Observations on the Mapping Plane |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Station | $\underline{X}(\mathrm{~m})$ | $\mathrm{Y}(\mathrm{m})$ | Type | From | TO | Value | $\underline{\square}$ |
| 1003 | 3265.0 | 645.0 | Dist. | 3 | 1003 | 395.840 m | 0.005 m |
| 3 | 3660.0 | 630.0 | Dist. | 4 | 1003 | 464.103 m | 0.006 m |
| 4 | 3635.0 | 355.0 |  |  |  |  |  |

Table 5.3 Initial Data for Distance Intersection
points 3 and 4, and the observed distances and their standard deviations are listed in Table 5.3. These standard deviations could be a result of six determinations of the distance with a lightwave instrument such as the Hewlett-Packard 3800 (assuming $\sigma_{r}=\sigma_{\ell}$ ).

The observation equation in the form of formulae (4-9) and
(4-10) is
$\underset{(2,1)}{V}=\left[\begin{array}{l}\left(\left(x_{1003}^{0}-x_{3}\right)^{2}+\left(y_{1003}^{0}-y_{3}\right)^{2}\right)^{1 / 2}-\ell_{3,1003} \\ \left(\left(x_{1003}^{0}-x_{4}\right)^{2}+\left(y_{1003}^{0}-y_{4}\right)^{2,1 / 2}-\ell_{4,1003}\right.\end{array}\right]+$

$$
\left[\begin{array}{cc}
\frac{\left(x_{\left.1003^{-x_{3}}\right)}^{0}\right.}{\ell_{3,1003}^{0}} & \frac{\left(y_{1003}^{0}-y_{3}\right)}{\ell_{3,1003}^{0}}  \tag{5-11}\\
\frac{\left(x_{1003}^{0}-x_{4}\right)}{\ell_{4,1003}^{0}} & \frac{\left(y_{1003}^{0}-y_{4}\right)}{\ell_{4,1003}^{0}}
\end{array}\right] \quad\left[\begin{array}{l}
\delta x_{1003} \\
\\
\delta y_{1003}
\end{array}\right]
$$

Using the data from Table 5.3, the $A, P$ and $W$ metrices are

$$
A=\left[\begin{array}{cc}
-0.9992797 & 0.0379473 \\
-0.7870559 & 0.6168817
\end{array}\right], W=\left[\begin{array}{c}
-0.55498 \mathrm{~m} \\
6.00325 \mathrm{~m}
\end{array}\right]
$$

and, assumed the a priori variance factor $\sigma_{0}^{2}=1$,

$$
P=\left[\begin{array}{cc}
40000 & 0 \\
0 & 27777.777
\end{array}\right]
$$

where $A$ is unitless and $P$ has units of $m^{-2}$.
Using these matrices to evaluate $\hat{X}$ for the first iteration (again using equation (AII-1l)) yields

$$
\hat{x}=\left[\begin{array}{l}
-0.97203 \mathrm{~m} \\
-10.97179 \mathrm{~m}
\end{array}\right]
$$

which results in the least squares estimate of the coordinates being

$$
\mathrm{x}=\mathrm{x}^{\circ}+\hat{\mathrm{X}}=\left[\begin{array}{l}
3265.0  \tag{5-12}\\
645.0
\end{array}\right]+\left[\begin{array}{c}
-0.97203 \\
-10.97179
\end{array}\right]=\left[\begin{array}{c}
3264.028 \mathrm{~m} \\
634.028 \mathrm{~m}
\end{array}\right] .
$$

Using these parameter values as approximate coordinates now, and reevaluating the $A$ and $W$ matrices for the second iteration gives

$$
A=\left[\begin{array}{cc}
-0.9999483 & 0.01017244 \\
-0.7991728 & 0.6011013
\end{array}\right] \quad W=\left[\begin{array}{l}
0.15283 \\
0.09187
\end{array}\right]
$$

Using these matrices (as well as $P$ ) to compute $\hat{X}$ again using equation (AII-1l) gives

$$
\hat{x}=\left[\begin{array}{l}
0.15336 \mathrm{~m}  \tag{5-13}\\
0.05105 \mathrm{~m}
\end{array}\right]
$$

which, when added to this iterations' approximate coordinates (eq. (5-12)) gives the parameters from the second iteration as

$$
X=X^{0}+\hat{X}=\left[\begin{array}{l}
3264.028  \tag{5-14}\\
634.028
\end{array}\right]+\left[\begin{array}{l}
0.15336 \\
0.05105
\end{array}\right]=\left[\begin{array}{cc}
3264.181 & \mathrm{~m} \\
634.079 & \mathrm{~m}
\end{array}\right]
$$

Because the correction or solution vector $X$ was not insignificant (i.e. less than 0.001 m ) on the second iteration (eq. (5-13)), a third iteration is required. Thus, the parameters of the second iteration (eq. (5-14)) now becomes the approximate coordinates $x^{\circ}$ for the third iteration, and the $A$ and $W$ matrices are again computed. They are
$A=\left[\begin{array}{ll}-0.9999469 & 0.01030534 \\ -0.7990006 & 0.6013303\end{array}\right], W=\left[\begin{array}{c}0.3496 \cdot 10^{-5} \mathrm{~m} \\ 0.1905 \cdot 10^{-4} \mathrm{~m}\end{array}\right]$.
The solution vector $\hat{X}$ from this third iteration is

$$
\hat{x}=\left[\begin{array}{c}
0.00000 \\
-0.00003
\end{array}\right]
$$

which is less than 0.001 m , and thus insignificant. The solution has converged, and the final least squares estimates for the coordinates of point 1003 are given by equation (5-14).

The variance covariance matrix of the parameters computed according to equation (AII-16) is

$$
C_{x}=\left[\begin{array}{ll}
0.2972685 \cdot 10^{-4} & 0.40380781 \cdot 10^{-4} \\
0.40380781 \cdot 10^{-4} & 0.13924377 \cdot 10^{-3}
\end{array}\right]
$$

which results in a standard error ellipse of

$$
\begin{aligned}
\mathrm{a}_{\mathbf{s}} & =0.01235 \mathrm{~m} \\
\mathrm{~b}_{\mathbf{s}} & =0.00406 \mathrm{~m} \\
\theta & =18^{\circ} 12^{\prime} 11^{\prime \prime}
\end{aligned}
$$

Increasing the confidence lebel to $95 \%$ as in the previous two examples yields

$$
\begin{aligned}
& \mathrm{a}=0.030 \mathrm{~m} \\
& \mathrm{~b}=0.010 \mathrm{~m}
\end{aligned}
$$

Figure 5.6 illustrates the result.


Figure 5.6 95\% Confidence Ellipse for Distance Intersection

### 5.4 Angle Resection

The angle resection considered here is depicted in figure 5.7.
There are two angle observations $B_{1007,2,1}$ and $B_{1007,1,3}$, where the subscripts stand for stations 'at, from, to'. Point 1007 is the unknown station and points 1,2 and 3 are fixed.

Table 5.4 lists the point coordinates as well as the observations

| Station | X | $\underline{Y}$ | At | From | To | Value | $\underline{\square}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1007 | 3160.0 | 865.0 | 1007 | 2 | 1 | $23^{\circ} 13^{\prime \prime} 37: 33$ | 3:0 |
| 1 | 2640.0 | 1150.0 | 1007 | 1 | 3 | 175*36'21:70 | 3:5 |
| 2 | 2530.0 | 935.0 |  |  |  |  |  |
| 3 | 3660.0 | 630.0 |  |  |  |  |  |


Figure 5.7 Angle Resection
and their standard deviations. These standard deviations could result from 2 sets of observations with a l" theodolite or 8 to 10 sets with a 20" instrument (see Table 2.4).

The observation equation as developed in section 3.3 (eqs.
(3-26) and (3-27)) is
$\mathrm{V},\left[\begin{array}{c}\arctan \left(\frac{x_{1}-x_{1007}^{0}}{y_{1}^{-y_{1007}^{0}}}\right)-\arctan \left(\frac{x_{2}^{-x_{1007}^{\circ}}}{y_{2}^{-}-y_{1007}^{\circ}}\right)-B_{1007,2,1} \\ \arctan \left(\frac{x_{3}-x_{1007}^{\circ}}{y_{3}-y_{1007}^{\circ}}\right)-\arctan \left(\frac{x_{1}-x_{1007}^{\circ}}{y_{1}-y_{1007}^{\circ}}\right)-B_{1007,1,3}\end{array}\right]^{\prime \prime}$


Substituting the initial data from Table 5.4 into the above expressions results in

$$
A=\left[\begin{array}{cc}
-134.3056 & 23.32723 \\
329.0484 & 637.9743
\end{array}\right] \quad W=\left[\begin{array}{c}
-2: 34886 \\
3.317
\end{array}\right]
$$

and, assuming $\sigma_{0}^{2}=1$,

$$
P=\left[\begin{array}{cc}
0.11111 & 0 \\
0 & 0.08163
\end{array}\right]
$$

The units for $P$ are (") ${ }^{-2}$ and for $A$ are ("). $m^{-1}$.
The solution vector $\hat{X}$ resulting from this first iteration (computed by eq. (AII-ll)) is

$$
\hat{X}=\left[\begin{array}{c}
-0.01638 \mathrm{~m}  \tag{5-16}\\
0.00351 \mathrm{~m}
\end{array}\right],
$$

which gives the parameter vector as

$$
X=\hat{X}^{\circ}+\hat{X}=\left[\begin{array}{c}
3161.0  \tag{5-17}\\
865.0
\end{array}\right]+\left[\begin{array}{c}
-0.01688 \\
0.00351
\end{array}\right]=\left[\begin{array}{c}
3159.983 \mathrm{~m} \\
865.004 \mathrm{~m}
\end{array}\right]
$$

These parameter values are now taken as new approvimate coordinates $X^{\circ}$, the $A$ and $W$ matrices are recomputed to enable the second iteration value for $\hat{X}$ to be found. They are

$$
A=\left[\begin{array}{cc}
-134.3128 & 23.32934 \\
329.0485 & 637.9719
\end{array}\right], W=\left[\begin{array}{c}
0.6428 \cdot 10^{-4} \\
-0.4607 \cdot 10^{-5}
\end{array}\right],
$$

which results in a solution vector of

$$
\hat{x}=\left[\begin{array}{l}
0.00000 \\
0.00000
\end{array}\right]
$$

Thus, the solution has converged, and the final least squares estimate of the coordinates of point 1007 is given by equation (5-17).

The variance covariance matrix of the parameters computed according to equation (AII-16) is

$$
C_{x}=\left[\begin{array}{lll}
0.42099214 \cdot 10^{-3} & -0.21233829 \cdot 10^{-3} \\
-0.21233829 \cdot 10^{-3} & 0.13714132 \cdot 10^{-3}
\end{array}\right]
$$

winch gives the standard error ellipse computed by the equations of Appendix III as


Figure $5.8 \quad 95 \%$ Confidence Ellipse from Angle Resection

$$
\begin{aligned}
\mathrm{a}_{\mathrm{s}} & =0.02312 \mathrm{~m} \\
\mathrm{~b}_{\mathrm{s}} & =0.00486 \mathrm{~m} \\
\theta & =-61^{\circ} 52^{\prime} 46^{\prime \prime}
\end{aligned}
$$

Increasing the confidence level to 95\% yields

$$
\begin{aligned}
& \mathrm{a}=0.057 \mathrm{~m} \\
& \mathrm{~b}=0.012 \mathrm{~m}
\end{aligned}
$$

The resultant confidence ellipse is depicted in Figure 5.8.
The variance covariance matrix for the parameters $C_{x}$ would be difficult to compute using the direct (explicit) mathematical model for an angle resection, mainly due to the complicated partial derivatives $\frac{\partial X}{\partial L}$ needed for the Jacobian of transformation in the covariance law (cf. Thomson et al. [1978], Appendix II). Thus, it is seen that the method of least squares offers a consistently convenient method of error propogation.

### 5.5 Open Traverse

Figure 5.9 shows the open ended traverse considered in this example. As in section 5.1, angle and distance observations are combined in the same model. There are three angles and three distances for a total of six observations, and three unknown points, or six unknown coordinates. Thus, $n=u$ and this is a unique case.

The fixed coordinates for points 1 and 2, the approximate coordinates for points 1001, 1002 and 1003, and the observations and their standard deviations are listed below in Table 5.5. The observational standard deviations are computed according to section 2.1.


Figure 5.9 Open Traverse

| Coordinates of Points |  |  |  | Observations on the Mapping Plane |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Station | $\underline{X}$ | $\underline{Y}$ | Type | At | From | To | Value | $\underline{\sigma}$ |
| 1001 | 2950.0 | 1.160 .0 | Dist. | 1 | 1 | 1001 | 307.997 m | 0.010 m |
| 1002 | 3280.0 | 1145.0 | Dist. | 1001 | 1001 | 1002 | 330.355 m | 0.012 m |
| 1003 | 3265.0 | 64.5 .0 | Dist. | 1002 | 1002 | 1003 | 500.243 m | 0.011 m |
| 1 | 2640.0 | 1160.0 | Angle | 1 | 2 | 1001 | $243^{\circ} 56{ }^{\prime} 55: 16$ | 3:5 |
| 2 | 2530.0 | 935.0 | Angle | 1001 | 1 | 1002 | 182*36'5: 44 | 4:0 |
|  |  |  | Angle | 1002 | 1001 | 1003 | 2690'51:39 | 3.0 |

Table 5.5 Initial Data for Open Traverse

Combining the mathematical models developed in section 3.3 and chapter 4, the observation equation for this example is

$$
\underset{(6,1)}{v}=\underset{(6,1)}{W}+\underset{(6,6)}{A} \underset{(6,1)}{\hat{X}}
$$

or evaluating the matrices individually,
(5-19)

Note that $\rho "$ has been omitted from the non-zero elements of the last three rows of $A$. The first three rows of $A$ are unitless, whereas the last three have units of (") $\mathrm{m}^{-1}$. The solution vector has the form

$$
\hat{x}=\left[\begin{array}{l}
\delta x_{1001}  \tag{5-20}\\
\delta y_{1001} \\
\delta \mathrm{x}_{1002} \\
\delta y_{1002} \\
\delta \mathrm{x}_{1003} \\
\delta y_{1003}
\end{array}\right]
$$

Using the initial values from Table 5.5 yields the following
$\mathrm{A}, \mathrm{W}$ and P matrices:
$A=\left[\begin{array}{cccccc}100.0 & 0.0 & 0.0 & 0.0 & 0.0 & 0.0 \\ -0.9989685 & 0.04540766 & 0.9989685 & -0.04540766 & 0.0 & 0.0 \\ 0.0 & 0.0 & 0.02998651 & 0.9995503 & -0.02998651 & -0.9995503 \\ 0.0 & -665.3703 & 0.0 & 0.0 & 0.0 & 0.0 \\ 28.35255 & 1289.126 & -28.35255 & -623.7561 & 0.0 & 0.0 \\ -28.35255 & -623.7561 & 440.5112 & 611.3914 & -412.1587 & 12.36476\end{array}\right]$
(note that $\rho "$ has been multiplied onto the appropriate elements for this A matrix)
$\mathrm{W}=\left[\begin{array}{c}2.0033911 \mathrm{~m} \\ -0.01411715 \mathrm{~m} \\ -0.01848697 \mathrm{~m} \\ -7.7426476 \\ 3.7866587 \\ 5.4711834\end{array}\right], \mathrm{P}=\left[\begin{array}{cccccc}10000.0 & 0 & 0 & 0 & 0 & 0 \\ 0 & 6944.44 & 0 & 0 & 0 & 0 \\ 0 & 0 & 8264.46 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0.08163 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0.0625 & 0 \\ 0 & 0 & 0 & 0 & 0 & 0.1111\end{array}\right]$.

The solution vector $\hat{X}$ computed (using eq. (AII-ll)) from these first iteration matrices is

$$
\hat{X}=\left[\begin{array}{l}
-2.00339 \mathrm{~m} \\
-0.01164 \mathrm{~m} \\
-1.98958 \mathrm{~m} \\
-0.1861 \mathrm{~m} \\
-1.98646 \mathrm{~m} \\
-0.03720 \mathrm{~m}
\end{array}\right]
$$

which results in the parameters X being

$$
X=x^{\circ}+\hat{X}=\left[\begin{array}{l}
2950.0 \\
1160.0 \\
3280.0 \\
1145.0 \\
3265.0 \\
645.0
\end{array}\right]+\left[\begin{array}{c}
-2.00339 \\
-0.01164 \\
-1.98958 \\
-01861 \\
-1.98646 \\
-0.03720
\end{array}\right]=\left[\begin{array}{c}
2947.997 \mathrm{~m} \\
1159.988 \mathrm{~m} \\
3278.010 \mathrm{~m} \\
1144.981 \mathrm{~m} \\
3263.014 \mathrm{~m} \\
644.963 \mathrm{~m}
\end{array}\right]
$$

Using these parameters as new approximate coordinates and recomputing $A$ and $W$ to get a new solution vector $\hat{X}$ results in

$$
\hat{x}=\left[\begin{array}{l}
0.00000 \\
0.00008 \\
0.00000 \\
0.00007 \\
0.00000 \\
0.00007
\end{array}\right] .
$$

All elements of $X$ are less than 0.001 m , and thus the solution has converged. Thus, the final least squares estimate of the coordinates of points 1001, 1002 and 1003 is given by equation (5-21).

