

An Empirical Approach for the Estimation of GPS Covariance Matrix of Observations

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BIOGRAPHY

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ABSTRACT

The covariance matrix of observations plays an important role in GPS data processing. For example, the weighting of the observations are based on the covariance matrix of the observations. The covariance matrix of the observations also plays a role in the estimation of the covariance matrix of the parameters. In this paper we present an empirical stochastic approach to create covariance matrices for GPS. Our approach aims at a more realistic and complete information about the stochastic behavior of GPS observations. As a result, improvement in quality and quality control of estimated coordinates using GPS measurements is obtained.

In the present case study, baselines are processed using pseudorange. The observations are considered as a time series after double differencing. The advantages and

disadvantages of this empirical stochastic approach are explored, comparing it with more usual methods used for building observations covariance matrix. An analysis is made in terms of obtained coordinates where baselines between points with known coordinates were processed.

Two baselines with length of approximately 2 km and 20 km were processed using three different weighting schemes for observations. In general, when the empirical approach was used a good improvement in the bias was obtained. The mean biases were at least 44 % smaller in the 2 km baseline case and at least 21 % in the 20 km baseline case. The model certainly brings improvement to the solution, with better results achieved for height determination, with improvements up to 72 % in mean bias for the shorter baseline.

INTRODUCTION

The covariance matrix of observations plays a fundamental role in GPS data adjustment. Usually the method of least squares is used to compute the coordinates of a given receiver. In the adjustment the covariance matrix of observations drives how each one of the observations will contribute for the update of the parameters. This matrix includes not just variances, but also a relation between all observations, the covariances. An usual approach to estimate the covariance matrix of observations is to set a weight for all observations and then, in case of baseline, propagate it using the double difference operator. It is not assumed any correlation between observations made in two different epochs, the so called autocorrelations. The weights can be set based on different types of information, such as elevation angle of the satellite or signal-to-noise ratio. Sometimes even the identity matrix (equally weighted observations) is used.

In this work, an approach to build the covariance matrix of observations for GPS data processing is presented. The main objective of such technique is to populate the covariance matrix with realistic information, estimated by means of a stochastic analysis of the raw data. Trying to

have as much as possible information inside the covariance matrix is a way to carry into the adjustment model a realistic picture of the quality and behavior of the observations. Eventually, the procured parameters will be adjusted using this matrix. The ultimate goal of this work is to improve the quality of GPS data processing.

In order to analyze the data assuming a stochastic behavior a few requirements need to be satisfied, such as stationarity. Because of that, the GPS data need to undergo some modifications before it can be used. In our Empirical Stochastic model (which will be called herein with ESto model) all analysis is made based on raw data, without any external information, before the adjustment.

In next Section a background explanation about stochastic processes and their analysis will be made. Next, the treatment given to GPS data in order to make its analysis as a stochastic process possible will be shown. A case study was carried out and its results are shown, as well as conclusions and recommendations for future work.

STOCHASTIC PROCESSES

A time series is a series of observations made through time. An important feature of a series is that, usually, the observations made at subsequent epochs are dependent on each other. The analysis of a time series is based on this dependency. In order to carry out the analysis, generally stochastic models are used.

The utilization of mathematical models to describe physical phenomena is very common. Sometimes a model based on physical laws can be capable of determining the value of a given variable for a specified time. This is possible when the variable (in our case the GPS observable) is totally deterministic.

A stochastic process is a statistical phenomena which occurs through time according to probability laws, [Box et. al., 1994]. A time series can be considered as a realization of a stochastic process.

There is a particular process, called a stationary process, which has a particular statistic equilibrium state. A stochastic process is called widely stationary when its properties remain unaffected when the time origin is changed. This means that the joint probability function of a process with n observations $z_{t_1}, z_{t_2}, \dots, z_{t_n}$, observed at time instants t_1, t_2, \dots, t_n is the same as a process with n observations $z_{t_1+k}, z_{t_2+k}, \dots, z_{t_n+k}$, observed at time instants $t_1+k, t_2+k, \dots, t_n+k$ for every integer k .