The variance covariance matrix of the parameters assuming the a posteriori variance factor known (eq. (AII-16)) is


Table 5.6 shows the standard and $95 \%$ station error ellipses as well as the relative error ellipses between the unknown points. Again, it is assumed that $\sigma_{0}^{2}$ is known, and the $c$ factor is 2.45. Station Ellipses

Standard 95\%

| Station | $\frac{a_{s}(m)}{}$ |  | $b_{s}(m)$ | $\theta$ | $\frac{a(m)}{}$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1001 | 0.010 | 0.005 | $-89^{\circ} 59^{\prime} 52^{\prime \prime}$ | 0.025 | 0.013 |
| 1002 | 0.016 | 0.013 | $-88^{\circ} 42^{\prime} 45^{\prime \prime}$ | 0.038 | 0.031 |
| 1003 | 0.024 | 0.014 | $62^{\circ} 4^{\prime} 46^{\prime \prime}$ | 0.058 | 0.034 |

## Relative Ellipses

|  | Standard |  | $95 \%$ |  |  |
| :--- | :--- | :--- | :---: | :---: | :---: | :---: |
| Station | $a_{s}(m)$ | $b_{s}(m)$ | $\theta$ | $a(m)$ | $b(m)$ |
| 1001 to 1002 | 0.012 | 0.009 | $-87^{\circ} 23^{\prime} 47^{\prime \prime}$ | 0.030 | 0.021 |
| 1002 to 1003 | 0.015 | 0.011 | $-88^{\circ} 16^{\prime} 56^{\prime \prime}$ | 0.036 | 0.027 |

Table 5.6 Station and Relative Ellipses for Open Traverse

The traverse with both station and relative 95\% confidence ellipses is shown in Figure 5.10.


Figure 5.10 95\% Confidence Ellipses for Open Traverse

## 6. SOLUTION OF OVERDETERMINED CASES

As already mentioned in the introduction, overdetermined cases include any network in which the number of observations $n$ is greater than the number of unknowns $u$. Any overdetermined network has more than one unique solution for the coordinates of the unknown points. Thus, the best solution based on all of the available information must be found. This is accomplished by the method of least squares (see Appendix II) which gives the minimum weighted sum of squares of the residuals (corrections to the observations), i.e. $\mathrm{V}^{\mathrm{T}} \mathrm{PV}=$ minimum. A simple example is the least squares line fitting technique shown in Figure 6.1. The observations are the $y$ coordinate (horizontal axis $t$ is known) and the unknowns are the slope of the line $a$ and the $y$ intersect $b$ (i.e. $y=a t+b)$. The least squares technique minimizes the sum square of the residuals, which in this case is the distance parallel to the $y$ axis from the observation point to the line.


Figure 6.1 Least Squares Line Fitting

Besides having at least as many observations as unknowns, a horizontal network must also have certain other basic information before it can be solved. A network must have scale, orientation, and one known position. Orientation is introduced by observing an azimuth, or by "fixing" one point along with the $x$ or $y$ coordinate of another point. The scale is provided by measuring at least one distance and including it in the network, or by "fixing" at least two points. At least one point must be assumed known to provide the minimum position information for a horizontal network. If two or more points are assumed known, then the scale and orientation are inherent as well. If this minimum information is not provided, then usually the normal equations matrix $N=A^{T} P A$ is singular, and its inverse cannot be found.

Classical horizontal networks were usually measured by triangulation methods; i.e. having only angular observations between stations. The scale was introduced by baselines measured with invar wires or tapes. Since the introduction of EDM equipment, trilateration networks composed mainly of distance observations have been measured, with the orientation provided by azimuth observations. Modern day horizontal networks are usually composed of a mixture of both angular and distance observations, and the phrase triangulateration network has been coined to characterize them. A traverse is a simple example of a triangulateration network. The following examples show some of the various network types for overdetermined cases.

### 6.1 Closed Traverse

Figure 6.2 depicts the closed traverse considered in this example.


There are four unknown points and four fixed points. The $x, y$ coordinates of the four unknown points along with the six orientation unknowns (see section 3.2) gives a total of fourteen unknown parameters. It is seen from Figure 6.2 and Table 6.1 that there are a total of seventeen observations, and the degrees of freedom is, therefore, three. Since the degrees of freedom is greater than zero, this problem is overdetermined.

| Coordinates of Points |  |  | Observations on the Mapping Plane |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Station | X | $\underline{Y}$ | Type | From | To | Value | $\underline{\sigma}$ |
| 1003 | 3265.0 | 645.0 | Dir. | 1 | 1006 | $0^{\circ} 00 \cdot 00: 0$ | 2:0 |
| 1004 | 3570.0 | 915.0 | Dir. | 1 | 2 | $66^{\circ} 1$ 1 $1: 0$ | 2:0 |
| 1006 | 2820.0 | 945.0 | Dir. | 1006 | 1007 | $0^{\circ} 00^{\prime} 00: 0$ | 2:0 |
| 1007 | 3160.0 | 865.0 | Dir. | 1006 | 1. | 216*53'42.0 | 2.0 |
| 1 | 2640.0 | 1160.0 | Dir. | 1007 | 1003 | $0^{\circ} 00^{\prime} 00.0$ | 2.0 |
| 2 | 2530.0 | 935.0 | Dir. | 1007 | 1006 | 128* $41^{\prime} 52.0$ | 2.0 |
| 3 | 3660.0 | 630.0 | Dir. | 1003 | 1004 | 0"00'00:0 | 2.0 |
| 4 | 3635.0 | 355.0 | Dir. | 1003 | 1007 | 286 ${ }^{\circ} 1^{\prime} 57.0$ | 2:0 |
|  |  |  | Dir. | 1004 | 3 | 0"00'00:0 | 2:0 |
|  |  |  | Dir. | 1004 | 1003 | 6548'6:0 | 2.0 |
|  |  |  | Dir. | 3 | 4 | 0"00'00:0 | 2.0 |
|  |  |  | Dir. | 3 | 1004 | 157* $25^{\prime} 5.0$ | 2:0 |
|  |  |  | Dist. | 3 | 1004 | 301.200 m | . 01 m |
|  |  |  | Dist. | 1 | 1006 | 279.747 m | .01 m |
|  |  |  | Dist. | 1006 | 1007 | 348.982 m | . 01 m |
|  |  |  | Dist. | 1007 | 1003 | 243.623 m | . 01 m |
|  |  |  | Dist. | 1003 | 1004 | 408.310 m | .01 m | Table 6.1 Initial Data for Closed Traverse

Again, it is assumed that the observations have already been reduced to the mapping plane, and that the standard deviations have been computed according to section 2.1 .

Combining the mathematical models developed in section 3.2 and chapter 4, the observation equation takes the form

$$
\underset{(17,1)}{\mathrm{V}}=\underset{(17,1)}{\mathrm{W}}+\underset{(17,14)}{\mathrm{A}} \underset{(14,1)}{\mathrm{X}}
$$

The $W, A$ and $\hat{X}$ matrices are evaluated exactly as for the unique case examples in chapter 5. The primary difference between the unique case and this overdetermined case is that the residuals $V$ are no longer zero. The $W$ matrix is computed exactly as in equations (3-15) (directions)
and (4-9) (distances). For a direction, the $W$ matrix element is

$$
\left.W_{i j}=\arctan \frac{x_{j}^{0}-x_{i}^{0}}{y_{j}^{0}-y_{i}^{0}}\right)-z_{i}^{0}-d_{i j}
$$

Taking the direction $d_{1,1006}$ as an example and using the initial approximate coordinates of Table 6.1 yields

$$
\begin{aligned}
& \mathrm{w}_{\mathrm{d}_{1,1006}}=\arctan \left(\frac{2820.0-2640.0}{945.0-1160.0}\right)-\arctan \left(\frac{2820.0-2640.0}{945.0-1160.0}\right)-0^{\circ} 00^{\prime} 00: 0 \\
& \mathrm{~W}_{\mathrm{d}_{1,1006}}=0: 0
\end{aligned}
$$

The approximate value for the orientation unknown $Z_{i}^{\circ}$ is always computed as the azimuth between the from and to stations of the first direction in the set of directions. This causes the $W$ matrix element for the first direction of a set to be zero assuming that the directions are reduced such that the first direction of the set always has an observed value of $0^{\circ} 00^{\prime} 00 \%$. In this case, $d_{1,1006}$ is the first direction of the set. For the second direction of the set $d_{1,2}$, the $W$ matrix element is $\mathrm{W}_{\mathrm{d}_{1,2}}=\arctan \left(\frac{2530.0-2640.0}{935.0-1160.0}\right)-\arctan \left(\frac{2820.0-2640.0}{945.0-1160.0}-66^{\circ} 1 \cdot 1 \%^{\prime} 0 \quad\right.$,
$W_{d_{1,2}}=206^{\circ} 03^{\prime} 12: 58-140^{\circ} 03^{\prime} 49: 02-66^{\circ} 1^{\prime} 1: 0$, $W_{d_{1,2}}=-97: 44$.

The distance $W$ matrix elements are computed as in chapter 5, for example $W_{\ell_{3,1004}}=\left((3570.0-3660.0)^{2}+(915.0-630.0)^{2}\right)^{1 / 2}-301.200 \quad$,
$\mathrm{W}_{\ell, 1004}=298.873-301.200$,
$W_{\ell_{3,1004}}=-2.327 \mathrm{~m}$.
After computing all of the elements of $W$ using the initial coordinates in Table 6.1, W is

$$
\begin{aligned}
\underset{(1,17)}{\mathrm{W}^{T}=} & (0: 0,-97: 44,0: 0,-258: 85,0: 0,203: 80,0: 0,-106: 71,0: 0,745: 82, \\
& 0: 0,-496: 99,-2.327 \mathrm{~m}, ~ 0.654 \mathrm{~m}, ~ 0.303 \mathrm{~m}, ~ 0.149 \mathrm{~m},-0.971 \mathrm{~m}),
\end{aligned}
$$

where the transpose of $W$ is given for ease of writing.
The A matrix is computed via formulae (3-15) and (4-9) as well.
For example, the direction $d_{1,1006}$ has nonzero elements -pa ${ }_{1,1006^{\prime}}$
$-0 b_{1,1006^{\prime}}-1$ in columns 5,6 and 9 of row 1 of A. Distance $\ell_{3,1004}$ has nonzero elements $-e_{3,1004}$ and $f_{3,1004}$ in columns 3 and 4 of row 13 of A. The entire $A$ matrix is given to five significant digits below.
where $Z_{1}, Z_{2}, Z_{3}, Z_{4}, Z_{5}$ and $Z_{6}$ refer to the orientation unknowns at stations 1, 1006, 1007, 1003, 1004, and 3, respectively.

The $P$ matrix is diagonal of size $17 \times 17$ with the inverse of the standard deviations (Table 6.1) squared on the diagonal. It is

Computing the solution vector $\hat{X}$ via equation (AII-1l) and the
above $W, A$ and $P$ matrices yields (in units of metres)

$$
\begin{aligned}
\hat{\mathrm{x}}^{\mathrm{T}}= & (-0.40492,1.4367,-0.01428,2.43956,-0.32453,0.58429,-0.49373, \\
& 1.23052)
\end{aligned}
$$

for this first iteration. The orientation unknowns are not given as they are considered as nuisance parameters. This results in the parameters

$$
\begin{aligned}
x^{T}=x^{o^{T}}+\hat{X}^{T}= & (3264.595,646.437,3569.986,917.440,2819.675,945.584, \\
& 3159.506,866.231),
\end{aligned}
$$

where $X^{\circ}$ are the approximate coordinates of Table 6.1.
These parameters X are taken as new approximate coordinates $\mathrm{X}^{\circ}$, and the $A$ and $W$ matrices are reevaluated using them. Equation (AII-1l) is again used to compute the solution vector for the second iteration as

$$
\begin{aligned}
\hat{\mathbf{x}}^{\mathrm{T}}= & (0.00451,-0.00163,0.00500,0.00123,-0.00103,0.00361, \\
& -0.00111),
\end{aligned}
$$

which results in the parameter vector

$$
\begin{aligned}
x^{T}= & x^{\circ} \mathrm{T}+\hat{X}^{T}=(3264.600,646.435,3569.991,917.441,2819.677,945.583 \\
& \quad 3159.510,866.229)
\end{aligned}
$$

These parameters are taken as new approximate coordinates $\mathrm{X}^{\circ}$ again, and the third solution vector is (after reevaluating $A$ and $W$ )

$$
\begin{aligned}
\hat{\mathrm{X}}^{\Gamma}= & (0.00000,0.00000,0.00000,0.00000,0.00000,0.00000,0.00000, \\
& 0.00000)
\end{aligned}
$$

and the solution has converged. The final parameters ( $x, y$ coordinates of the unknown points) are thus given by equation (6-1).

The variance covariance matrix of the parameters is computed via equation (AII-16) assuming the a priori variance factor is l. To five significant digits, $C_{x}$ is


The $95 \%$ (c factor $=2.45$ ) station and relative error ellipses computed by the equations of Appendix III from the above $C_{x}$ matrix are listed in Table 6.2, and plotted in Figure 6.3.


Figure 6.3 Plot of 95\% Error Ellipses for Closed Traverse

| Station | $a(m)$ | b (m) | $\theta$ | Station | Station | $\mathrm{a}(\mathrm{m})$ | $\mathrm{b}(\mathrm{m})$ | $\theta$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1003 | . 021 | . 019 | -9³9'53' | 1003 | 1004 | . 021 | . 012 | $69^{\circ} 24^{\prime} 07 \prime$ |
| 1004 | . 022 | . 008 | $-21^{\circ} 19^{\prime} 45^{\prime \prime}$ | 1003 | 1007 | . 022 | . 008 | $-30^{\circ} 39^{\prime} 26^{\prime \prime}$ |
| 1006 | . 021 | . 007 | $-40^{\circ} 51{ }^{\prime \prime} 12^{\prime \prime}$ | 1007 | 1006 | . 020 | . 010 | $-67^{\circ} 07^{\prime} 34^{\prime \prime}$ |
| 1007 | . 022 | . 016 | $-39^{\circ} 39^{\prime} 44^{\prime \prime}$ |  |  |  |  |  |

Table 6.2 95\% Error Ellipses for Closed Traverse

The residuals $V$ computed by equation (AII-17) are

$$
\begin{aligned}
\left.\mathrm{v}^{\mathrm{T}} \mathrm{~T}, 17\right) & (2.07,-2.07,1.48,-1.48,0.91,-0.91,0.41,-0.41,0.45,-0.45 \\
& -0.14,0.14,0.004,0.000,-0.007,0.003,-0.011)
\end{aligned}
$$

where the units are arcseconds for the first 12 residuals (i.e. for directions), and metres for the last five (distance residuals). Using these residuals and the $P$ matrix computed earlier, the a posteriori variance factor $\hat{\sigma}_{0}^{2}$ is computed via equation (AII-18) as

$$
\begin{equation*}
\hat{\sigma}_{0}^{2}=\frac{v^{T} \mathrm{PV}}{\mathrm{df}}=1.9214 \tag{6-3}
\end{equation*}
$$

Both the residuals and a posteriori variance factor are used in chapter 9 for the post analysis procedures.

### 6.2 Network

Figure 6.4 shows the network considered in this example. It consists of 10 unknown stations and one fixed station with 38 directions, 17 distances, and 2 azimuths observed. Accounting for the 11 orientation unknowns, then, the degrees of freedom is 26 . The initial point coordinates are listed in Table 6.3 and the observations and their standard deviations are given in Table 6.4.


| Station | $\underline{X}$ | $\underline{Y}$ | $\underline{\text { Station }}$ | $\underline{X}$ | $\underline{Y}$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 2640.0 | 1160.0 | 1003 | 3265.0 | 645.0 |
| 2 | 2530.0 | 935.0 | 1004 | 3570.0 | 915.0 |
| 3 | 3660.0 | 630.0 | 1005 | 2770.0 | 655.0 |
| 4 | 3635.0 | 355.0 | 1006 | 2820.0 | 945.0 |
| 1001 | 2950.0 | 1160.0 | 1007 | 3160.0 | 865.0 |
| 1002 | 3280.0 | 1145.0 |  |  |  |

Table 6.3 Initial Coordinates for Network Stations

Combining the mathematical models of sectiors 3.1 and 3.2 and chapter 4, the matrix form of the observation equations is

$$
\underset{(57,1)}{V}=\underset{(57,1)}{\mathrm{W}}+\underset{(57,31)}{\mathrm{A}} \underset{(31,1)}{ } \hat{\mathrm{X}} .
$$

Using the same techniques as in the previous example in section 6.1, the $W$ matrix is computed as

$$
\begin{aligned}
& W^{T}=(327.58,928: 78,0.191 \mathrm{~m}, ~ 0.003 \mathrm{~m}, ~ 0.010 \mathrm{~m},-0.613 \mathrm{~m},-0.48 \mathrm{~m} \text {, } \\
& (1,57) \\
& 0.827 \mathrm{~m},-0.588 \mathrm{~m}, ~ 0.448 \mathrm{~m}, ~ 0.287 \mathrm{~m}, ~ 0.571 \mathrm{~m}, ~ 0.223 \mathrm{~m}, ~ 0.242 \mathrm{~m} \text {, } \\
& -1.011 \mathrm{~m},-0.198 \mathrm{~m},-0.112 \mathrm{~m},-0.603 \mathrm{~m},-0.774 \mathrm{~m}, ~ 0.0,-215.98,-388.51 \text {, } \\
& -337: 42,0.0,-344: 78,-371: 73,-527: 22,0.0,-59: 62,-240: 83, \\
& -369: 48,0: 0,544: 73,-304: 18,0: 0,-497: 86,-446: 99,-367: 48, \\
& \text { 0:0, -161:"91, 60:"93, 41:22, 0:0, -67:31, -1361:01, -1193:51, } \\
& \text {-276:.02, 0."0, 236:"22, -1184:82, 0!0, -204":24, 0.0, -688:.87, } \\
& \text {-582.:99, 0."0, 522.:36) , }
\end{aligned}
$$

where the order of the elements of $W$ is the same as the order of the observations in Table 6.4. The A matrix is too large to put conveniently on one page, so the nonzero elements of it are given on the following

| Type | From | To | Value | $\underline{\sigma}$ | Type | From | To | Value | $\underline{\sigma}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Az | 1 | 2 | 205*57'45:0 | 5.0 | Dir | 1006 | 1005 | $158^{\circ} 41^{\prime} 24: 0$ | 2.0 |
| Az | 1003 | 1004 | 48ำ $3^{\prime} 31: 0$ | 5:0 | Dir | 1006 | 1 | 289 ${ }^{\circ} 00^{\prime} 25.0$ | 2:0 |
| Dist | 1003 | 3 | 395.094 m | . 01 m | Dir | 1005 | 1006 | $0^{\circ} 00^{\prime} 00.0$ | 2:0 |
| Dist | 3 | 4 | 276.131 m | . 01 m | Dir | 1005 | 1003 | $81^{\circ} 13^{\prime} 25.0$ | 2.0 |
| Dist | 1 | 2 | 250.440 m | . 01 m | Dir | 1005 | 1 | 33551'58:0 | 2:0 |
| Dist | 4 | 1003 | 470.719 m | . 01 m | Dir | 1007 | 1002 | $0^{\circ} 00^{\prime} 00.0$ | 2:0 |
| Dist | 2 | 1005 | 369.262 m | . 01 m | Dir | 1007 | 1003 | 131"25'33.0 | 2:0 |
| Dist | 1 | 1001 | 309.173 m | . 01 m | Dir | 1007 | 1006 | 260 $08^{\prime} 58.0$ | 2:0 |
| Dist | 1 | 1005 | 522.052 m | . 01 m | Dir | 1007 | 1001 | $301^{\circ} 27^{\prime} 28.0$ | 2.0 |
| Dist | 1 | 1006 | 279.953 m | . 01 m | Dir | 1002 | 1004 | $0^{\circ} 00^{\prime} 00.0$ | 2.0 |
| Dist | 1001 | 1006 | 250.960 m | . 01 m | Dir | 1002 | 1003 | $53^{\circ} 20^{\prime} 43.0$ | 2.0 |
| Dist | 1001 | 1002 | 329.770 m | .01 m | Dir | 1002 | 1007 | 7445'49:0 | 2.0 |
| Dist | 1001 | 1007 | 361.889 m | . 01 m | Dir | 1002 | 1001 | 144* $10^{\prime} 23^{\prime}$ ! 0 | 2:0 |
| Dist | 1006 | 1007 | 349.043 m | . 01 m | Dir | 1003 | 1002 | 0"00'00:0 | 2:0 |
| Dist | 1006 | 1005 | 295.290 m | . 01 m | Dir | 1003 | 1004 | 46* $7^{\prime}$ '01:0 | 2:0 |
| Dist | 1007 | 1002 | 304.829 m | . 01 m | Dir | 1003 | 3 | 90으'04:0 | 2.0 |
| Dist | 1007 | 1003 | 243.884 m | . 01 m | Dir | 1003 | 4 | $126^{\circ} 42^{\prime} 07: 0$ | 2.0 |
| Dist | 1002 | 1004 | 370.738 m | . 01 m | Dir | 1003 | 1007 | $332{ }^{\circ} 50^{\prime} 40.0$ | 2:0 |
| Dist | 1003 | 1004 | 408.113 m | . 01 m | Dir | 1004 | 1003 | $0^{\circ} 00^{\prime} 00 \cdot 0$ | 2:0 |
| Dir | 1 | 1001 | $0^{\circ} 0^{\prime} 00: 0$ | 2.0 | Dir | 1004 | 1002 | 79*52'09:0 | 2.0 |
| Dir | 1 | 1006 | $50^{\circ} 07!25^{\prime \prime} 0$ | 2:0 | Dir | 1004 | 3 | 294*19'13:0 | 2.0 |
| Dir | 1 | 1005 | $75^{\circ} 40^{\prime} 19.0$ | 2:0 | Dir | 2 | 1 | $0^{\circ} 00 \cdot 00: 0$ | 2.0 |
| Dir. | 1 | 2 | 116 $08^{\prime} 50.0$ | 2.'0 | Dir | 2 | 1005 | $113^{\circ} 24^{\prime} 07!0$ | 2:0 |
| Dir | 1001 | 1002 | $0^{\circ} 0^{\prime} 00: 0$ | 2.0 | Dir | 3 | 4 | $0^{\circ} 00^{\prime} 00 \cdot 0$ | 2.0 |
| Dir | 1001 | 1007 | $52^{\circ} 02^{\prime} 51: 0$ | 2.0 | Dir | 3 | 1003 | $87^{\circ} 10^{\prime} 18: 0$ | 2:0 |
| Dir | 1001 | 1006 | 118³9'36.0 | 2.0 | Dir | 3 | 1004 | $157^{\circ} 26^{\prime} 31: 0$ | 2.0 |
| Dir | 1001 | 1 | 177032'38:0 | 2.0 | Dir | 4 | 1003 | $0^{\circ} 00^{\prime} 00 \cdot 0$ | 2.0 |
| Dir | 1006 | 1001 | $0^{\circ} 0^{\prime} 00 \cdot 0$ | 2.'0 | Dir | 4 | 3 | $56^{\circ} 57^{\prime} 38.0$ | 2.0 |
| Dir | 1006 | 1007 | $72^{\circ} 05^{\prime} 52.0$ | 2.0 |  |  |  |  |  |

Table 6.4 Observations on the Mapping Plane for the Network
two pages. Note that the row number is given under the observation type, and the column numbers are given between dashes (e.g.-18-) over top of the actual number. Station 1 is fixed, and thus there are no columns appearing in $A$ for it. All elements resulting from azimuth or direction observations are multiplied by $\rho$ " as was done in section 6.1.

The $P$ matrix is diagonal of size $57 \times 57$ with the inverse of the standard deviations (see Table 6.4) squared on the diagonal. The first two diagonal elements are 0.04, the next 17 are 10000.0 , and the last 38 elements are 0.25. $P$ thus takes the shape


Using the above $A, P$ and $W$ matrices in equation (AII-11) to compute the solution vector for the first iteration yields
$\hat{\mathrm{X}}^{\mathrm{T}} \mathbf{2 0}^{\mathrm{T}}=(0.3616,-0.1768,0.8559,1.6254,1.2828,1.5808,-0.8268,1.0078$, $-1.3181,2.9488,1.0736,2.3229,0.4415,4.2027,0.8414,-0.3917$, $0.1877,0.7411,0.2573,2.0613)$,
in units of metres. The orientation unknowns are not given as they are not required. Adding this solution to the approximate coordinates of

| Azputh |  |  |  |
| :---: | :---: | :---: | :---: |
|  |  |  |  |
| distance |  | ${ }^{3} .99927970{ }^{3-0}(x, y)$ |  |
| ${ }_{4}^{\text {distance }}$ | ${ }_{0}^{3} .40535750-02^{30} \times 1 \times 10.95589320+00$ |  |  |
| distance |  |  |  |
| distance |  |  |  |
| distiance | ${ }_{-0}^{2} .650791^{-10}+00^{(x, r)} 0.759256000^{2-00}$ |  |  |
| otistance | ${ }_{-0.1}^{1} \cdot 10000000+01^{(\overline{X, Y}} 0.0{ }^{\text {FIXED }}$ |  |  |
| disgiance |  |  |  |
| Disitance | $-1.64193670+00^{(x, Y)}$ O. $7667 \times 15770+00$ |  |  |
| distance | ${ }_{0}^{1091} 51741930+00^{7-}(x, r) .35573200{ }^{-8-00}$ |  |  |
| distitance | ${ }_{-0.9918963550}^{100}+00^{(x, r)} 0.45407660-01$ |  |  |
| distance | $-1009199311^{7-0}+00^{(x, r)} 0.814665^{8-0}+00$ |  |  |
| distiance | $-10066^{10} 44^{1720}+00^{(x, y)} 0.22903930+00$ |  |  |
| OISIANCE |  |  |  |
|  |  |  |  |
| $\frac{\text { Oistance }}{17}$ |  |  |  |
| ${ }_{\text {distance }}$ |  | $1004{ }^{10.78349770}+00^{-13}(x, y) .0 .62139470+00$ |  |
| distance | ${ }_{-0.74876220}^{10} 0$ |  |  |
| $\mathrm{OI}_{20} \mathrm{O}_{\text {ection }} 1$ |  |  | -0.20 ${ }^{-21-01}$ |
| - Direction 2 |  |  | ${ }^{-0.10}$ |
| direction 3 |  | ${ }_{-0.305306050}^{10}+03^{(x, y)}-0.98609630-02$ | -0.20 $0^{-21-01}$ |
|  |  |  | -0.10 ${ }^{-21}+01$ |
| $\mathrm{dipection}^{24}$ |  |  | $-0.20{ }^{22-01}$ |
| - 25 2fection 2 |  |  |  |
| ${ }_{\text {direction }}{ }^{\text {d }}$ |  |  | ${ }_{-0.10}^{-220}$ |
| - dikectich-4 |  |  | $-0.10^{22-01}$ |
| $\mathrm{Dighectila}_{28}$ |  |  | $-0.10{ }^{23}+01$ |
| difiction 2 | 2. $0_{0}^{1006} 135255500^{-17-0} \frac{(x, y)}{0.57463630-03}$ |  | -0.10 ${ }^{23-01}$ |
|  |  |  | $-0.80{ }^{-23-01}$ |



A matrix, rows 31-57

Table 6.3 results in the parameter vector

$$
\begin{aligned}
& \mathrm{X}^{\mathrm{T}}=\mathrm{X}^{\circ}{ }^{\mathrm{T}}+\hat{\mathrm{X}}^{\mathrm{T}}=(2530.362,934.823,3660.856,631.625,3636.283,356.581, \\
& 2949.173,1161.008,3278.682,1147.949,3266.074,647.323, \\
& 3570.442,919.203,2770.841,654.608,2820.188,945.741,3160.257, \\
&867.061) .
\end{aligned}
$$

Taking these parameters as new approximate coordinates $\mathrm{X}^{\circ}$, reevaluating the $A, P$ and $W$ matrices, and computing a second iteration solution vector via equation (AII-ll) results in

$$
\begin{aligned}
\hat{x}^{T}= & (-0.00006,-0.00004,-0.00898,-0.00012,-0.00821,0.00147,-0.00136, \\
& -0.00253,-0.00649,-0.00518,-0.00375,-0.00082,-0.00707,0.00096, \\
& 0.00073,-0.00062,-0.00155,-0.00031,-0.00369,-0.00124),
\end{aligned}
$$

which yields a second iteration parameter vector of

$$
\begin{aligned}
\mathrm{x}^{\mathrm{T}}= & \mathrm{x}^{\circ \mathrm{T}}+\mathrm{x}^{\mathrm{T}}=(2530.362,934.823,3660.847,631.625,3636.275,356.582, \\
& 2949.172,1161.005,3278.675,1147.944,3266.070,647.322,3570.434, \\
& 919.204,2770.842,654.608,2820.186,945.741,3160.254,867.060) .
\end{aligned}
$$

Using these parameters as new approximate coordinates results in a zero vector for the third iteration solution vector, and thus the parameters in equation $(6-4)$ are the final adjusted coordinates.

The variance covariance matrix $C_{x}$ of the parameters is computed by equation (AII-16) (assuming a priori variance factor equals l) and is given on the following page. Note that since the $C_{x}$ matrix is too large to display in a normal fashion, it has been printed in rows of six columns at a time.

The 95\% station and relative error ellipses are computed using the equations of Appendix III. They are listed below in Table 6.5 and plotted in Figure 6.5.


STATION ELLIPSES
RELATIVE ELLIPSES

| Station | a (m) | $\mathrm{b}(\mathrm{m})$ | Station Station |  |  | a (m) | b (m) |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2 | . 016 | . 011 | $16^{\circ} 07^{\prime} 25^{\prime \prime}$ | 2 | 1005 | . 020 | . 014 | $36^{\circ} 20^{\prime \prime} 31{ }^{\prime \prime}$ |
| 3 | . 052 | . 023 | $26^{\circ} 14^{\prime \prime} 12^{\prime \prime}$ | 1005 | 1006 | . 015 | . 011 | -69*5 ${ }^{\prime}$ 41" |
| 4 | . 058 | . 025 | $38^{\circ} 28^{\prime \prime} 49^{\prime \prime}$ | 1001 | 1006 | . 012 | . 008 | $-63^{\circ} 44^{\prime} 52^{\prime \prime}$ |
| 1001 | . 015 | . 011 | $-4^{\circ} 27^{\prime} 10^{\prime \prime}$ | 1001 | 1002 | . 016 | . 011 | $2^{\circ} 01^{\prime} 21^{\prime \prime}$ |
| 1002 | . 030 | . 017 | $-0^{\circ} 28^{\prime} 49^{\prime \prime}$ | 1001 | 1007 | . 018 | . 011 | 59*14'16" |
| 1003 | . 038 | . 018 | $37^{\circ} 53^{\prime \prime} 16^{\prime \prime}$ | 1006 | 1007 | . 017 | . 012 | $11^{\circ} 45^{\prime \prime} 51^{\prime \prime}$ |
| 1004 | . 043 | . 020 | 14* $48^{\prime \prime}{ }^{\prime \prime}$ | 1002 | 1007 | . 015 | . 009 | -65*34'59" |
| 1005 | . 025 | . 013 | 76³7'28' | 1002 | 1003 | . 024 | . 013 | -84* $45^{\prime} 14^{\prime \prime}$ |
| 1006 | . 014 | . 009 | $50^{\circ} 14^{\prime \prime} 28^{\prime \prime}$ | 1002 | 1004 | . 018 | . 013 | $42^{\circ} 27^{\prime \prime} 4{ }^{\prime \prime}$ |
| 1007 | . 028 | . 016 | $29^{\circ} 25^{\prime} 52^{\prime \prime}$ | 1003 | 1004 | . 019 | . 010 | $-39^{\circ} 02^{\prime} 59^{\prime \prime}$ |
|  |  |  |  | 1003 | 1007 | . 013 | . 012 | -21* $58^{\prime} 22^{\prime \prime}$ |
|  |  |  |  | 3 | 1004 | . 015 | . 013 | $58^{\circ} 21^{\prime \prime} 17$ |
|  |  |  |  | 3 | 1003 | . 020 | . 011 | $1^{\circ} 26^{\prime} 04^{\prime \prime}$ |
|  |  |  |  | 3 | 4 | . 015 | . 012 | -76 ${ }^{\circ} 04^{\prime} 26^{\prime \prime}$ |
|  |  |  |  | 4 | 1003 | . 024 | . 014 | $38^{\circ} 01^{\prime} 35^{\prime \prime}$ |

Table 6.5. 95\% Error Ellipses for Network

The residuals $V$ computed using equation (AII-17) are

$$
\begin{aligned}
\underset{(1,57)}{\mathrm{v}^{\mathrm{T}}=} & (-3.89,3.89,-0.005,0.007,0.010 ; 0.005,-0.004,0.000,0.003, \\
& 0.001,-0.010,-0.008,-0.006,0.008,-0.005,-0.002,0.005, \\
& -0.002,0.002,0.13,-0.72,-1.31,1.90,-0.01,1.34,-0.55,-0.78, \\
& 0.45,-1.29,-0.51,1.36,1.02,-0.33,-0.69,-0.37,0.45,1.68, \\
& -1.76,-1.20,0.84,0.05,0.32,-0.08,1.17,0.31,-0.35,-1.05, \\
& -1.98,1.72,0.26,-0.88,0.88,-0.16,0.03,0.13,-0.73,0.73),(6-5)
\end{aligned}
$$

where the units are arcseconds for the first two residuals, metres for the next 17, and arcseconds for the last 38 residuals. The a posteriori variance factor $\hat{\sigma}_{o}^{2}$ is computed using the above residuals, the $P$ matrix given above, and equation (AII-18) as

$$
\begin{equation*}
\hat{\sigma}_{o}^{2}=\frac{\mathrm{v}^{T} \mathrm{PV}}{\mathrm{df}}=0.58488 \tag{6-6}
\end{equation*}
$$

The residuals and a posterior variance factor are considered again in Chapter 9 for postanalysis of the network.