If a process is stationary, we can say that:

$$\rho(z_t) = \rho(z_{t+k}) \forall t \geq 0, k \geq 0, \quad (1)$$

where $\rho(z_t)$ is the probability density function, and because of that, we can conclude that the mean of a stationary process is constant:

$$\mu_t = E(z_t) = \int_{-\infty}^{\infty} z \cdot \rho(z) dz, \quad (2)$$

and therefore:

$$\mu_t = \mu_{t+k}, \quad (3)$$

for every integer k . The mean of a process \bar{z} can be estimated with the equation:

$$\bar{z} = \frac{1}{n} \cdot \sum_{t=1}^n z_t. \quad (4)$$

The variance of a process, defined as:

$$s^2 = E[(z_t - \bar{z})^2] = \int_{-\infty}^{\infty} (z - \bar{z})^2 \rho(z) dz, \quad (5)$$

can be estimated with:

$$s^2 = \frac{1}{n} \cdot \sum_{t=1}^n (z_t - \bar{z})^2. \quad (6)$$

Based on these equations and the assumption behind them, the variances for GPS observables are computed in the ESto model.

The stationarity of a process implies that the joint probability function $\rho(z_{t_1}, z_{t_2})$ is the same for any instants t_1 and t_2 , for the same time interval between them. This means that we can estimate the joint probability function of a process for different time intervals k . The covariance between the values z_t and z_{t+k} , separated by a time interval k , which is constant for every t for a stationary process, is called the auto-covariance function, and can be defined as:

$$\gamma_k = \text{cov}(z_t, z_{t+k}) = E[(z_t - \bar{z})(z_{t+k} - \bar{z})]. \quad (7)$$

It is possible to estimate the auto-correlation function of a process for a given lag k with the following equation:

$$\rho_k = \frac{E[(z_t - \bar{z})(z_{t+k} - \bar{z})]}{\sqrt{E[(z_t - \bar{z})^2] \cdot E[(z_{t+k} - \bar{z})^2]}} = \frac{\sum_{t=1}^{n-k} [(z_t - \bar{z})(z_{t+k} - \bar{z})]}{s^2}. \quad (8)$$

ESTO MODEL

In ESto model, the previously described stochastic assumptions are used to estimate the correlations between different observations (made at different time and/or from different satellite pairs), as well as the variances for each one of the observed satellites. The input of the model is the raw data, and the output is the stochastic parameters.

As stated in the previous section, a stochastic process needs to have certain characteristics in order to be used for estimation of variances and covariances. Since the objective of this work is to analyze GPS data as a stochastic process, it is necessary to modify the original time series of observations to satisfy the assumptions made in stochastic analysis. The main assumption that will be discussed here is the stationarity of the process. It is clear that GPS observations are not stationary, because they are related to the satellite orbits and always vary with time. Figure 1 shows an example of double differenced pseudoranges through time.

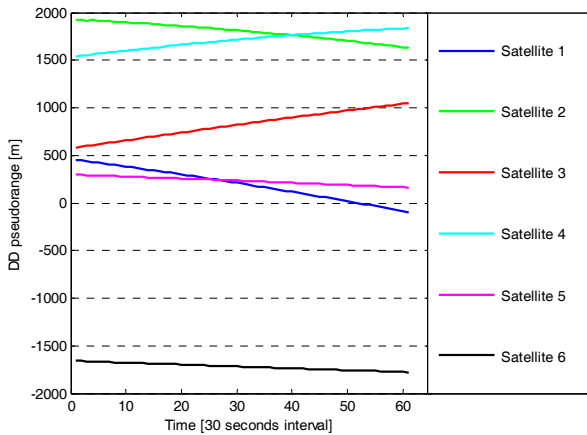


Figure 1. Double difference pseudoranges time series.