Figure 6.5 Plot of 95\% Error Ellipses for the Network

## 7. A PRIORI KNOWLEDGE OF PARAMETERS

This chapter considers points which have some independent a priori estimate of their position. The coordinates of these points are treated as observables, and have the following simple observation equation:

$$
\begin{equation*}
L_{x}=x \tag{7-1}
\end{equation*}
$$

with associated variance covariance matrix $C_{L_{X}}$, the accuracy estimate of these so-called weighted parameters. By expanding the matrices of Appendix II to include these new observables, the least squares estimate of the solution vector (cf. eq. (AII-ll)) becomes

$$
\left.x=-\left[\begin{array}{ll}
{\left[A^{T}\right.} & I
\end{array}\right]\left[\begin{array}{ll}
C_{L} & 0  \tag{7-2}\\
0 & C_{L_{x}}
\end{array}\right]^{-1}\left[\begin{array}{l}
A \\
I
\end{array}\right]\right]^{-1} \quad\left[\begin{array}{ll}
A^{T} & I]
\end{array} \begin{array}{ll}
C_{L} & 0 \\
0 & C_{L_{x}}
\end{array}\right]^{-1}\left[\begin{array}{l}
W \\
W_{L_{X}}
\end{array}\right]
$$

or, multiplying the matrices together,

$$
\begin{equation*}
X=-\left[A^{T} P A+P_{x}\right]^{-1}\left[A^{T} P W+P_{x} W_{L_{x}}\right] \tag{7-3}
\end{equation*}
$$

where $P_{x}=C_{L_{x}}^{-1}$,
and $W_{L_{x}}=X^{0}-L_{x}$.
Note that $W_{L_{x}}=0$ for the first iteration if $x^{0}$ is taken equal to $L_{x}$. This is not the case for the second and subsequent iterations.

Although not done here, it can be shown (e.g. Krakiwsky [1975]), that the corresponding variance covariance matrix of the parameters $X$ (cf. eq. (AII-16)) is

$$
\begin{equation*}
C_{x}=\left[A^{T} P A+P_{x}\right]^{-1} \tag{7-4}
\end{equation*}
$$

Thus, the only difference between a priori knowledge and no a priori
knowledge of the parameters for the accuracy estimate of the parameters is the addition of the $\mathrm{P}_{\mathrm{x}}$ matrix to the normal equations.

Since the weighted parameters are treated as observables, the degrees of freedom for the adjustment change from df $=n-u$ to

$$
\begin{equation*}
\mathrm{d} f=n-u+u_{x} \tag{7-5}
\end{equation*}
$$

where $u_{x}=$ number of weighted parameters.
Another consequence of this is that the a posteriori variance factor (cf. eq. (AII-18)) is now computed as

$$
\begin{equation*}
\hat{\sigma}_{0}^{2}=\frac{V^{T} P V+V_{L_{x}} P_{x} V_{L_{x}}}{d f} \tag{7-6}
\end{equation*}
$$

where $V_{L_{X}}=\hat{A X}+W_{L_{X}}$ (cf. eq. (AII-I7)) are the residual corrections to the weighted coordinates. Noting that the $A$ matrix for $L_{x}$ is equal to $I$, then $V_{L_{x}}$ is simply

$$
\begin{equation*}
V_{L_{x}}=\hat{X}+W_{I_{x}}=\sum_{i=1}^{m} \hat{X}_{i} \tag{7-7}
\end{equation*}
$$

where $\mathrm{m}=$ number of iterations in the adjustment.
The example shown in Figure 7.1 is exactly the same as that of section 5.1 except that station 1 is now weighted with an a priori variance covariance matrix of

$$
C_{L_{x_{1}}}=\left[\begin{array}{cc}
0.4455 .10^{-1} & -0.709 .10^{-3}  \tag{7-8}\\
-0.709 .10^{-3} & 0.9535 .10^{-1}
\end{array}\right]
$$

in units of $\mathrm{m}^{2}$.


This corresponds to a standard error ellipse $a=0.309 \mathrm{~m}, \mathrm{~b}=0.211 \mathrm{~m}$ and $\theta=-0^{\circ} 47^{\prime} 58^{\prime \prime}$. The initial data (i.e. approximate point coordinates, observations and their standard deviations) are given in Table 5.1. The observation equation is

$$
\begin{equation*}
\underset{(2,1)}{V}=\underset{(2,1)}{W}+\underset{(2,4)}{A} \underset{(4,1)}{\hat{X}} \tag{7-9}
\end{equation*}
$$

or, explicitly
where the units are

$$
\underset{(2,1)}{\mathrm{V}}=\left[\begin{array}{l}
\mathrm{n}  \tag{7-11}\\
\mathrm{~m}
\end{array}\right]+\left[\begin{array}{cccc}
(") \cdot \mathrm{m}^{-1} & (") \cdot m^{-1} & (") \cdot m^{-1} & (") \cdot m^{-1} \\
- & - & - & -
\end{array}\right]\left[\begin{array}{l}
m \\
m \\
m \\
m
\end{array}\right] .
$$

Evaluating $A$ and $W$ using the coordinates and observations of Table 5.1 yields

$$
A=\left[\begin{array}{cccc}
-59.17941 & 57.65335 & 59.17941 & -57.65335 \\
-0.6978111 & -0.7162819 & 0.6978111 & 0.7162819
\end{array}\right], \quad W=\left[\begin{array}{c}
-23.0324 \mathrm{~m} \\
0.11655
\end{array}\right] .
$$

The weight matrix $P$ of the observations is identical to that of section 5.1. Employing formula (7-3) to compute the solution vector $\hat{X}$ gives $\hat{X}$ for the first iteration as

$$
\hat{X}=\left[\begin{array}{c}
0.00000 \mathrm{~m} \\
0.00000 \mathrm{~m} \\
0.11835 \mathrm{~m} \\
-0.27802 \mathrm{~m}
\end{array}\right]
$$

which yield parameters X of
$X=X^{\circ}+\hat{X}=\left[\begin{array}{l}377164.887 \\ 862395.774 \\ 378907.0 \\ 864184.0\end{array}\right]+\left[\begin{array}{c}0.00000 \\ 0.00000 \\ 0.11835 \\ -0.27802\end{array}\right]=\left[\begin{array}{c}377164.877 \mathrm{~m} \\ 862395.774 \mathrm{~m} \\ 378907.118 \mathrm{~m} \\ 864183.722 \mathrm{~m}\end{array}\right]$.

Using these parameter values as new approximate coordinates, and recomputing $A$ and $W$ gives the second iteration solution vector as

$$
\hat{x}=\left[\begin{array}{c}
0.00000 \\
0.00000 \\
-0.00002 \\
0.00000
\end{array}\right]
$$

and the solution has converged. Thus, the final least squares estimate of the parameters is given by equation (7-12), which is identical to the solution obtained in section 5.1 .

The variance covariance matrix of the parameters computed according to equation $(7-4)$ is

$$
C_{x}=\left[\begin{array}{llll}
0.4455 .10^{-1} & -0.709 .10^{-3} & 0.4455 .10^{-1} & -0.709 .10^{-3} \\
& 0.9535 .10^{-1} & -0.709 .10^{-3} & 0.9535 .10^{-1} \\
& & 0.46855185 .10^{-1} & -0.21014856 .10^{-2} \\
\text { Symmetric } & & & 0.97583041 .10^{-1}
\end{array}\right],(7-13)
$$

which gives the following standard error ellipses:

| Point \#1: | $a=0.309 \mathrm{~m}$ | $\mathrm{~b}=0.211 \mathrm{~m}$ | $\theta=-0^{\circ} 47^{\prime} 58^{\prime \prime}$ |
| :--- | :--- | :--- | :--- |
| Point \#2: | $a=0.313 \mathrm{~m}$ | $\mathrm{~b}=0.216 \mathrm{~m}$ | $\theta=-2^{\circ} 22^{\prime} 05^{\prime \prime}$ |
| Relative 1-2: | $a=0.061 \mathrm{~m}$ | $\mathrm{~b}=0.030 \mathrm{~m}$ | $\theta=-45^{\circ} 44^{\prime} 31^{\prime \prime}$. |

From these error ellipses, it is seen that the relative error ellipses give the precision of the actual surveying being done, whereas the station ellipses reflect the fact that point 2 cannot be established more accurately than the accuracy of the starting point \#l. Increasing the confidence level to 95\% (c factor $=2.45$ ) gives

$$
\begin{aligned}
\text { Point \# 1: } & \mathrm{a}=0.756 \mathrm{~m} \quad \mathrm{~b}=0.517 \mathrm{~m} \\
\text { Point \# 2: } & \mathrm{a}=0.766 \mathrm{~m} \quad \mathrm{~b}=0.530 \mathrm{~m} \\
\text { Relative 1-2: } & \mathrm{a}=0.148 \mathrm{~m} \quad \mathrm{~b}=0.072 \mathrm{~m} .
\end{aligned}
$$

These error ellipses are plotted in Figure 7.2.

The residuals $V$ are still zero because this is a unique case. These results compare identically with the example of section 4.8 .1 in Thomson et al. [1978] which uses the direct formulae, and propagation of errors to arrive at the result. Thus, the equivalence of the least squares method and the direct approach for the unique case is seen.


Figure 7.2 Plot of 95\% Error Ellipses for Direct Case with :leichさec: Parameters

The following example depicted in Figure 7.3 is almost the same as the closed traverse in section 6.1. The only difference is that the points which were considered fixed (i.e pts. 1, 2, 3 and 4) in section 6.1 are now weighted. The initial data in Table 6.1 is the same for this case with weighted parameters. Additional initial data includes the weight matrix $P_{x}$ (cf. eq. (7-3)) which is (for points $1,2,3$ and 4, respectively)


Note that points 1 and 2 are considered uncorrelated to points 3 and 4. Figure 7.3 shows the $95 \%$ error ellipses represented by $P_{x}$ above for the four weighted points.

Strictly speaking, the general matrix form of the observation equations for this example is

$$
\underset{(25,1)}{\mathrm{V}}=\underset{(25,22)}{\mathrm{A}} \quad \underset{(22,1)}{\hat{X}}+\underset{(25,1)}{\mathrm{W}}
$$

where $W$ includes $W_{L_{X}}$. Realizing that the rows of $A$ corresponding to the $L_{x}$ observations reduce to the unity matrix (cf. eq. (7-2)), and that $W_{L_{x}}$ reduces to zero for the first iteration (i.e. $L_{x}=X^{\circ}$ ), then the observation equations are written as

$$
\begin{equation*}
\underset{(17,1)}{V}=\underset{(17,22)}{A} \quad \hat{(22,1)} \hat{X}+\underset{(17,1)}{W} \tag{7-14}
\end{equation*}
$$



Figure 7.3 Initial Configuration for Closed Traverse with Weighted Points

The A matrix of section 6.1 is part of the A matrix for this example. Here, however, there are eight more columns for stations 1, 2, 3 and 4. The A matrix for this example is
where the numbers have been rounded to five significant figures and only the first eight columns are given (the final 14 are identical to A of section 6.1). The $W$ matrix in equation (7-14) as well as the $P$ matrix are both identical to those in section 6.1.

Noting that $W_{L_{X}}$ is zero for this first iteration, equation (7-3) is employed to compute the first solution vector as

$$
\begin{aligned}
\hat{\mathrm{x}}^{\mathrm{T}}= & (0.00271,-0.00239,-0.00436,0.00215,0.00170,0.00018,-0.00007, \\
& 0.00002,-0.40785,1.42952,-0.01027,2.43608,-0.32709,0.57935, \\
& -0.49390,1.22106),
\end{aligned}
$$

for stations $1,2,3,4,1003,1004,1006,1007$, respectively. Thus, the updated parameter vector is

$$
\begin{aligned}
\mathrm{x}^{\mathrm{T}}=\mathrm{x}^{\circ \mathrm{T}}+\hat{\mathrm{x}}^{\mathrm{T}}= & (2640.003,1159.998,2529.996,935.002,3660.002,630.000, \\
& 3635.000,355.000,3264.592,646.430,3569.990,917.436, \\
& 2819.673,945.579,3159.506,866.221) .
\end{aligned}
$$

Using this parameter vector as new approximate coordinates $X^{\circ}$, the $A$ and $W$ matrices are reevaluated (note that $W_{L_{x}}$ is no longer zero) and equation (7-3) is again employed to give a second iteration solution vector of

$$
\begin{aligned}
\hat{\mathrm{x}}^{\mathrm{T}}= & (0.00003,0.00014,0.00048,-0.00023,-0.00069,0.00014,0.00019, \\
& -0.00004,0.00441,-0.00096,0.00361,0.00156,0.00150,-0.00063, \\
& 0.00322,-0.00030),
\end{aligned}
$$

which yields the second iteration parameter vector as

$$
\begin{aligned}
\mathrm{X}^{\mathrm{T}=}=\mathrm{X}^{\circ \mathrm{T}}+\hat{\mathrm{X}}^{\mathrm{T}}= & (2640.003,1159.998,2529.996,935.002,3660.001,630.000 \\
& 3635.000,355.000,3264.597,646.429,3569.993,917.438 \\
& 2819.674,945.579,3159.509,866.221) .
\end{aligned}
$$

Evaluating $A, W$ and $W_{L_{X}}$ a third time to compute the third iteration solution vector $\hat{X}$ yields a zero solution vector. The final parameter vector of adjusted coordinates is thus given by equation (7-16).

The variance covariance matrix $C_{x}$ of the parameters is computed using equation (7-¢) to yielc.

where $C_{x}$ has been printed in rows of six columns at a time. The equations of Appendix III ( $C$ factor $=2.45$ ) are used to compute the $95 \%$ station and relative error ellipses listed in Table 7.1 and plotted in Figure 7.4.

Station Ellipses

| Station | $\mathrm{a}(\mathrm{m})$ | $\mathrm{b}(\mathrm{m})$ | $\theta$ | Station | Station | $\mathrm{a}(\mathrm{m})$ | $\mathrm{b}(\mathrm{m})$ | $\theta$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | .011 | .010 | $50^{\circ} 14^{\prime} 35^{\prime \prime}$ | 1 | 2 | .016 | .011 | $32^{\circ} 18^{\prime} 42^{\prime \prime}$ |
| 2 | .012 | .010 | $26^{\circ} 36^{\prime} 31^{\prime \prime}$ | 1 | 1006 | .022 | .011 | $-44^{\circ} 01^{\prime} 33^{\prime \prime}$ |
| 3 | .012 | .010 | $-10^{\circ} 26^{\prime} 14^{\prime \prime}$ | 1006 | 1007 | .021 | .012 | $-65^{\circ} 23^{\prime} 33^{\prime \prime}$ |
| 4 | .012 | .010 | $-5^{\circ} 13^{\prime} 45^{\prime \prime}$ | 1003 | 1007 | .022 | .009 | $-31^{\circ} 42^{\prime} 07^{\prime \prime}$ |
| 1003 | .024 | .021 | $9^{\circ} 48^{\prime} 58^{\prime \prime}$ | 1003 | 1004 | .024 | .014 | $73^{\circ} 00^{\prime} 12^{\prime \prime}$ |
| 1004 | .025 | .017 | $-31^{\circ} 49^{\prime} 52^{\prime \prime}$ | 3 | 1004 | .023 | .012 | $-24^{\circ} 38^{\prime} 31^{\prime \prime}$ |
| 1006 | .022 | .013 | $-36^{\circ} 09^{\prime} 57^{\prime \prime}$ | 3 | 4 | .017 | .011 | $-3^{\circ} 28^{\prime} 13^{\prime \prime}$ |
| 1007 | .025 | .019 | $-26^{\circ} 37^{\prime} 59^{\prime \prime}$ |  |  |  |  |  |

Table 7.1 95\% Ercor Ellipses for Closed Traverse with Weighted Parameters


Figure 7.4 Plot of 95\% Error Ellipses for Closed Traverse with Weighted Points
the residuals $V$ are computed using equation (AII-17) and the solution vector of equation (7-15) to be

$$
\begin{aligned}
\mathrm{v}^{\mathrm{T}}= & (0.84,-0.84,0.58,-0.58,0.40,-0.40,0.16,-0.16,0.31,-0.31, \\
& 0.02,-0.02,0.000,-0.002,-0.004,-0.000,-0.005)
\end{aligned}
$$

where the units of the first 12 elements are arcseconds and the final five elements are in metres. Using these residuals and the summation of $\hat{X}$ 's for all of the iterations (cf: eq. (7-7)) the a posteriori variance factor $\hat{\sigma}_{0}^{2}$ is computed via equation (7-6) as

$$
\hat{\sigma}_{0}^{2}=\frac{v^{T} P V+V_{L_{X}}^{T} P_{x} V_{L_{X}}}{d f}=0.79712
$$

The results (i.e. adjusted coordinates and accuracy estimates)
for this example and that of section 6.1 where points $1,2,3$ and 4 were considered fixed are significantly different. With weighted points, the observations are allowed to affect the final coordinates of the weighted points to a degree dictated by the observation as well as coordinate weight matrices ( P and $\mathrm{P}_{\mathrm{x}}$ ). This gives a more realistic least squares solution than does the fixed point approach.