Clearly these observations could not be used directly in a stochastic analysis, and therefore need to be modified. In order to get a stationary series from the measurements, they are reduced by using the approximate geometric distances between receiver and satellite antennas, according to:

$$\Delta \nabla P_{red} = \Delta \nabla P_{obs} - \Delta \nabla \rho, \quad (9)$$

where $\Delta \nabla P_{obs}$ is the double differenced observed pseudorange, $\Delta \nabla \rho$ is the double differenced geometric distance between receiver and satellite antennas and $\Delta \nabla P_{red}$ is the double differenced reduced pseudorange. Figure 2 shows the double differenced C/A pseudoranges before and after the reduction made in the ESto model, for each one of the satellites shown above.

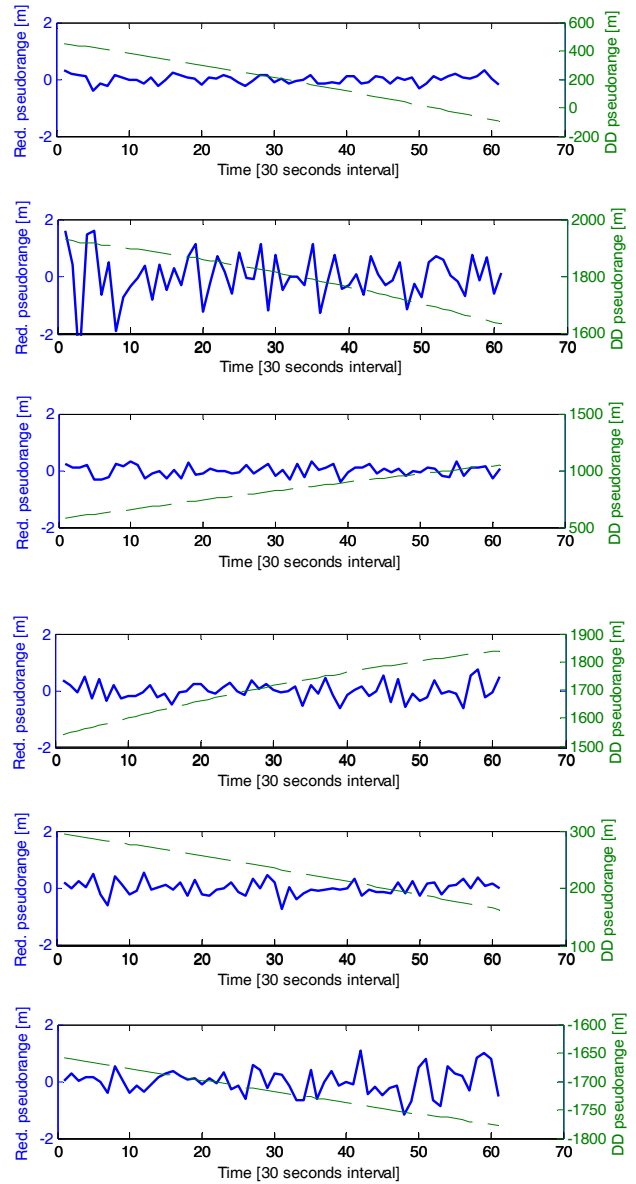


Figure 2. Reduction of DD pseudoranges, in the ESto model.

In the plots above, the dashed thin lines (green) represent the original series, and the continuous thick lines (blue) represent the reduced series. After this step the reduced series are used in the stochastic analysis. This part of the ESto model is based on the approaches mentioned in the previous Section.

As an illustration of the stochastic part of the model, Figure 3 shows the auto-correlation function estimated with ESto model and also computed based on the adjustment residuals, for satellites with different elevation angles (65, 31 and 17 degrees). The observables are C/A pseudoranges, observed for a short baseline (5 km) during half hour. The sampling interval is 30 seconds, thus one epoch lag represents 30 seconds in time.