## 8. PREANALYSIS

Preanalysis is the study of the design of a network. The design is carried out prior to the establishment of the network in the field, and thus no observations are necessary for a preanalysis. By optimizing the accuracy and distribution of the observables before entering the field, the required accuracy of the network points is achieved most expediently.

Preanalysis is based on equation (AII-16), i.e.

$$
C_{x}=\left[A^{T} C_{L}^{-1} A\right]^{-1}
$$

Since the variance covariance matrix of the parameters $C_{x}$ does not require knowledge of the actual observations (the only place where the observations are necessary is for computation of $W$ ), it can be computed knowing the approximate coordinates of the unknown points along with some proposed observations (and their standard deviations) amongst them. If some parameters are weighted, then equation (7-4) applies, i.e.

$$
C_{x}=\left[A^{T} P A+P_{x}\right]^{-1}
$$

Computing the station and relative error ellipses from $C_{x}$, the results of a network design are readily apparent. The design can be altered by proposing different observations and standard deviations and/or changing the position or number of unknown points, and recomputing $C_{x}$.

Better use of the already existing design is made when the sequential design approach [e.g. Nickerson et al., 1978] is used. This method is characterized by the following equation:

$$
\begin{equation*}
C_{x_{i}}=C_{x_{i-1}}-C_{x_{i-1}} A_{i}^{T}\left( \pm C_{L_{i}}+A_{i} C_{x_{i-1}} A_{i}^{T}\right)^{-1} A_{i} C_{x_{i-1}} \tag{8-1}
\end{equation*}
$$

where $C_{x_{i}}=$ covariance matrix of the parameters utilizing all observables, $C_{x_{i-1}}=$ previous covariance matrix which is being altered, $A_{i}=$ design matrix for the observables to be added or deleted, $C_{L_{i}}=$ variance covariance matrix of the observables being added or deleted.

The plus and minus signs preceding $C_{L_{i}}$ refer to addition and deletion of observables, respectively. The size of inverse to be computed (usually the most time consuming task) is equal to the number of observables being added or deleted, not the number of parameters as in the nonsequential equations. The standard deviations of specific observables are changed by subtracting the old observable with its standard deviation, and adding it back with the new standard deviation. The following examples illustrate the preanalysis process.

### 8.1 Traverse Design

Figure 8.1 depicts the initial information (see Table 8.1) for the traverse design. There are four fixed points and three unknown

## UNKNOWN POINTS

| Station | $x(m)$ | $y(m)$ | Des.Acc. $(m)$ | Station | $x(m)$ | $y(m)$ |
| :---: | ---: | ---: | :--- | :--- | :--- | :--- |
| 1 | 293682 | 225293 | 0.05 | 1102 | 293054.171 | 225214.674 |
| 2 | 293976 | 225607 | 0.05 | 1116 | 293571.011 | 225598.373 |
| 3 | 294421 | 225284 | 0.05 | 1105 | 295267.293 | 225419.706 |
|  |  |  |  | 1106 | 295004.038 | 225951.144 |

Table 8.1 Initial Data for Traverse Design


Figure 8.1 Initial Data Plot

| OBSERVABLE TYPE | STANDARD DEVİATION | FROM STATION | TO STATION |
| :--- | :--- | :--- | :---: |
| direction | $3: 0$ | 1116 | 1102 |
| airection | $3: 0$ | 1116 | 1 |
| direction | $2!0$ | 1 | 1116 |
| direction | 2.0 | 1 | 2 |
| direction | 2.0 | 2 | 1 |
| direction | 2.0 | 2 | 3 |
| direction | 0.02 m | 1116 | 1 |
| direction | 0.02 m | 2 | 1 |
| direction | 0.02 m | 2 | 3 |

Table 8.2 Initial Observables
points. The required accuracy is represented by the design circles of 5 cm . radius around the unknown points.

The proposed initial observables between these points are listed in Table 8.2 and plotted in Figure 8.2. The design is first treated as an open ended traverse similar to the example in section 5.5. The main difference is that directions are used here instead of angles. Thus, the A matrix is of size $9 \times 9$ (i.e. $n=9, u=9$ ( 6 unknown coordinates, 3 orientation unknowns)), and is given as


Figure 8.2 Initial observables Plot
$\underset{(9,9)}{A}=\left[\begin{array}{cccccccc}0.0 & 0.0 & 0.0 & 0.0 & 0.0 & 0.0 & -1.0 & 0.0 \\ -596.637 & -216.850 & 0.0 & 0.0 & 0.0 & 0.0 & -1.0 & 0.0 \\ -596.637 & -216.850 & 0.0 & 0.0 & 0.0 & 0.0 & 0.0 & -1.0 \\ -350.0322 & 327.7371 & 350.0322 & -327.7371 & 0.0 & 0.0 \\ -350.0322 & 327.7371 & 350.0322 & -327.7371 & 0.0 & 0.0 & 0.0 & 0.0 \\ 0.0 & 0.0 & 220.3494 & 303.5774 & -220.3494 & -303.5774 & 0.0 & 0.0 \\ 0.3415916 & -0.9398485 & 0.0 & 0.0 & 0.0 & 0.0 & 0.0 & 0.0 \\ -0.683477 & -0.729972 & 0.683477 & 0.729972 & 0.0 & 0.0 & 0.0 & 0.0 \\ 0.0 & 0.0 & -0.8092862 & 0.5874145 & 0.8092862-0.5874145 & 0.0 & 0.0 & 0.0 \\ 0.0\end{array}\right]$.

This results from a combination of the direction mathematical model developed in section 3.2 (for the first six rows of $A$ ) and the distance mathematical model of chapter 4 (last three rows of $A$ above). The last three columns of $A$ are for the orientation unknowns $Z_{1116}, Z_{1}$, and $Z_{2}$. The $P$ matrix is (assuming $\sigma_{0}^{2}=1$ )


Computing ( $A^{T} P A$ ) and taking its inverse yields


The first six rows and columns are the variance covariance matrix for the coordinates of points 1,2 and 3 . The last three rows and columns represent the variance and covariance of the orientation unknowns, and are of no practical concern. Since the orientation unknowns are nuisance parameters, only the first (6 $x$ ) submatrix of $C_{x}$ will be considered as representative of the traverse being designed. The 99\% error ellipses
assuming $\sigma_{0}^{2}$ known and nonsimultaneous ellipses ( $c$ factor $=3.035$, see Appendix II) are listed in Table 8.3. From Figure 8.3

STATION ELLIPSES (99\%)
RELATIVE ELLIPSES (99\%)

Point Semimajor $(\mathrm{m})$ Semiminor $(\mathrm{m}) \quad \underline{\theta} \quad$ Points Semimajor $(\mathrm{m})$ Semiminor $(\mathrm{m})$.

| 1 | 0.061 | 0.021 | $-19^{\circ} 58^{\prime} 26^{\prime \prime}$ | $1-2$ | 0.061 | 0.033 | $43^{\circ} 06^{\prime} 57^{\prime \prime}$ |
| :--- | :--- | :--- | :---: | :--- | :--- | :--- | :--- |
| 2 | 0.079 | 0.049 | $7^{\circ} 19^{\prime} 43^{\prime \prime}$ | $2-3$ | 0.061 | 0.049 | $-54^{\circ} 01^{\prime} 35^{\prime \prime}$ |
| 3 | 0.103 | 0.073 | $2^{\circ} 09^{\prime} 09^{\prime \prime}$ |  |  |  |  |

Table 8.3 Confidence Ellipses from Initial Observables
and the above table it is seen that all of the station ellipses lie outside the required accuracy circle of 5 cm radius. New confidence ellipses


Figure 8.3 Plot of $99 \%$ Confidence Ellipses from Initial Observables.
are computed using equation (8-1) to update the design. Table 8.4 shows the update observables which are depicted in Figure 8.4. Referring to
OBSERVATION TYPE STANDARD DEVIATION FROM STATION TO STATION

| Direction | $1: 5$ | 3 | 2 |
| :--- | :---: | :---: | :---: |
| Direction | $1: 5$ | 3 | 1105 |
| Direction | $1: 5$ | 1105 | 3 |
| Direction | $1: 5$ | 1105 | 1106 |
| Distance | 0.02 m | 3 | 2 |

Table 8.4 Update Observables


Figure 8.4 Plot of Update Observables
equation ( $8-1$ ), these new observables are added by first computing $A_{i}$ and $C_{L_{i}}$, and in turn $C_{X_{i}}$ taking equation (8-2) as $C_{x_{i-1}}$. After performing these operations, the new $C_{X_{i}}$ corresponding to the 3 unknown points is (to 5 significant digits)


The $99 \%$ error elinpses (again assuming $\sigma_{0}^{2}$ known and nonsimultaneous ellipses for a $c$ factor of 3.035 ) computed from the above $C_{x}$ matrix are listed in

| STATION ELLIPSES (99\%) |  |  |  | RELATIVE ELLIPSES (99\%) |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Point | Semimajor (m) | Semiminor (m) | $\underline{\theta}$ | Points | Semimajor(m) | Semiminor (m) | ) $\underline{\theta}$ |
| 1 | 0.047 | 0.014 | $-24^{\circ} 18^{\prime} 25^{\prime \prime}$ | 1-2 | 0.046 | 0.018 | $51^{\circ} 07^{\prime} 30^{\prime}$ |
| 2 | 0.041 | 0.040 | 50oㄴ '24" | 2-3 | 0.047 | 0.019 - | $-59^{\circ} 19^{\prime} 18^{\prime}$ |
| 3 | 0.045 | 0.022 | $85^{\circ} 01^{\prime} 08^{\prime \prime}$ |  |  |  |  |

Table 8.5 99\% Error Ellipses After Update

Table 8.5 and plotted in Figure 8.5. Obviously, all of the error ellipses


Figure 8.5 Plot of Updated 99\% Error Ellipses
now meet the required accuracy, and the design is finished. A summary of the pronosed observations is given in Table 8.6.

| FROM | TO | TYPE | $\sigma$ |
| ---: | ---: | ---: | :---: |
|  |  |  | $\sigma$ |
| 1116 | 1102 | 1 | 3.0000 |
| 1116 | 1 | 1 | 3.0000 |
| 1116 | 1 | 2 | .0200 |
| 1105 | 3 | 1 | 1.5000 |
| 1105 | 1106 | 1 | 1.5000 |
| 1105 | 3 | 2 | .0200 |
| 1 | 1106 | 1 | 2.0000 |
| 1 | 2 | 1 | 2.0000 |
| 2 | 1 | 1 | 2.0000 |
| 2 | 3 | 1 | 2.0000 |
| 2 | 1 | 2 | .0200 |
| 2 | 3 | 2 | .0200 |
| 3 | 2 | 1 | 1.5000 |
| 3 | 1105 | 1 | 1.5000 |

Note: Type $1=$ Direction
Type 2 = Distance

Table 8.6 Observable Summary

### 8.2 Property Survey Design

The initial data of this design (see Table 8.7 and Figure 8.6)
is characteristic of a simple lot layout often encountered in practice.

| UNKNOWN POINTS |  |  | KNOWN POINTS |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Station | $x(m)$ | $y(m)$ | Des.Acc. | Station | $x(m)$ | $y(m)$ |
| 1 | 155721.0 | 119687.0 | 0.05 | 1004 | 155221.688 | 119515.558 |
| 2 | 156019.0 | 119595.0 | 0.05 | 1005 | 155493.110 | 119604.892 |
| 3 | 156027.0 | 119386.0 | 0.05 |  |  |  |
| 4 | 156204.0 | 119596.0 | 0.05 |  |  |  |
| 5 | 156213.0 | 119388.0 | 0.05 |  |  |  |

Table 8.7 Initial Data for Property Survey Design

Fixed points 1004 and 1005 represent two second order monuments, point
1 an intermediate point, and points 2, 3, 4 and 5 the four lot corners to be established. Again, the required accuracy circle is of 5 cm . radius, but this time at a confidence level of $95 \%$.


Figure 8.6 Initial Data Plot for Property Survey Design

The initial observables considered for the design are listed in Table 8.8 and plotted in Figure 8.7. These initial accuracies JBSERVABLE TYPE STANDARD DEVIATION FROM STATION TO STATION

| Direction | 2.0 | 1005 | 1004 |
| :--- | :--- | :---: | :---: |
| Direction | 2.0 | 1005 | 1 |
| Direction | 2.0 | 1 | 1005 |
| Direction | 2.0 | 1 | 2 |
| Distance | 0.02 m | 1 | 1005 |
| Distance | 0.02 m | 1 | 2 |
| Direction | 3.0 | 2 | 1 |
| Direction | 3.0 | 2 | 4 |
| Direction | 3.0 | 2 | 3 |
| Distance | 0.02 m | 2 | 4 |
| Direction | 3.0 | 4 | 2 |
| Direction | 3.0 | 4 | 5 |
| Direction | 3.0 | 5 | 4 |
| Direction | 3.0 | 5 | 3 |
| Distance | 0.02 m | 5 | 4 |
| Distance | 0.02 m | 5 | 3 |

Table 8.8. Initial Observables for Property Survey Design
could be achieved using 4 sets of direction observables at station 1005 and station 1 , and 2 sets at stations 2,4 and 5 with a 1 " theodolite (see Table 2.4). The distance's accuracy of 0.02 m is : easily achièved using either lightwave or microwave EDM with normal meteorological readings (see Table 2.8). The A matrix for the initial $C_{x}$ is (16 $x$ 15) since there are 16 observables


Figure 8.7 plot of Initial Observables for Property Survey Design
(see Table 8.8) and 15 unknowns (10 coordinates and 5 orientation unknowns). The A matrix is formed using the equations developed in section 3.2 and chapter 4. After computing $A$ and $P$, equation (AII-16) is used to compute $\mathrm{C}_{\mathrm{X}}$ as


Figure 8.8 depicts the $95 \%$ (assuming $\sigma_{0}^{2}$ known, nonsimultaneous ellipses, $c$ factor $=2.45$ ) error ellipses resulting from this initial design.


Figure 8.8 Initial Design Results
point $l$ is the only point which meets the required accuracy. From the shape of the relative error ellipses (i.e. long and skinny along the line of sight), it is obvious that the distances are less accurate relative to the direction observables. Thus, the design is altered by
changing the standard deviation of most of the distance observables from 0.02 m to 0.01 m . This would correspond to using a more accurate instrument for observing distances (e.g. lightwave EDM with standard meteorological readings (cf. table 2.8)). These changes are effected by first subtracting (use - sign in eq. (8-1)) the existing distances with 0.02 m for a standard deviation, and then adding them back again (+ sign in eq. (8-l)) with a standard deviation of 0.01 m . After performing this operation for each distance observable except that from point 5 to point 3 , the $C_{x}$ matrix is


As can be seen in Figure 8.9, all of the station ellipses now fall within the required accuracy circle, and the design is finished. A list
of the final 95\% station and relative error ellipse (using the same assumptions as for the initial design) is made in Table 8.9.


Figure 8.9 Final Plot of Property Survey Design

STATION ELLIPSES (95\%)

| Point Semimajor $(\mathrm{m})$ | Semiminor $(\mathrm{m})$ | $\theta$ | Points | Semimajor $(\mathrm{m})$ | Semiminor $(\mathrm{m})$ | $\theta$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 0.025 | 0.008 | $70^{\circ} 11^{\prime} 09^{\prime \prime}$ | $1-2$ | 0.025 | 0.015 | $-72^{\circ} 50^{\prime} 36^{\prime \prime}$ |
| 2 | 0.033 | 0.023 | $87^{\circ} 54^{\prime} 11^{\prime \prime}$ | $2-3$ | 0.028 | 0.014 | $-2^{\circ} 15^{\prime} 34^{\prime \prime}$ |
| 3 | 0.040 | 0.034 | $78^{\circ} 50^{\prime} 42^{\prime \prime}$ | $2-4$ | 0.022 | 0.013 | $-89^{\circ} 04^{\prime} 23^{\prime \prime}$ |
| 4 | 0.040 | 0.032 | $88^{\circ} 14^{\prime} 20^{\prime \prime}$ | $3-4$ | 0.025 | 0.017 | $-12^{\circ} 02^{\prime} 02^{\prime \prime}$ |
| 5 | 0.047 | 0.037 | $78^{\circ} 24^{\prime} 09^{\prime \prime}$ | $4-5$ | 0.025 | 0.017 | $-2^{\circ} 28^{\prime} 02^{\prime \prime}$ |

Table 8.9 Final 95\% Error Ellipses for Property Survey Design
This chapter has shown how the preanalysis or design of a horizontal network can be done. This technique is used wherever a priori knowledge of the expected accuracies of the points in a network is desired. The next chapter treats the equally important matter of testing the results of an adjustment of a network to see if they are reasonable. This process is known as postanalysis.