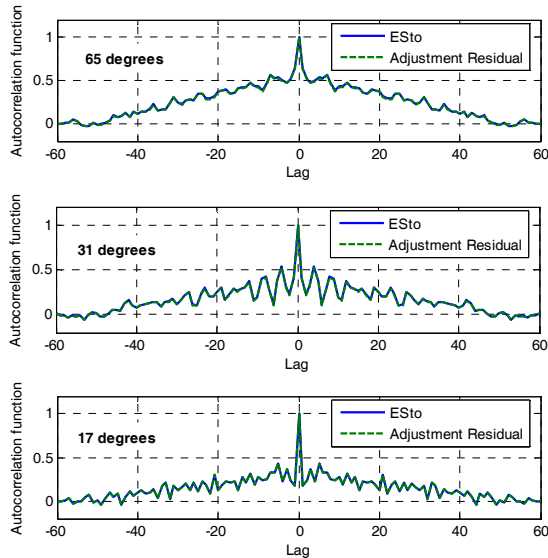


Figure 3. Auto-correlation function for satellites with different elevation angles.

It can be seen that ESto provides a good approximation of the residuals auto-correlation (the two curves overlap most of the time) and therefore for the observations. It can be also noticed how the auto-correlations get smaller as the elevation angle gets lower.

Another issue is the variance which is assigned to each satellite. The ESto model computes the variance by analyzing the raw data. Figure 4 shows the pseudorange standard deviations estimated by ESto for different elevation angles.

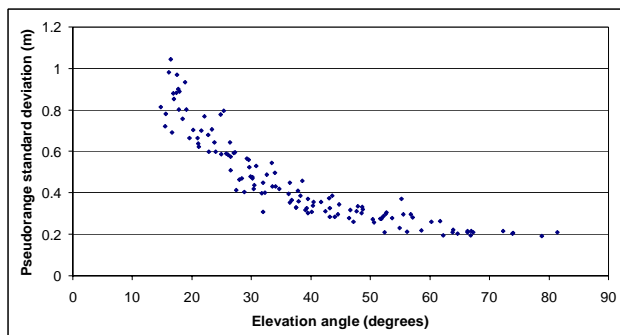


Figure 4. Estimated standard deviation for different elevation angles.

The standard deviation varies with the elevation angle, with an exponential shape, between around 0.2 m and 1 m. Figure 5 shows the agreement between the estimated standard deviations and the residuals standard deviation after the adjustment.

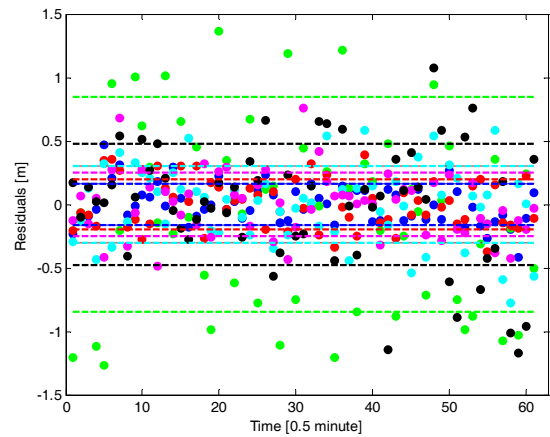


Figure 5. Standard deviations estimated with ESto and adjustment residuals.

In Figure 5, the dots represent the residuals of each of the observations during a half hour observation session with 30 seconds sampling interval (7 satellites observed). The dashed lines represent the one sigma standard deviations estimated with ESto. For each satellite, the color of the lines and dots is the same. Table 1 shows the computed values for both cases.

Table 1 – Comparing adjustment residuals and ESto standard deviations.

Satellite # (color on plot) – elevation angle	Standard deviation	
	Residuals [m]	ESto [m]
1 (blue) – 65°	0.1588	0.1589
2 (green) – 17°	0.8452	0.8453
3 (red) – 52°	0.1931	0.1931
4 (cyan) – 44°	0.3043	0.3043
5 (magenta) – 45°	0.2516	0.2516
6 (black) – 31°	0.4764	0.4763

ESto model has been developed in order to allow a estimation of fully populated observation covariance matrices for GPS. In terms of validation, it has been compared with other techniques used to build the same matrices. The first one, called here Formal DD, the simplest of the ones explored in this work, is a matrix generated by propagating an identity matrix with the double difference operator, thus all satellites have the same weight. The second approach uses the elevation angle as a parameter to estimate the standard deviation of the observations (here this approach is being called Elevation based). After the standard deviations are estimated, they are propagated using the double difference operator as well. Figure 6 shows a representation of covariance matrices estimated using ESto, Formal DD and elevation based:

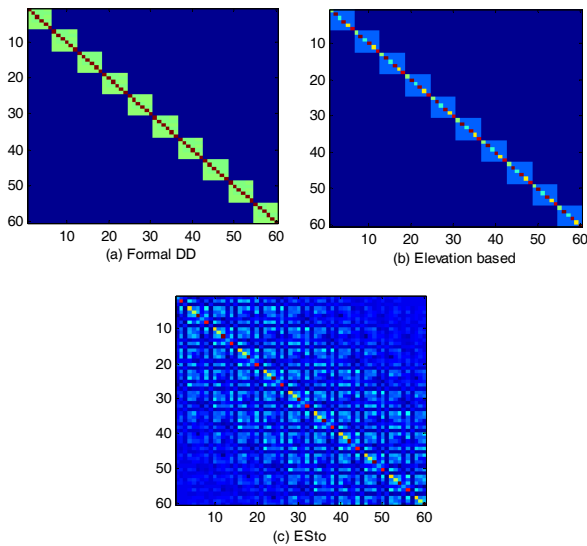


Figure 6. Representation of covariance matrices estimated using different schemes.

In the figure above, different colors represent different values within the matrices. Each one of them has its own scale, and they were generated from the same data set. In the case of Formal DD (Figure 6a), it can be noticed that the elements on the main diagonal have always the same color, what means the same variance. The green squares around the main diagonal are the correlation between different satellite pairs within the same epoch. This correlation exists with this pattern due to the use of a common reference satellite. In the second case (Elevation based – Figure 6b), the main difference from the previous in the different weights for different satellites (this can be noticed by the different colors in the main diagonal). The blue squares represent correlation between different satellite pairs within the same epoch, as in the previous case. In Figures 6a and 6b the dark blue means elements with value equal to zero. When the ESto model is used (Figure 6c), it can be noticed that a large amount of information is placed into off-diagonal elements of the covariance matrix. These values were derived from raw data, and represent correlation among observations for the same epochs and for different epochs. The next step to validate the ESto model is an analysis with real data, explored in the next Section.

DATA PROCESSING

In order to investigate the efficiency of this empirical stochastic approach with respect to other conventional techniques, two short baselines were processed using C/A pseudoranges. The choice of such data set is justified by the elimination of several errors in the measurements, such as clock errors and atmospheric refraction, which would require a special treatment in the stochastic analysis if they were not eliminated. Datasets of 24 hours were processed in half hour batch adjustments along the

day. This was done for two baselines, UNB1-FRDN and GAGS-NRC1, with approximately 2 km and 20 km, respectively. Figure 7 shows the results of the adjustment for the UNB1-FRDN baseline. The mean errors (as compared to the known station coordinates) and their RMS are shown in Table 2.

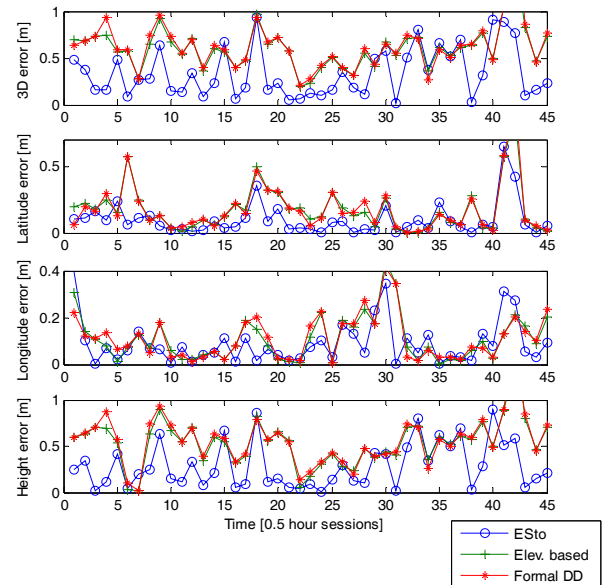


Figure 7. Results for the adjustment of the baseline UNB1-FRDN.