## 9. POSTANALYSIS

Postanalysis of a horizontal network tests whether the results of an adjustment are reliable, and is based on multivariate analysis. One important test which is performed is the chi-square test of the variance factor. This test takes the form

$$
\begin{equation*}
\frac{d f \hat{\sigma}_{o}^{2}}{x_{d f, 1-\frac{\alpha}{2}}^{2}}<\sigma_{o}^{2}<\frac{d f \hat{\sigma}_{o}^{2}}{x_{\overline{d f}, \frac{\alpha}{2}}} \tag{9-1}
\end{equation*}
$$

where $d f=$ degrees of freedom of the adjustment,
$\hat{\sigma}_{0}^{2}=$ a posteriori variance factor computed by equation (AII-18),
$\alpha=$ significance level (e.g. 0.05),
$\sigma_{0}^{2}=$ a priori variance factor (usually assumed to be 1 ), $\chi^{2}=$ chi-square distribution value from Table AIII-1 (replace u by df).

Equation (9-1) tests the null hypothesis $H_{o}$
$H_{0}: \sigma_{0}^{2}=\sigma_{0}^{2}$ hypothesized,
i.e. is the actual value of $\sigma_{0}^{2}$ equal to what it was assumed to be (e.g. 1)? If the test fails, then this hypothesis is rejected at the (l-ג) \% confidence level. Two possible reasons for its failure are

1) Incorrect a priori covariance matrix $C_{L}$ of the observations (i.e. wrong weights for the observations),
2). Observations are not normally distributed.

The first reason given is usually the first to be investigated. If it is found that the weights are chosen correctly and the test still fails, then the second reason is examined. The observations are examined by testing the residuals for outliers similar to the data screening process explained
in section 2.3. Each residual $\mathrm{v}_{\mathrm{i}}$ from the adjustment is tested as follows:

$$
\begin{equation*}
n_{\frac{\alpha}{2}}{ }^{\sigma} v_{i}<v_{i}<n_{1-\frac{\alpha}{2}} \sigma_{v_{i}} \tag{9-2}
\end{equation*}
$$

where $n_{\frac{\alpha}{2}}=$ values of the normal distribution for probability $\frac{\alpha}{2}$
(see Table 2.14),
$\alpha=$ significance level (e.g. 0.05),
$\sigma_{v_{i}}=$ known a priori standard deviation of the observation whose residual is being tested,
$v_{i}=$ residual being tested (computed by eq. (AII-17)).
Assuming that the a priori standard deviation $\sigma_{v_{i}}$ is known, then if this test fails the observation does not come from a normal distribution. This usually implies that some systematic bias has affected the observation, and it should be reobserved.

Another test which is useful is one for comparing two determinations of the same set of parameters to see if they are significantly different. This test assumes that the difference vector ( $\mathrm{X}_{2}-\mathrm{X}_{1}$ ) between the two determinations is a random variable which is normally distributed. The test is

$$
\begin{equation*}
\left(x_{2}-x_{1}\right)^{T} c_{x_{2}}^{-1}\left(x_{2}-x_{1}\right)<x_{u, 1-\alpha}^{2} \tag{9-3}
\end{equation*}
$$

where $X_{2}=$ vector of parameters being tested,

$$
\begin{aligned}
x_{1} & =\text { originally determined parameters, } \\
C_{x_{2}}= & \text { variance covariance matrix of the parameters being compared, } \\
x_{u, 1-\alpha}^{2}= & \text { chi-square distribution with } u \text { degrees of freedom at probability } \\
& \text { level } 1-\alpha \text { (see Table AIII.1) , }
\end{aligned}
$$

```
u = number of parameters being compared (i.e. dimension of vector
```

$$
\begin{aligned}
& \left.\left(X_{2}-x_{1}\right)\right), \\
\alpha= & \text { significance level (e.g. 0.05). }
\end{aligned}
$$

If the test fails, then the two parameter determinations are considered different at the (1- $\alpha$ ) $\%$ confidence level. If it passes, then the two sets of parameters cannot be considered significantly different (again at (1-a)\% confidence level). One precautionary note when using this test is that the two sets of parameters should be determined using approximately the same level of accuracy; i.e. $C_{x_{1}}$ and $C_{x_{2}}$ should not be greatly different. If a network is being designed or adjusted specifically for the purpose of comparing to a previous adjustment of the same network, then the simultaneous error ellipses (Appendix III, eqs. (AIII-12) and (AIII-13)) should be computed since all of the points are required to be inside the $(1-\alpha) \%$ confidence ellipse simultaneously.

The following examples illustrate some of the postanalysis concepts described above.

Using the a posteriori variance factor from section 6.1 and equation (9-1), the chi-square test of the variance factor (for $\alpha=0.05$ ) yields

$$
\begin{gathered}
\frac{3 \cdot 1.9214}{x_{3,0.975}^{2}}<1<\frac{3 \cdot 1.9214}{x_{3,0.025}^{2}} \\
\frac{5.76}{9.35}<1<\frac{5.76}{0.216} \\
0.62<1<26.69
\end{gathered}
$$

and the test passes. The hypothesis that the a priori variance factor is 1 cannot be rejected at the $95 \%$ confidence level.

The chi-square test on the variance factor should always be performed after an adjustment as an overall check on the validity of the results. If the test fails, then there is a good chance that something is wrong in the adjustment.

As an example of testing of the residual for outliers, the first residual of equation (6-2) is tested. From equation (9-2), the test is

$$
n_{0.025} 2: 0<2: 07<n_{0.975} 2: 0
$$

where $\alpha$ is assumed to be 0.05 , and the a priori standard deviation for this direction (from 1 to 1006) is 2:0. From Table 2.14, the value of $n_{0.025}=-1.96$, and $n_{0.975}=1.96$. The test becomes

$$
-3.92<2: .07<3.92
$$

which is true, and the test passes. This teat should be performed on all residuals of an adjustment even though the chi-square test on the variance factor passes. It is easily verified that all of the residuals for the examples in sections 6.1 and 6.2 pass the outlier test. It is said, then, that we are $95 \%$ confident that the residuals come from a normal distribution.

## REFERENCES

Bomford, G. (1975). Geodesy. London, Oxford University Press, 3rd ed.

Burnside, C.D. (1971). Electromagnetic Distance Measurement. London, Crosby Lockwood Staples.

Cooper, M.A.R. (1971). Modern Theodolites and Levels. London, Crosby Lockwood \& Son Ltd.

Energy, Mines and Resources (1973). Specifications and Recommendations for Control Surveys and Survey Markers. Surveys and Mapping Branch, Misc. Ser. 73/3.

Faig, W. (1972). Advanced Surveying I. Lecture Notes no. 27, Dept. of Surveying Engineering, University of New Brunswick, Fredericton.

Halmos, F. (1977). High Precision Measurement and Evaluation Method for Azimuth Determination with Gyro-Theodolites. Manuscripta Geodetica, vol. 2, no. 3, p. 213.

Hogg, R.V. and A.T. Craig (1970). Introduction to Mathematical Statistics. New York, Macmillan, 3rd ed.

Krakiwsky, E.J. (1975). A Synthesis of Recent Advances in the Method of Least Squares. Lecture Notes no. 42, Dept. of Surveying Engineering, University of New Brunswick, Fredericton, N.B.

Krakiwsky, E.J. (1978). Statistical Techniques and Doppler Satellite Positioning. Invited paper to the Royal Society Discussion Meeting on Satellite Doppler Tracking and Geodetic Applications, London, England.

Krakiwsky, E.J., D.B. Thomson and R.R. Steeves (1977). A Manual for Geodetic Coordinate Transformations in the Maritimes. Technical Report no. 48, Dept. of Surveying Engineering, University of New Brunswick, Fredericotn, N.B.

Krakiwsky, E.J. and D.B. Thomson (1978). Mathematical Models for Horizontal Geodetic Networks. Lecture Notes no. 48, Dept. of Surveying Engineering, University of New Brunswick, Fredericton, N.B.

Kreyszig, E. (1972). Advanced Engineering Mathematics. New York, John Wiley and Sons.

Mikhail, E.M. (1976). Observations and Least Squares. New York., IEP Series in Civil Engineering.

Mueller, I.I. (1969). Spherical and Practical Astronomy as Applied to Geodesy. New York, Frederick Ungar.

Nickerson, B.G. (1978). A Priori Estimation of Variance for Surveying Observables. Technical Report no. 57, Dept. of Surveying Engineering, University of New Brunswick, Fredericton, N.B.

Nickerson, B.G., E.J. Krakiwsky, D.B. Thomson, M.L. Syverson-Krakiwsky, and J.M. Crawford (1978). Design of Survey Networks Using Interactive Computer Graphics. Proceedings of the American Congress on Surveying and Mapping, Washington.

Pope, A.J. (1976). The Statistics of Residuals and the Detection of Outliers. NOAA Technical Report NOS 65 NGS 1, U.S. Dept. of Commerce, Rockville, U.S.A.

Rainsford, H.F. (1957). Survey Adjustments and Least Squares. London, Constable.

Saastamoinen, J. (ed.) (1967). Surveyor's Guide to Electromagnetic Distance Measurement. Toronto, University of Toronto Press.

Smith, J.R. (1970). Optical Distance Measurement. London, Crosby Lockwood and Son Ltd.

Steeves, R.R. (1978). A Users' Manual for Program GEOPAN Geodetic Plane Adjustment and Analysis. Technical Report no. 54, Dept. of Surveying Engineering, University of New Brunswick, Fredericton, N.B.

Thomson, D.B. (1978). Introduction to Geodetic Astronomy. Lecture Notes no. 49, Dept. of Surveying Engineering, University of New Brunswick, Fredericton, N.B.

Thomson, D.B., E.J. Krakiwsky and J.R. Adams (1978). A Manual for Geodetic Position Computations in the Maritime Provinces. Technical Report no. 52, Dept. of Surveying Engineering, University of New Brunswick, Fredericton, N.B.

Vanicek, P. and E.J. Krakiwsky (in prep.); Concepts of Geodesy. Amsterdam, North-Holland.

Vincenty, T. (1973). Three Dimensional Adjustment of Geodetic Networks. Internal Report, DMAAC Geodetic Survey Squadron.

Wells, D.E. and E.J. Krakiwsky (1971). The Method of Least Squares. Lecture Notes no. 18, Dept. of Surveying Engineering, University of New Brunswick, Fredericton, N.B.

## Taylor's Series

The Taylor series is used to linearize non-linear mathematical models. Given a single function $f(x)$ of a single variable $x$ (unidimensional case), and a known value of this function $f(a)$ at $x=a$, then the function $f(x)$ is given by Taylor's series as

$$
f(x)=f(a)+\left.\frac{\partial f}{\partial x}\right|_{a}(x-a)+\left.\frac{\partial^{2} f}{\partial x^{2}}\right|_{a} \frac{(x-a)^{2}}{2!}+\ldots+\left.\frac{\partial^{n} f}{\partial x^{n}}\right|_{a} \frac{(x-a)^{n}}{n!}+\ldots
$$

(AI-1)

For values of $x$ close to $a$, the linear approximation is used. This is

$$
\begin{equation*}
f(x)=f(a)+\left.\frac{\partial f}{\partial x}\right|_{a}(x-a) \tag{AI-2}
\end{equation*}
$$

From Figure AI.l, the geometric meaning of this linear approximation is clear. The function $f(x)$ is approximated by a straight tangent to $f(x)$ at a


Figure AI.l. Geometric Interpretation of Linear Taylor's Series
with slope $\left.\frac{\partial f}{\partial x}\right|_{a}$. When $(x-a)$ is given (or evaluated), the value of the function $f(x)$ is approximated by $b$, and the exact value is $c$. Thus, the error arising from using the linear approximation is c-b.

If $f$ is a function of more than one variable, say $f\left(x_{1}, x_{2}\right)$ and
its value is known at $x_{1}=a_{1}$ and $x_{2}=a_{2}$, then for values of $\left(a_{1}, a_{2}\right)$ close to $\left(x_{1}, x_{2}\right)$ the linear approximation is

$$
f\left(x_{1}, x_{2}\right)=f\left(a_{1}, a_{2}\right)+\left.\frac{\partial f}{\partial x_{1}}\right|_{a_{1}, a_{2}}\left(x_{1}-a_{1}\right)+\left.\frac{\partial f}{\partial x_{2}}\right|_{a_{1}, a_{2}} \quad\left(x_{2}-a_{2}\right)
$$

(AI-3)

Setting

$$
x=\left[\begin{array}{l}
x_{1} \\
x_{2}
\end{array}\right], \Delta x=\left[\begin{array}{c}
x_{1}-a_{1} \\
x_{2}-a_{2}
\end{array}\right], \quad x^{\circ}=\left[\begin{array}{l}
a_{1} \\
a_{2}
\end{array}\right], \quad \frac{\partial f}{\partial x}=\left[\begin{array}{ll}
\frac{\partial f}{\partial x_{1}} & \frac{\partial f}{\partial x_{2}}
\end{array}\right]
$$

then

$$
\begin{equation*}
f(x)=f\left(x^{\circ}\right)+\left.\frac{\partial f}{\partial x}\right|_{X^{\circ}} \quad \Delta X \tag{AI-4}
\end{equation*}
$$

If there is more than one function of $X\left(e . g . f_{1}, f_{2}\right)$ then the following set of equations exists:

$$
\begin{aligned}
& f_{1}(X)=f_{1}\left(X^{\circ}\right)+\left.\frac{\partial f_{1}}{\partial x}\right|_{x^{\circ}} \Delta x \\
& f_{2}(X)=f_{2}\left(X^{\circ}\right)+\left.\frac{\partial f_{2}}{\partial X}\right|_{X^{\circ}} \Delta x
\end{aligned}
$$

Setting

$$
F=\left[\begin{array}{c}
f_{1} \\
f_{2}
\end{array}\right], \quad \frac{\partial F}{\partial x}=\left[\begin{array}{cc}
\frac{\partial f_{1}}{\partial x_{1}} & \frac{\partial f_{1}}{\partial x_{2}} \\
\frac{\partial f_{2}}{\partial x_{1}} & \frac{\partial f_{2}}{\partial x_{2}}
\end{array}\right]
$$

then

$$
\begin{equation*}
f(X)=F\left(X^{\circ}\right)+\left.\frac{\partial F}{\partial X}\right|_{X^{\circ}} \Delta X \tag{AI-5}
\end{equation*}
$$

Equation (AI-5) is the matrix form of the Taylor's series linear approximation. Equations (AI-2) and (AI-4) can be thought of as special cases of this matrix form.

## Least Squares Method

The least squares method is usually used to give a unique solution for an overdetermined case (i.e. number of observations $n$ greater than number of parameters u). Only the inverse (explicit) form of mathematical model (cf. eq. (l-2)), which is sometimes called the parametric case of adjustment, is considered here. In matrix form, this model is expressed as

$$
\begin{aligned}
& L= F(X), \\
& \text { or } \quad F(X)-L=0, \\
& \text { where } \quad L= \text { vector of observations, } \\
& F(X)= \text { non-linear functions of the } \\
& \text { parameters } X .
\end{aligned}
$$

The linearized form of this inverse (explicit) model is (see also Appendix I and chapters 3 and 4)

$$
\begin{equation*}
\underset{(n, u)}{A} \underset{(u, 1)}{\hat{X}}+\underset{(n, 1)}{.}-\underset{(n, 1)}{V}=0, \tag{AII-2}
\end{equation*}
$$

where $\mathrm{V}=$ vector of residuals or corrections to the observations,
$A=\left.\frac{\partial F}{\partial X}\right|_{X^{\circ}}=$ design matrix or Jacobian of transformation from observation space to parameter space,
$\hat{X}=$ solution vector of corrections, which, when added to the approximate values $\mathrm{X}^{0}$ gives the parameters X (see eq. AII-12),
$W=F\left(X^{\circ}\right)-L=$ misclosure vector.

The least squares estimate for $X$ is obtained subject to the condition

$$
\mathrm{V}^{\mathrm{T}} \mathrm{PV}=\text { minimum }
$$

(AII-3)
where $\underset{(n, n)}{P}=\sigma_{0}^{2} C_{L}^{-1}$ is called the weight matrix of the observables, $\sigma_{0}^{2}=$ a priori variance factor,

$$
C_{L}=\text { variance covariance matrix of the observables }
$$

$$
(n, n)
$$

The variation function $\phi$ relating the unknown quantities $\hat{X}$ and $V$ to the known quantities $A, W$, and $P$ is

$$
\begin{equation*}
\phi=V^{T} P V+2 K^{T}(A \hat{X}+W-V) \tag{AII-4}
\end{equation*}
$$

where $K=$ unknown vector of Lagrange correlates.
To find the minimum of the variation function, the derivatives with respect to $\hat{x}$
to $X$ and $V$ are found and set to zero. Thus

$$
\begin{align*}
& \frac{1}{2} \frac{\partial \phi}{\partial v}=v^{T} P-K^{T}=0  \tag{AII-5}\\
& \frac{1}{2} \frac{\partial \phi}{\partial \hat{x}}=K^{T} A=0 \tag{AII-6}
\end{align*}
$$

The transpose of the above two equations and the linearized mathematical model (eq. (AII-2)) make up the following least squares normal equations system:

$$
\begin{aligned}
& P V-K=0 \\
& A^{T} K=0 \\
& A \hat{X}+W-V=0
\end{aligned}
$$

Writing these equations in hypermatrix form yields the most expanded matrix form of the normal equations system as

$$
\left[\begin{array}{ccc}
P & -I & 0  \tag{AII-7}\\
-I & 0 & A \\
0 & A^{T} & 0
\end{array}\right]\left[\begin{array}{l}
V \\
K \\
\hat{X}
\end{array}\right]+\left[\begin{array}{c}
0 \\
W \\
0
\end{array}\right]=0
$$

The solution vector $\hat{\mathrm{X}}$ is obtained by using a matrix elimination technique [e.g Thompson, 1969]. Given the matrix equation system

$$
\left[\begin{array}{lr}
A & B \\
\hdashline C & D
\end{array}\right]\left[\begin{array}{c}
X \\
-Y
\end{array}\right]+\left[\begin{array}{c}
U \\
- \\
V
\end{array}\right]=0
$$

$X$ is eliminated by forming a modified coefficient matrix and known vector as follows:

$$
\begin{equation*}
\left[D-C A^{-1} B\right] Y+\left[V-C A^{-1} U\right]=0 \tag{AII-9}
\end{equation*}
$$

Applying this method to equation (AIII-7) to first eliminate $V$ gives

$$
\left\{\left[\begin{array}{ll}
0 & A \\
A^{T} & 0
\end{array}\right]-\left[\begin{array}{c}
-I \\
0
\end{array}\right] \quad P^{-1}[-I \quad 0]\right\}\left[\begin{array}{l}
K \\
\hat{X}
\end{array}\right]+\left\{\left[\begin{array}{l}
W \\
0
\end{array}\right]-\left[\begin{array}{c}
-I \\
0
\end{array}\right] \quad P^{-1}[0]\right\}=0
$$

or

$$
\left[\begin{array}{cc}
-P^{-1} & A \\
A^{T} & 0
\end{array}\right]\left[\begin{array}{l}
K \\
\hat{X}
\end{array}\right]+\left[\begin{array}{l}
W \\
0
\end{array}\right]=0
$$

(AII-10)

Using the same technique to eliminate K from equation (AII-10) yields

$$
\left[A^{T} P A\right] \hat{X}+A^{T} P W=0
$$

or, solving for $\hat{X}$

$$
\hat{X}=-\left[A^{T} P A\right]^{-1} \quad A^{T} P W
$$

(AII-11)
This is the least squares estimate for the solution vector $\hat{X}$. The parameters X are now computed as

$$
x=x^{\circ}+\hat{x}
$$

(AII-12)

Usually, the solution vector $\hat{X}$ is iterated (i.e. now $X^{\circ} \leftarrow X$, and a new $\hat{X}$ is computed) until it is very small (e.g. < 1 mm ). This is necessary because the use of a linear approximation is not exact (see Appendix I). The final expression for X is

$$
x=X^{0}-\left[A^{T} P A\right]^{-1} A^{T} P W
$$

(AII-13)

To find the variance covariance matrix $C_{x}$ of the parameters, the covariance law [e.g. Thomson et al., 1978] is used to propogate errors through equation (AII-13). Since the only independent random variable in equation (AII-13) is $L$ (because $W=F\left(X^{\circ}\right)-L$ ), then

$$
C_{x}=\left(\frac{\partial X}{\partial L}\right) \quad C_{L} \quad\left(\frac{\partial X}{\partial L}\right)^{T}
$$

(AII-14)

Realizing that

$$
\frac{\partial X}{\partial L}=\left[A^{T} P A\right]^{-1} \quad A^{T} P
$$

(AII-15)
since

$$
\frac{\partial W}{\partial L}=\frac{\partial\left(F\left(X^{\circ}\right)-L\right)}{\partial L}=-I
$$

then

$$
C_{x}=\left[A^{T} P A\right]^{-1} A^{T} P C_{L} P A \quad\left[A^{T} P A\right]^{-1}
$$

Noting that $P=\sigma_{0}^{2} C_{L}^{-1}$, then

$$
\begin{aligned}
& C_{x}=\sigma_{0}^{2}\left[A^{T} P A\right]^{-1} A^{T} P A\left[A^{T} P A\right]^{-1} \\
& C_{x}=\left[A^{T} C_{L}^{-1} A\right]^{-1}
\end{aligned}
$$

(AII-16)

Thus, $C_{x}$ is simply the normal equations inverse of the solution vector (see eq. (AII-11)).

Once the final solution vector is found, then the residual vector V is computed as

$$
V=\hat{A X}+W
$$

(AII-17)

From the residuals V , the a posteriori variance factor $\hat{\sigma}_{o}^{2}$ is evaluated as

$$
\begin{equation*}
\hat{\sigma}_{0}^{2}=\frac{\mathrm{v}^{\mathrm{T}} \mathrm{PV}}{\mathrm{df}} \tag{AII-18}
\end{equation*}
$$

$$
\text { where } \mathrm{df}=\text { degrees of freedom }=\mathrm{n}-\mathrm{u} .
$$

This a posteriori variance factor is useful when performing a post analysis of adjustment results (see chapter 9).

## APPENDIX III

## Error Ellipses

Error ellipses (see Figure AIII-1), are characterized by the length of their semimajor and semiminor axes $a$ and $b$, respectively and the azimuth $\theta$ of the semimajor axis a. These ellipses are representative of the error of a point in a network (sometimes called station ellipses) or of the error in the difference of coordinates between two points (relative error ellipses). These error ellipses are computed knowing the variance covariance matrix $C_{x}$ of the parameters, and the so-called $c$ factor. The c factor is used to increase the confidence level of the ellipse from standard(~39\%) to a desired (e.g. 95\%) confidence level in the following way :

$$
\begin{align*}
& \mathrm{a}=\mathrm{ca}  \tag{AII-1}\\
& \mathrm{~b}=\mathrm{cb} \\
& \mathrm{~s}
\end{align*}
$$

where $a_{s}$ and $b_{s}$ are the semimajor and semiminor axes of the standard error ellipse.

The basis of error ellipse computation lies in multivariate statistics [e.g. Wells and Krakiwsky, 1971; Hogg and Craig, 1970]. The quadratic form of the parameters for the a priori variance factor $\sigma_{0}{ }^{2}$ known is distributed as

$$
\begin{equation*}
x^{T} C_{x}^{-1} x=x_{u, 1-\alpha}^{2} \tag{AII-2}
\end{equation*}
$$

```
where \(X=\) difference between the least squares estimate of the parameters
    and the true value of the parameters,
    \(x_{u, 1-\alpha}^{2}=\) random variable with a chi-square distribution and degrees
    of freedom \(u\) (see Table AIII-l),
```



Figure AIII.l Error Ellipse

| $4{ }^{1-\alpha}$ | . 005 | . 010 | . 025 | . 050 | . 100 | . 250 | . 500 | . 750 | . 900 | . 950 | . 975 | . 900 | . 095 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | . 0000393 | . 000157 | . 000982 | . 00393 | . 0158 | . 102 | . 455 | 1.32 | 2.71 | 3.84 | 5.02 | 6.63 | 7.88 |
| 2 | . 0100 | . 0201 | . 0506 | . 103 | . 211 | . 575 | 1.39 | 2.77 | 4.61 | 5.99 | 7.38 | 9.21 | 10.6 |
| 3 | . 0717 | . 115 | . 216 | . 352 | . 584 | 1.21 | 2.37 | 4.11 | 6.25 | 7.81 | 9.35 | 11.3 | 12.8 |
| 4 | . 207 | . 297 | . 484 | . 711 | 1.06 | 1.92 | 3.36 | 5.39 | 7.78 | 9.49 | 11.1 | 13.3 | 14.9 |
| 5 | . 412 | . 554 | . 831 | 1.15 | 1.61 | 2.67 | 4.35 | 6.63 | 9.24 | 11.1 | 12.8 | 15.1 | 16.7 |
| 6 | . 676 | . 872 | 1.24 | 1.64 | 2.20 | 3.45 | 5.35 | 7.84 | 10.6 | 12.6 | 14.4 | 16.8 | 18.5 |
| 7 | . 989 | 1.24 | 1.69 | 2.17 | 2.83 | 4.25 | 6.35 | 9.04 | 12.0 | 14.1 | 16.0 | 18.5 | 20.3 |
| 8 | 1.34 | 1.65 | 2.18 | 2.73 | 3.49 | 5.07 | 7.34 | 10.2 | 13.4 | 15.5 | 17.5 | 20.1 | 22.0 |
| 9 | 1.73 | 2.09 | 2.70 | 3.33 | 4.17 | 5.90 | 8.34 | 11.4 | 14.7 | 16.9 | 19.0 | 21.7 | 23.6 |
| 10 | 2.16 | 2.50 | 3.25 | 3.94 | 4.87 | 6.74 | 9.34 | 12.5 | 16.0 | 18.3 | 20.5 | 23.2 | 25.2 |
| 11 | 2.60 | 3.05 | 3.82 | 4.57 | 5.58 | 7.58 | 10.3 | 13.7 | 17.3 | 19.7 | 21.9 | 24.7 | 26.8 |
| 12 | 3.07 | 3.57 | 4.40 | 5.23 | 6.30 | 8.44 | 11.3 | 14.8 | 18.5 | 21.0 | 23.3 | 26.2 | 28.3 |
| 13 | 3.57 | 4.11 | 5.01 | 5.89 | 7.04 | 9.30 | 12.3 | 16.0 | 19.8 | 22.4 | 24.7 | 27.7 | 29.8 |
| 14 | 4.07 | 4.66 | 5.63 | 6.57 | 7.79 | 10.2 | 13.3 | 17.1 | 21.1 | 23.7 | 26.1 | 29.1 | 31.3 |
| 15 | 4.60 | 5.23 | 6.26 | 7.26 | 8.55 | 11.0 | .4.3 | 18.2 | 22.3 | 25.0 | 27.5 | 30.6 | 32.8 |
| 16 | 5.14 | 5.81 | 0.91 | 7.96 | 9.31 | 11.9 | 15.3 | 19.4 | 23.5 | 26.3 | 28.8 | 32.0 | 34.3 |
| 17 | 5.70 | 6.41 | 7.56 | 8.67 | 10.1 | 12.8 | 16.3 | 20.5 | 24.8 | 27.6 | 30.2 | 33.4 | 35.7 |
| 18 | 6.26 | 7.01 | 8.23 | 9.39 | 10.9 | 13.7 | 17.3 | 21.6 | 26.0 | 28.9 | 31.5 | 34.8 | 37.2 |
| 19 | 6.84 | 7.63 | 8.91 | 10.1 | 11.7 | 14.6 | 18.3 | 22.7 | 27.2 | 30.1 | 32.9 | 36.2 | 38.6 |
| 20 | 7.43 | 8.26 | 9.59 | 10.9 | 12.4 | 15.5 | 19.3 | 23.8 | 28.4 | 31.4 | 34.2 | 37.6 | 40.0 |
| 21 | 8.03 | 8.90 | 10.3 | 11.6 | 13.2 | 10.3 | 20.3 | 24.9 | 29.6 | 32.7 | 35.5 | 38.9 | 41.4 |
| 22 | 8.64 | 9.54 | 11.0 | 12.3 | 14.0 | 17.2 | 21.3 | 26.0 | 30.8 | 33.9 | 36.8 | 40.3 | 42.8 |
| 23 | 0.26 | 10.2 | 11.7 | 13.1 | 14.8 | 18.1 | 22.3 | 27.1 | 32.0 | 35.2 | 38.1 | 41.6 | 44.2 |
| 24 | 9.89 | 10.9 | 12.4 | 13.8 | 15.7 | 19.0 | 23.3 | 28.2 | 33.2 | 36.4 | 39.4 | 43.0 | 45.6 |
| 25 | 10.5 | 11.5 | 13.1 | 14.6 | 16.5 | 10.9 | 24.3 | 29.3 | 34.4 | 37.7 | 40.6 | 44.3 | 46.9 |
| 26 | 11.2 | 12.2 | 13.8 | 15.4 | 17.3 | 20.8 | 25.3 | 30.4 | 35.6 | 38.9 | 41.9 | 45.6 | 48.3 |
| 27 | 11.8 | 12.9 | 14.6 | 16.2 | 18.1 | 21.7 | 26.3 | 31.5 | 36.7 | 40.1 | 43.2 | 47.0 | 49.6 |
| 28 | 12.5 | 13.6 | 15.3 | 16.9 | 18.9 | 22.7 | 27.3 | 32.6 | 37.9 | 41.3 | 44.5 | 48.3 | 51.0 |
| 29 | 13.1 | 14.3 | 16.0 | 17.7 | 19.8 | 23.6 | 28.3 | 33.7 | 39.1 | 42.6 | 45.7 | 49.6 | 52.3 |
| 30 | 13.8 | 15.0 | 16.8 | 18.5 | 20.6 | 24.5 | 29.3 | 34.8 | 40.3 | 43.8 | 47.0 | 50.9 | 53.7 |

Table AIII. $1 \chi^{2}$ Distribution

```
l-\alpha = desired confidence level (e.g. 0.95),
    u = dimensionality of the problem.
```

For the case of horizontal geodetic networks, $u=2$, and equation (AIII-2) is written as

$$
\left[\begin{array}{ll}
x & y
\end{array}\right]\left[\begin{array}{ll}
\sigma_{x}^{2} & \sigma_{x y} \\
\sigma_{x y} & \sigma_{y}^{2}
\end{array}\right]^{-1}\left[\begin{array}{l}
x \\
y
\end{array}\right]=x_{2,1-\alpha}^{2}
$$

(AIII-3)

The $C_{x}$ element in equations (AIII-2) and (AIII-3) is the submatrix for a single point of the full $C_{x}$ matrix for the whole network. An eigenvalue problem [e.g. Kreyszig,1972; Mikhail, 1976] is performed an equation (AIII-3) to transform it to an equation without cross product terms as follows;

$$
\left.\left[\begin{array}{ll}
x^{\prime} & y^{\prime}
\end{array}\right]\left[\begin{array}{ll}
\sigma_{\max }^{2} & 0 \\
0 & \sigma_{\min }^{2}
\end{array}\right]^{-1}\left[x^{\prime}\right] y^{\prime}\right]=x_{2,1-\alpha}^{2}
$$

where $x^{\prime}, y^{\prime}=$ transformed coordinates with respect to the rotated coordinate axes resulting from the eigenvalue problem, $\sigma_{\text {max }}^{2}=$ largest eigenvalue of $C_{x}$ (see eq. (AIII-5)), $\sigma_{\text {min }}^{2}=$ smallest eigenvalue of $C_{x}$ (see eq. (AIII-6)).