Table 2 indicates that when ESto was used, the mean bias was lower than the other solutions. However, the standard deviation didn't improve as much. The ESto solutions show some errors quite larger than their mean values, which can be affecting standard deviation values. The use of ESto model provided better solutions in 90% of the situations for this baseline.

Table 2

		Formal DD	Elev. based	ESto
Mean error (m)	Latitude	0.146	0.137	0.051
	Longitude	-0.081	-0.076	-0.024
	Height	0.543	0.523	0.153
	3D	0.632	0.614	0.341
RMS (m)	Latitude	0.257	0.248	0.155
	Longitude	0.146	0.146	0.128
	Height	0.623	0.599	0.380
	3D	0.689	0.664	0.430

Tables 3 and 4 show the absolute and relative improvements achieved by the use of ESto with respect to the other models.

Table 3

		w.r.t. Formal DD	w.r.t. Elev. based
Mean error (m)	Latitude	0.10	0.09
	Longitude	0.06	0.05
	Height	0.39	0.37
	3D	0.29	0.27
RMS (m)	Latitude	0.10	0.09
	Longitude	0.02	0.02
	Height	0.24	0.22
	3D	0.26	0.23

Table 4

		w.r.t. Formal DD	w.r.t. Elev. based
Mean error	Latitude	66%	64%
	Longitude	70%	68%
	Height	72%	71%
	3D	46%	44%
RMS	Latitude	40%	38%
	Longitude	12%	12%
	Height	39%	37%
	3D	38%	35%

ESto model provided improvements ranging from 5 cm to 37 cm in bias and 2 cm to 26 cm in RMS. The greatest improvements were achieved for height determination, and consequently also for three-dimensional positioning. In terms of relative results, the improvements are very good, ranging from 44 % to 72 % for mean bias and 12 % to 40 % for RMS. These results show that the use of the empirical approach can definitely bring advantages in terms of quality of the estimated coordinates.

Further analysis was carried out with the 20km baseline, in order to validate the model. Figure 8 shows the results for the GAGS-NRC1 baseline. Table 5 summarizes the statistics of the results.

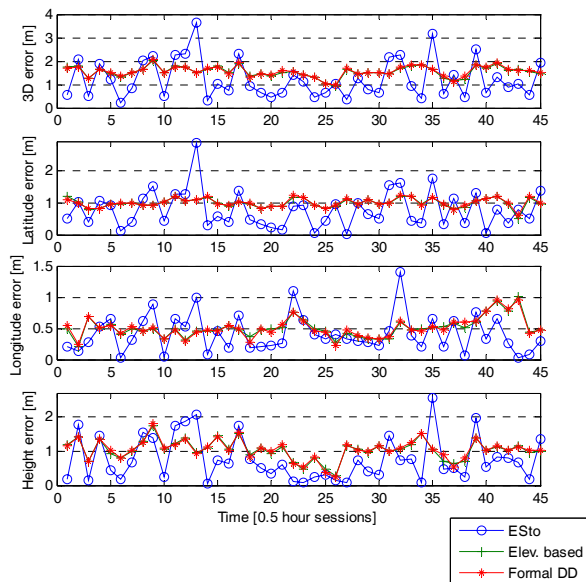


Figure 8. Results for the adjustment of the baseline GAGS-NRC1.

Likewise the previous situation when ESto was used the mean bias (with respect to known coordinates) was lower than for other solutions. The RMS is also lower for ESto solutions. The use of ESto results in some solutions with values larger than the mean values of the errors, which could explain the standard deviation values. The use of ESto model provided better solutions in only 70% of the situations in this case.

Table 5

		Formal DD	Elev. based	ESto
Mean error (m)	Latitude	-0.993	-0.986	-0.762
	Longitude	0.508	0.506	0.395
	Height	1.050	1.036	0.696
	3D	1.558	1.543	1.223
RMS (m)	Latitude	1.003	0.996	0.948
	Longitude	0.533	0.529	0.508
	Height	1.090	1.076	1.000
	3D	1.574	1.560	1.473

Tables 6 and 7 show the absolute and relative improvements achieved by the use of ESto with respect to the other models.