Specfically,the eigenvalues are

$$
\begin{aligned}
& \sigma_{\max }^{2}=\frac{1}{2}\left[\left(\sigma_{x}^{2}+\sigma_{y}^{2}\right)+\left\{\left(\sigma_{x}^{2}-\sigma_{y}^{2}\right)^{2}+4 \sigma_{x y}^{2}\right\}^{1 / 2}\right] \\
& \sigma_{\min }^{2}=\frac{1}{2}\left[\left(\sigma_{x}^{2}+\sigma_{y}^{2}\right)-\left\{\left(\sigma_{x}^{2}-\sigma_{y}^{2}\right)^{2}+4 \sigma_{x y}^{2}\right\}^{1 / 2}\right\}
\end{aligned}
$$

(AIII-5)
(AIII-6)

Writing out equation (AIII-4) explicitly results in

$$
\frac{\left(x^{\prime}\right)^{2}}{\sigma_{\max }^{2} x_{2,1-\alpha}^{2}}+\frac{\left(y^{\prime}\right)^{2}}{\sigma_{\min }^{2} x_{2,1-\alpha}^{2}}=1
$$

(AIII-7)
which is the familiar equation of an ellipse with axes $\left(\sigma_{\max }^{2} \chi_{2,1-\alpha}^{2}\right)^{1 / 2}$ and $\left(\sigma_{\min }^{2} X_{2,1-\alpha}^{2}\right)^{1 / 2}$. The standard error ellipse is found when $\chi_{2,1-\alpha}^{2}$ is equal to 1 , which corresponds to $(1-\alpha)=0.3935$, or a $39.35 \%$ confidence level. Thus

$$
\begin{aligned}
& a_{s}=\left(\sigma_{\max }^{2}\right)^{1 / 2} \\
& b_{s}=\left(\sigma_{\min }^{2}\right)^{1 / 2}
\end{aligned}
$$

(AIIE-8)
are the axes of the standard exror ellipse.
It is obvious from equation. (AIII-7) that the required c factor to compute a and b is

$$
c=\left(x_{2,1-\alpha}^{2}\right)^{1 / 2}
$$

(AIII-9)
for the case of $\sigma_{0}{ }^{2}$ assumed known. If the a priori variance factor is assumed unknown, however, then the a posterior variance factor $\hat{\sigma}_{0}^{2}$ (see Appendix II, eq. (AII-18)) is used to estimate the variance covariance matrix as $\hat{C}_{x}$, where the ${ }^{\wedge}$ stands for an estimated quantity. In this case, the quadratic form of the parameters is distributed as

$$
x^{T} \hat{C}_{x}^{-1} \quad x=u F_{u, d f, 1-\alpha}
$$

(AIII-10)
where $F=$ random variable with a Fischer distribution (see table AIII.2)
and degrees of freedom $u$ and $d f$,
df $=$ degrees of freedom of the adjustment.
Using the same development as for $\sigma_{0}^{2}$ known, the $c$ factor when using
the estimated variance factor $\hat{\sigma}_{o}^{2}$ to estimate $\hat{C}_{x}$ is

|  | 1 | 2 | 3 | 4 | 3 | 0 | 7 | 8 | 9 | 10 | 12 | 15 | 20 | 24 | 30 | 10 | 60 | 120 | $\infty$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 161.4 | 199.5 | 215.7 | 224.6 | 230.2 | 234.0 | 230.8 | 238.9 | 240.5 | 21.9 | 243.9 | 215.9 | $2: 30$ | $2 \cdot 13.1$ | 250.1 | 251.1 | 252.2 | 253.3 | 25.3 |
| . | 18.51 | 19.00 | 19.16 | 10.25 | 19.30 | 19.33 | 19.35 | 10.37 | 19.38 | 19.40 | 19.41 | 19.43 | 19.45 | 19.45 | 19.15 | 19.47 | 19.45 | 10.49 | 18.50 |
| 3 | 10.13 | 9.55 | 9.28 | 9.12 | 9.01 | 8.94 | 8.83 | 8.85 | 8.81 | 8.78 | 8.74 | 8.70 | 8.60 | 8.64 | 8.62 | 8.50 | 3.57 | 8.55 | 8.53 |
| 4 | 7.71 | 6.56 | 6.59 | 5.39 | 6.20 | 6.16 | 0.03 | 6.01 | 6. 0 | 5.00 | 3.91 | 3.86 | 5.90 | 5.77 | 5.75 | 3.72 | 5.62 | 3.6 | 3.63 |
| 5 | 0.61 | 5.79 | 5.41 | 5.19 | 5.05 | 4.95 | 4.83 | 4.32 | 4.77 | 4.74 | 4.68 | 4.62 | 4.56 | 4.53 | 4.50 | 4.46 | 4.43 | 8.40 | 4.38 |
| 6 | 5. 99 | 5.14 | 4.76 | 4.53 | 4.39 | 4.28 | 4.21 | 4.15 | 4.10 | 4.05 | 4.00 | 3.94 | 3.37 | 3.34 | 3.81 | 3.77 | 3.16 | $3 . \%$ | 3.67 |
| 7 | 5.53 | 4.74 | 4.35 | 4.12 | 3.07 | 3.87 | 3.79 | 3.73 | 3.65 | 3.64 | 3.5 ? | 3.51 | 3.44 | 3.41 | 3.34 | 3.34 | 3.30 | 3.27 | 3.43 |
| 8 | 5.32 | 4.10 | 4.07 | 3.84 | 3.69 | 3.58 | 3.50 | 3.44 | 3.39 | 3.35 | 3.28 | 3.22 | 3.15 | 3.12 | 3.8 | 3.05 | 3.01 | 3.07 | 2.95 |
| , | 4.12 | 4.20 | 3.86 | 3.03 | 3.48 | 3.37 | 3.25 | 3.23 | 3.15 | 3.14 | 3.07 | 3.04 | 2.94 | 2.80 | 2.85 | 2.3 | 2.30 | 2.75 | 2.11 |
| 10 | 4.05 | 4.10 | 3.71 | 3.48 | 3.33 | 3.22 | 3.14 | 3.67 | 3.02 | 2.98 | 2.51 | 2.95 | 2.77 | 2.74 | 2.70 | 2.66 | 2.62 | 2.58 | 2.51 |
| 11 | 4.54 | 3.08 | 3.54 | 3.35 | 3.20 | 3.09 | 3.01 | 2.45 | 2.30 | 2.85 | 2.79 | 2.72 | 2.65 | 2.61 | 2.57 | 2.53 | 2.49 | 2.45 | 2.40 |
| 12 | 4.75 | 3.83 | 3.43 | 3.25 | 3.11 | 3.00 | 2.91 | 2.35 | 2.80 | 2.75 | 2.43 | 2.62 | 2.54 | 2.31 | 2.47 | 2.45 | 2.35 | 2.35 | 2.30 |
| 13 | 4.67 | 3.81 | 3.41 | 3.18 | 3.03 | 2.92 | 2.83 | 2.77 | 2.71 | 2.67 | 2.0 | 2.33 | 2.45 | 3.42 | 2.35 | 2.34 | 2.02 | 2.25 | 2.21 |
| 14 | 4.60 | 3.74 | 3.34 | 3.11 | 2.26 | 2.85 | 2.75 | 2.70 | 2.65 | 2.60 | 2.53 | 2.46 | 2.39. | 2.35 | 3.31 | 2.27 | 2.22 | 2.16 | 2.13 |
|  | 4.54 | 3.98 | 3.23 | 3.05 | 2.00 | 2.70 | 2.71 | 2.64 | 2.59 | 2.54 | 2.18 | 2.40 | 2.33 | 2.29 | 2.25 | 2.22 | 2.6 | 2.11 | 2.07 |
| 16 | 4.43 | 3.63 | 3.24 | 3.04 | 2.85 | 2.74 | 2.63 | 2.59 | 2.54 | 2.40 | 2.42 | 2.35 | 2.20 | 8 | \%. 18 | 2.15 | 2.11 | 2.03 | 2.05 |
| 17 | 4.45 | 3.59 | 3.20 | 2.96 | 2.81 | 2.70 | 2.61 | 2.5 | 2.43 | 2.45 | 2.38 | 2.31 | 2.23 | 2.15 | 2.6 | 4.10 | 2.6 | 2.01 | 1.45 |
| 18 | 4.41 | 3.55 | 3.19 | 2.93 | 9.17 | 2.06 | 2.5 | 2.54 | 2.46 | 2.45 | 2.84 | 2.77 | 2.19 | 2.15 | \% | 2. ${ }^{2}$ | 3.02 | 1. 6 | - 02 |
| 19 | 4.38 | 3.52 | 3.13 | 2.00 | 2.74 | 2.63 | 2.51 | 2.45 | 2.42 | 2.8 | 2.31 | 2.23 | 2.16 | 2.14 | 2.07 | 2.4 | 1.63 | 1.93 | 1.83 |
| 20 | 4.35 | 3.43 | 3.10 | 2.87 | 2.71 | 2.60 | 2.51 | 2.45 | 2.39 | 2.3.3 | 2.23 | 2.20 | 2.:2 | 2.05 | 2.64 | 2.93 | 1.85 | 1.00 | 1.84 |
| 21 | 4.32 | 3.47 | 3.07 | 2.64 | 2.88 | 2.57 | 2.14 | 2.42 | 2.37 | 2.32 | 2.25 | 2.18 | 2.10 | 2, 2.0 | 2.01 | 1.j5 | 1.92 | 1.57 | 1.81 |
| 23 | 4.30 | 3.4 \% | 3.05 | 2.82 | 2.64 | 2.5 | 2.46 | 2.40 | 2.34 | 2.30 | 2.23 | 2.15 | 2.07 | 2.03 <br> 2.01 | 1.03 | 1.46 | 1.80 1.80 | 1.84 1.84 | 1.78 1.76 |
| 23 | 4.28 | 3.12 3.40 | 3.03 3.01 | 2.80 2.76 | 2.64 2.62 | 2.54 |  | 2.37 |  | 2.27 2.25 | 近 | 2.13 | 2.05 | ? | 1.96 | 1.91 | 1.85 1.51 | 1.81 | 1.76 1.73 |
| 24 | 4.26 | 3.40 | 3.04 | 2.78 | 2.62 | 2.51 | 2.42 |  | 2.30 | 2.25 | 2.15 | 2.1 | 2.6 | 1.05 | ..4t | 1.30 | 1.5 | 1.10 | 1.43 |
| 25 | 4.24 | 3.33 | 2.29 | 2.75 | 2.00 | 2.19 | 2.10 | 2.34 | 2.23 | 2.21 | 2.46 | 2.0 | 2.01 | 1.00 | 1.92 | 1.87 | 1.82 | $1.7 \%$ | 1.74 |
| 26 | 4.23 | 3.37 | 2.03 | 2.74 | 2.59 | 2.47 | 2.35 | 2.32 | 2.27 | 2.22 | 2.15 | 2.07 | 1.92 | 1.45 | 1.80 | 1.35 | 1.3 |  |  |
| 27 | 4.21 | 3.35 | 2.0.6 | 2.73 | 2.57 | 2. 46 | 2.37 | 2.31 | 2.25 | 2.20 | 2.13 <br> 2.12 <br> 2.12 | 2.6 | 1.97 | 1.93 | 1.68 | 1.85 | 1.79 | 1.71 | 1.67 1.65 |
| 28 | 4.14 | 3.45 | 2. 2.45 | 2.71 | 2.55 2.55 | 2.15 2.43 | 2.35 2.35 |  | 2.22 | 2.18 | 2.12 2.10 |  | 1.96 | 1.00 | 1.35 | 1.82 | :\% | 1.7 | 1.64 |
| 24 | 4.15 | 3.33 | 2.43 | 2.70 | 2.55 | 2.43 | 2.35 | 2.28 | 2.22 | 2.15 | 2.0 | 2.03 | 1.41 | 1.0 |  |  |  |  |  |
| 30 | 4.17 | 3.32 | 2.02 | 2.63 | 2.53 | 2.42 | 2.33 | 2.27 | 2.21 | 2.16 | 2.09 | 2.01 | 1.93 | 1.39 | 1.84 | 1.79 | 1.74 | 1. 68 | 1. 62 |
| 40 | 4.03 | 3.23 | 2.84 | 2.61 | 2.45 | 2.34 | 2.25 | 2.18 | 2.12 | 2.08 | 2.00 | 1.92 | 1.84 | 1.79 1.70 | 1.74 1.85 | 1.69 | 1.6 | 1.53 | 1.51 |
| 60 | $4 . \infty$ | 3.15 | 2.76 | 2.53 | 2.37 | 2.25 | 2.17 | 2.10 | 2.64 | 1.99 | 1.92 | 1.84 | 1.75 <br> 1.65 | 1.70 1.61 | 1. 1.55 | 1.50 | 1.5 | 1.18 | 1.25 |
| 120 | 3.92 3.51 | 3.07 3.6 | 2. 2.60 | 2.45 2.37 | 2.21 | 2.17 2.10 | 2.02 | 2.02 | 1.56 | 1.91 | 1.75 | 1.76 | 1.57 | 1.61 | 1.45 | 1.02 | 1.32 | 1.22 | $1 . \infty$ |

Table AIII. 2 F Distribution for $(1-\alpha)=0.95$

$$
\begin{equation*}
\mathrm{c}=\left(2 \mathrm{~F}_{2, \mathrm{df}, 1-\alpha}\right)^{1 / 2} \tag{AIII-1I}
\end{equation*}
$$

Equations (AIII-9) and (AIII-11) above give the $c$ factor for computation of a single error ellipse without regard to other stations in the network. If, however, it is required that $N$ station ellipses all have the desired confidence level (1- 1 ) simultaneously, then the $c$ factor is computed as

$$
\begin{equation*}
c=\left(\chi_{2,1-\alpha / \bar{N}}^{2}\right)^{1 / 2} \tag{AIII-12}
\end{equation*}
$$

or

$$
\begin{equation*}
c=\left(2 F_{2, d f, 1-\alpha / N}\right)^{1 / 2} \tag{AIII-13}
\end{equation*}
$$

where $\alpha$ has been replaced by $\alpha / N$. This is a direct result of Bonferroni's inequality [Vanicek and Krakiwsky, in prep.] which states that the given confidence level is at least 1- $\alpha$ for the simultaneous case.

The orientation of the error ellipse is given by the normalized eigenvector corresponding to the eigenvalue $\sigma^{2} \max ^{2}$ of the eigenvalue problem performed on equation (AIII-3). The azimuth of the semimajor axis is thus

$$
\begin{equation*}
\theta=\operatorname{sign}\left(\sigma_{x y}\right) \cdot \arccos \left[\frac{\left(\sigma_{\max }^{2}-\sigma_{x}^{2}\right)}{\left(\sigma_{x y}^{2}+\left(\sigma_{\max }^{2}-\sigma_{x}^{2}\right)^{2}\right) 1 / 2}\right] \tag{AIII-14}
\end{equation*}
$$

where $\sigma_{\max }^{2}$ is given by equation (AIII-5).
The computation of relative error ellipses is facilitated by
applying the covariance law to the following expressions:

$$
\begin{align*}
& \Delta x_{i j}=x_{j}-x_{i},  \tag{AIII-15}\\
& \Delta y_{i j}=y_{j}-y_{i} .
\end{align*}
$$

This gives

$$
C_{\Delta x, \Delta y}=B C_{x, y} B^{T}
$$

where $B=$

$$
\left[\begin{array}{cccc}
\frac{\partial \Delta x_{i j}}{\partial x_{i}} & \frac{\partial \Delta x_{i j}}{\partial y_{i}} & \frac{\partial \Delta x_{i j}}{\partial x_{j}} & \frac{\partial \Delta x_{i j}}{\partial y_{j}} \\
\frac{\partial \Delta y_{i j}}{\partial x_{i}} & \frac{\partial \Delta y_{i j}}{\partial y_{i}} & \frac{\partial \Delta y_{i j}}{\partial x_{j}} & \frac{\partial \Delta y_{i j}}{\partial y_{j}}
\end{array}\right]
$$

or carrying out the partial derivatives and writing $C_{x, y}$ in full

$$
\begin{aligned}
& \text { (AIII-16) }
\end{aligned}
$$

- Thus, to compute the standard relative error ellipse between points i and $j$, the equations for station ellipses (i.e. eqs. (AIII-5), (AIII-6), (AIII-14)) are employed, but making the substitutions

$$
\begin{aligned}
& \sigma_{x}^{2}=\sigma_{x_{i}}^{2}-2 \sigma_{x_{i} x_{j}}+\sigma_{x_{j}}^{2} \\
& \sigma_{x y}=\sigma_{x_{i} y_{i}}-\sigma_{y_{i} x_{j}}-\sigma_{x_{i} y_{j}}+\sigma_{x_{j} y_{j}} \\
& \sigma_{y}^{2}=\sigma_{y_{i}}^{2}-2 \sigma_{y_{i} y_{j}}+\sigma_{y_{j}}^{2}
\end{aligned}
$$

(AIII-17)

The Surveys and Mapping Branch of the Dept. of Energy, Mines and Resources uses relative error ellipses to classify different order surveys [Energy, Mines and Resources, 1973]. A survey station of a network is classified according to whether the semimajor axis of the $95 \%$ confidence ellipse with respect to other stations of the network is less than or equal to

$$
r=\mathrm{k} \mathrm{~d}
$$

(AIII-18)
where $r=$ radius of an error circle in $c m$ (see Figure AIII. 2 ),
$\mathrm{d}=$ distance in km to any station,
$\mathbf{k}=$ factor assigned according to the order of survey (see Table AIII.3).


Figure AIII.2. Error Circle and Confidence Ellipse

Thus, if $a<r$ for all the relative error ellipses (at $95 \%$ confidence level) between station $i$ and the rest of the network stations, then station $i$ is classified in that specific order of survey. For example, a

| order | k | $r$ in pprn |
| :---: | :---: | :---: |
| 1st | 2 | 20 |
| 2nd | 5 | 50 |
| 3rd | 12 | 120 |
| 4th | 30 | 300 |

Table AIII.3. Horizontal Survey Classification.
second order survey station must have the serimajor axis a of the 95\% relative error ellipse less than 25 cm for stations 5 km apart.