Table 6

		w.r.t. Formal DD	w.r.t. Elev. based
Mean error (m)	Latitude	0.23	0.22
	Longitude	0.11	0.11
	Height	0.35	0.34
	3D	0.34	0.32
RMS (m)	Latitude	0.06	0.05
	Longitude	0.03	0.02
	Height	0.09	0.08
	3D	0.10	0.09

Table 7

		w.r.t. Formal DD	w.r.t. Elev. based
Mean error (m)	Latitude	23%	23%
	Longitude	22%	22%
	Height	34%	33%
	3D	22%	21%
RMS (m)	Latitude	5%	5%
	Longitude	5%	4%
	Height	8%	7%
	3D	6%	6%

For the GAGS-NRC1 baseline ESto model provided improvements ranging from 11 cm to 34 cm in bias and 2 cm to 10 cm in RMS. The greatest improvements were achieved for height determination, and consequently also for three-dimensional positioning, as in the previous solution. In terms of relative results, the improvements are very good, ranging from 21 % to 34 % for mean bias and 4 % to 8 % for RMS. These results show that for this case the empirical approach did not show great improvements, although the results are generally better than using other weighting schemes.

CONCLUSIONS AND FUTURE WORK

It was shown that the empirical model is capable of providing a good stochastic modeling of observables without any external source of information, before the adjustment process.

The implementation of the model is computationally efficient, and can be easily implemented for GPS data processing. The increase in computational cost is largely compensated by the benefit in terms of adjustment results. The additional computational time is due to the computation of auto-correlation and cross-correlation functions that involve several combinations in order to have fully populated covariance matrices.

Two baselines (2 km and 20 km long) were processed using three different weighting schemes for observations. In general, when ESto was used an improvement in bias was obtained. The mean biases were at least 44 % smaller in the first case and at least 21 % in the second one. The RMS values had improvements of at least 12 % for the shorter baseline and at least 4 % for the longer one.

The larger improvement was achieved for height determination (up to 72% in mean bias for the shorter baseline).

In terms of standard deviation of the solutions ESto showed little or no improvement. The solutions still show large variability with a consequent increase in their spread. There is room for improvement. A decrease in the spread would help to make the standard deviations lower.

The results for the shorter baseline (2 km) provided the best results so far. For the other longer baseline (20 km), additional treatment in the stochastic analysis is necessary to improve the results.

Future work is needed to investigate the source of solution outliers and enhancements in the stochastic analysis for longer baselines.

Further research to apply ESto to point positioning and carrier phase measurements is also a planned future step in the development and validation of the model.

REFERENCES

Box, G.E.P., G. M. Jenkins and G. C. Reinsel (1994). Time Series Analysis – Forecasting and Control. Prentice-Hall International, London.

Collins, J. P. and R. B. Langley (1999). Possible weighting schemes for GPS carrier phase observations in the presence of multipath. Final contract report for the

U.S. Army Corps of engineers Topographic Engineering Center, No DAAH04-96-C-0086 / TCN 98151, March, 33pp.

El-Rabbany, A. (1994). The effect of physical correlations on the ambiguity resolution and accuracy estimation in GPS differential positioning. Ph.D. Thesis - University of New Brunswick, Department of Geodesy and Geomatics Engineering, Fredericton, NB.

Hofmann-Wellenhof, B., H. Lichtenegger and J. Collins (2001). Global Positioning System: Theory and Practice. Springer-Verlag Wien New York.

Kim, D and R. Langley (2001). Estimation of the Stochastic Model for Long-Baseline Kinematic GPS Applications. ION National Technical Meeting, January 22-24, 2001, Long Beach-CA.

Langley, R. B. (1997). GPS observation noise. GPS World, Vol. 8, No. 6, April, pp 40-45.

Vanicek, P. and E. Krakiwsky (1982). Geodesy: The Concepts. North-Holland Publishing Company, Amsterdam – New York – Oxford.